

FINANCIAL LITERACY AND THE TIMING OF TAX-PREFERRED SAVINGS ACCOUNT WITHDRAWALS

MARIANNE **LAURIN**DEREK **MESSACAR**PIERRE-CARL **MICHAUD** 

2021s-36
CAHIER SCIENTIFIQUE

CS



Center for Interuniversity Research and Analysis on Organizations

The purpose of the **Working Papers** is to disseminate the results of research conducted by CIRANO research members in order to solicit exchanges and comments. These reports are written in the style of scientific publications. The ideas and opinions expressed in these documents are solely those of the authors.

Les cahiers de la série scientifique visent à rendre accessibles les résultats des recherches effectuées par des chercheurs membres du CIRANO afin de susciter échanges et commentaires. Ces cahiers sont rédigés dans le style des publications scientifiques et n'engagent que leurs auteurs.

**CIRANO** is a private non-profit organization incorporated under the Quebec Companies Act. Its infrastructure and research activities are funded through fees paid by member organizations, an infrastructure grant from the government of Quebec, and grants and research mandates obtained by its research teams.

Le CIRANO est un organisme sans but lucratif constitué en vertu de la Loi des compagnies du Québec. Le financement de son infrastructure et de ses activités de recherche provient des cotisations de ses organisations-membres, d'une subvention d'infrastructure du gouvernement du Québec, de même que des subventions et mandats obtenus par ses équipes de recherche.

#### CIRANO Partners – Les partenaires du CIRANO

#### Corporate Partners - Partenaires corporatifs

Autorité des marchés financiers

Bank of Canada

Bell Canada

**BMO Financial Group** 

Business Development Bank of Canada

Caisse de dépôt et placement du Québec

Desiardins Group

Énergir

Hydro-Québec

Innovation, Science and Economic Development Canada

Intact Financial Corporation

Manulife Canada

Ministère de l'Économie, de la Science et de l'Innovation Ministère des finances du Québec

National Bank of Canada

Power Corporation of Canada

PSP Investments

Rio Tinto

Ville de Montréal

#### Academic Partners - Partenaires universitaires

Concordia University

École de technologie supérieure

École nationale d'administration publique

HEC Montréal

McGill University

National Institute for Scientific Research

Polytechnique Montréal

Université de Montréal

Université de Sherbrooke

Université du Québec

Université du Québec à Montréal

Université Laval

CIRANO collaborates with many centers and university research chairs; list available on its website. Le CIRANO collabore avec de nombreux centres et chaires de recherche universitaires dont on peut consulter la liste sur son site web.

© September 2021. Marianne Laurin, Derek Messacar and Pierre-Carl Michaud. All rights reserved. *Tous droits réservés*. Short sections may be quoted without explicit permission, if full credit, including © notice, is given to the source. *Reproduction partielle permise avec citation du document source, incluant la notice* ©.

The observations and viewpoints expressed in this publication are the sole responsibility of the authors; they do not necessarily represent the positions of CIRANO or its partners. Les idées et les opinions émises dans cette publication sont sous l'unique responsabilité des auteurs et ne représentent pas nécessairement les positions du CIRANO ou de ses partenaires.

# Financial Literacy and the Timing of Tax-Preferred Savings Account Withdrawals

Marianne Laurin \*, Derek Messacar †, Pierre-Carl Michaud ‡

#### Abstract/Résumé

Tax deductions on contributions to registered savings vehicles are a common policy tool used by governments in many industrialized countries to encourage people to save for retirement. However, these plans do not typically lock in funds, which means savers may also withdraw before retirement when their marginal tax rates are still high and forgo the tax benefit. In this paper, we investigate the extent to which pre-retirement savings withdrawals respond to changes in the net-of-tax benefit of withdrawing and whether such behavior depends on the saver's financial literacy. To that end, we link respondents of a nationally representative financial capability survey from Canada to over 15 years of administrative tax data. Our results show that the correlation between savings withdrawals and the effective marginal tax rate is negative for those with higher financial literacy, but much weaker and sometimes statistically insignificant for those with lower financial literacy. The findings suggest that financial literacy is an important determinant of the extent to which tax-deductible savings plans are used efficiently.

Les déductions fiscales sur les contributions aux véhicules d'épargne enregistrés sont un outil politique commun utilisé par les gouvernements de nombreux pays industrialisés pour encourager les gens à épargner pour leur retraite. Cependant, ces plans ne bloquent généralement pas les fonds, ce qui signifie que les épargnants peuvent également retirer leurs fonds avant la retraite lorsque leur taux d'imposition marginal est encore élevé et renoncer à l'avantage fiscal. Dans cet article, nous étudions dans quelle mesure les retraits d'épargne avant la retraite répondent à des changements dans l'avantage net d'impôt du retrait et si ce comportement dépend de la littératie financière de l'épargnant. À cette fin, nous relions les répondants d'une enquête sur les capacités financières représentative au niveau national au Canada à plus de 15 ans de données fiscales administratives. Nos résultats montrent que la corrélation entre les retraits d'épargne et le taux d'imposition marginal effectif est négative pour les personnes ayant une meilleure culture financière, mais beaucoup plus faible et parfois statistiquement non significative pour les personnes ayant une moindre culture financière. Les résultats suggèrent que la littératie financière est un déterminant important de la mesure dans laquelle les plans d'épargne déductibles de l'impôt sont utilisés efficacement.

**Keywords/Mots-clés:** Tax-preferred savings accounts, Retirement savings, Financial literacy / Comptes d'épargne à fiscalité privilégiée, Epargne-retraite, Education financière

JEL Codes/Codes JEL: D14, H31, H51, H53, I10, I31

<sup>\*</sup> HEC Montréal

<sup>&</sup>lt;sup>†</sup> Statistics Canada and Memorial University of Newfoundland

<sup>&</sup>lt;sup>‡</sup> HEC Montréal and National Bureau of Economics Research (NBER)

# Financial Literacy and the Timing of Tax-Preferred Savings Account Withdrawals<sup>1</sup>

Marianne Laurin<sup>2</sup>, Derek Messacar<sup>3</sup> and Pierre-Carl Michaud<sup>4</sup>

September 8th, 2021

#### **Abstract**

Tax deductions on contributions to registered savings vehicles are a common policy tool used by governments in many industrialized countries to encourage people to save for retirement. However, these plans do not typically lock in funds, which means savers may also withdraw before retirement when their marginal tax rates are still high and forgo the tax benefit. In this paper, we investigate the extent to which pre-retirement savings withdrawals respond to changes in the net-of-tax benefit of withdrawing and whether such behavior depends on the saver's financial literacy. To that end, we link respondents of a nationally representative financial capability survey from Canada to over 15 years of administrative tax data. Our results show that the correlation between savings withdrawals and the effective marginal tax rate is negative for those with higher financial literacy, but much weaker and sometimes statistically insignificant for those with lower financial literacy. The findings suggest that financial literacy is an important determinant of the extent to which tax-deductible savings plans are used efficiently.

**Keywords:** tax-preferred savings accounts, retirement savings, financial literacy

JEL Codes: G53, G51, D14

<sup>&</sup>lt;sup>1</sup> The authors wish to thank the Financial Consumer Protection Agency of Canada for financial support for this research, the Retirement and Savings Institutes at HEC Montréal and Statistics Canada for performing the data linkage for this study.

<sup>&</sup>lt;sup>2</sup> HEC Montréal

<sup>&</sup>lt;sup>3</sup> Statistics Canada and Memorial University of Newfoundland

<sup>&</sup>lt;sup>4</sup> HEC Montréal and National Bureau of Economics Research (NBER)

# 1. Introduction

In most industrialized countries, tax incentives are the main tool used by governments to encourage people to save for their retirement. In Canada, Registered Retirement Savings Plans (RRSPs) allow taxpayers to make savings contributions with before-tax income and defer paying tax until these funds are withdrawn. Although this program is explicitly labeled as a retirement savings vehicle, funds do not lock in and there are no early withdrawal penalties imposed by the tax system. As a result, pre-retirement withdrawals are very common in practice. On average, one dollar is withdrawn each year for every five dollars contributed by working-aged savers (Giles and Maser 2004; Messacar 2017).<sup>5</sup>

The tax advantage offered by an RRSP is largely determined by the contributor's marginal effective tax rate (METR) at the time of contributing and when the withdrawals are made. Since RRSPs allow tax on the amounts saved to be deferred to a future period, the net value of a withdrawal is larger when it is made in a year when the METR is relatively low compared with other years. Due to tax progressivity, most taxpayers have higher tax rates during their working lives than in retirement, where income replacement rates are typically around 60 to 80 percent. However, the METRs that apply during retirement can sometimes be higher than the ones paid by individuals during their working lives (Shillington 2003; Laurin and Poshmann 2010). This is particularly true among low-income earners since public pensions and other transfer programs can lead to income replacement rates that exceed income from working years. Therefore, the optimal time to make an RRSP withdrawal can vary based on personal circumstances.

In a previous study analyzing RRSP withdrawals, Mawani and Paquette (2011) show that such behavior occurs frequently when tax rates are high. They find RRSPs are used for precautionary saving and income-smoothing purposes during times of financial hardship, such as after a job loss or divorce. The extent to which these tax mistakes are driven by low financial literacy is an unresolved empirical issue. The economic literature repeatedly underscores the critical role that

<sup>&</sup>lt;sup>5</sup> This differs with Individual Retirement Accounts (IRAs) in the United States, which impose a 10 percent penalty on withdrawals made before the age of 59.5. In Canada, funds held in tax-deductible pension plans set up by employers lock in funds notwithstanding special cases of financial hardship.

financial literacy plays in financial decision-making and retirement readiness (Lusardi and Mitchell 2007; Boisclair et al. 2017). In this study, we investigate the extent to which savings withdrawals respond to changes in the net-of-tax benefit of withdrawing and whether financial literacy is a relevant determinant of such behavior.

To this end, we link a nationally-representative financial capability survey conducted by Canada's central statistics agency to over 15 years of administrative tax records. The survey asks numerous questions designed to measure respondents' knowledge of financial issues related to inflation, interest compounding, diversification and related financial market concepts. We use a linear regression model with fixed effects that estimates the effect of a per-unit change in the METR on RRSP withdrawals. Then, we assess whether there are differences in how withdrawals respond to changes in tax rates based on saver's level of financial literacy. Using this information-rich dataset, we provide an overall picture of withdrawal choices made by taxpayers, and we draw conclusions on the optimality of those retirement savings decisions.

This study provides an innovative perspective on the impact of tax incentives for retirement savings by incorporating financial literacy into the analysis. Our approach also stands apart in its use of panel data. More broadly, our study contributes to a growing literature on the various impacts of financial literacy. Many studies have shown that there is a positive relationship between financial literacy and better savings decisions. For example, Lusardi and Mitchell (2011) find that individuals with a high level of financial literacy are more likely to gather the information needed to prepare and successfully implement financial plans for their retirement. More recently, a study by Boyer et al. (forthcoming) looks at the impact of financial literacy on the choice between the two types of tax-advantaged savings accounts in Canada: the RRSP and the Tax-Free Savings

\_

<sup>&</sup>lt;sup>6</sup> In recent years, a number of studies have examined the connection between financial literacy and retirement readiness and found a strong association between the two (Lusardi and Mitchell 2007, 2011, 2014; van Rooij et al. 2012; Mullock and Turcotte 2012; Lalime and Michaud 2014). For example, a study by Boisclair et al. (2017) found that individuals who correctly answered questions about interest rates, inflation and risk were 10% more likely to have saved for their retirement. Financial literacy has also been associated in the literature with greater wealth accumulation, partly because individuals with a better financial education are better informed about how the pension system works. They also tend to have more diversified retirement investments, pay lower investment fees, and earn a higher annual return (Chan and Stevens 2008; Hastings et al. 2011; Clark et al. 2017). Mullock and Turcotte (2012) estimate that the likelihood of knowing how much to save for reaching one's retirement goals increases by 30% when comparing the group with the lowest financial literacy (score lower than 50%) to the group with the highest (score greater than 85%), and that is among those who have retirement savings.

Account (TFSA). To our knowledge, our study is the first to establish a link between the interconnected issues of saving, tax incentives and financial literacy.

The rest of this study is structured as follows. The next section focuses on the economic theory pertaining to the optimal timing of an RRSP withdrawal. Section 3 presents the characteristics of the database used, along with relevant descriptive statistics, while section 4 describes the methodology used for the data analysis. Section 5 is dedicated to presenting the results. Lastly, we conclude with a brief discussion of the implications of our findings, plus avenues for future research.

# 2. Optimal Timing of RRSP Withdrawal

Tax-preferred savings accounts are defined based on whether the amount saved is subject to tax (Yoo and de Serres 2004). The RRSP<sup>7</sup> is considered an EET (exempt-exempt-taxed) account because the amount contributed is tax exempt (E), the accumulated returns are as well (E), but the amount withdrawn is taxable (T). Annual contributions to the plan are limited by a ceiling; however, any unused contribution room accumulates and can be carried forward indefinitely. Unlike other retirement savings plans around the world, there are no penalties on RRSP withdrawals made before retirement. However, the amounts withdrawn cannot be re-contributed. In other words, the contribution room is reduced by the amount withdrawn. Also, by withdrawing funds from an RRSP, the saver is forgoing the capitalization of the income from those tax-sheltered investments, which can significantly impact the value of the long-term savings.

However, in certain situations, making an RRSP withdrawal before retirement can be a rational decision. In order to decide on the optimal time to withdraw from an RRSP, the individual must simultaneously factor in their current METR and what they anticipate their METR will be in future periods. That said, adequately estimating the METR involves a comprehensive understanding of

<sup>7</sup> 

<sup>&</sup>lt;sup>7</sup> In Canada, private tax-preferred retirement savings vehicles such as the RRSP and TFSA are one of the three pillars of the retirement income system, along with employment-based pension plans and government-administered plans. Any Canadian with earned income is allowed to contribute to an RRSP. Consumers can have multiple RRSP and TFSA accounts with different financial institutions. Contributions are limited on total contributions. The maximum age to make contributions to RRSP is 70. One cannot obtain loans from RRSP or TFSA but can withdraw for either accounts without penalty (other than tax withholding). In 2016, nearly 80% of Canadians planned to contribute to an RRSP during their lifetime and 31% expected that the RRSP would be their main source of income in retirement (Uppal, 2016).

the tax system and of the clawback rates of the assistance programs for which the taxpayer may be a recipient (Duclos et al. 2008; Blancquaert et al. 2017). It is repeatedly pointed out in the literature that a large proportion of taxpayers have only a partial understanding of the tax system. In fact, it has been shown that the progressive nature of the tax system is often misunderstood and that, as a result, taxpayers underestimate the increase in marginal rates resulting from by a higher income (Gideon 2017; Rees-Jones and Taubinsky 2019). In addition, Feldman et al. (2016) find that taxpayers have difficulty interpreting an METR variation that results from a predictable change in eligibility for a tax credit, using the example of the Child Tax Credit in the United States. Miller and Mumford (2015) show that taxpayers respond to a change in a tax measure in isolation and usually do not factor in the interactions with the other tax measures, even if they affect their METRs.

Numerous studies also address optimal choice between the two main types of tax-advantaged savings accounts, specifically in the Canadian context: the RRSP (EET-type) and the tax free savings account (TFSA; TEE-type). The problem of the optimal time for an RRSP withdrawal is very similar to that of choosing between these two types of savings vehicles because, in both cases, a rational decision must be made by comparing the METR at contribution time and at withdrawal time. In Quebec, the TFSA is chosen as a savings vehicle about 30% of the time; whereas it is predicted to be the optimal choice in about 70% of cases (Marchand, 2018). Boyer et al. (forthcoming) show that Canadians do not choose any better than if they had made a random decision when faced with choosing between an RRSP and a TFSA. That said, they find that a financial literacy intervention providing information about the tax implications of both types of accounts can increase both the understanding of how these accounts work and the quality of the savings decision.

Empirical evidence indicates that multiple Canadians use their RRSP as a contingency fund and make a withdrawal before retirement when a negative income shock occurs unexpectedly (Mawani and Paquette 2011). In a context where TEE-type tax-advantaged savings accounts (e.g., TFSA) are available, the economic agent can decide to make an RRSP withdrawal at a given period without affecting their ability to save for future periods. In fact, a portion of the amount withdrawn can simply be put into a TFSA so that the funds continue to generate a tax-sheltered compound interest return.

Given the tax treatment of RRSP withdrawals, individuals should opt to withdraw when their effective marginal rate is the lowest in order to minimize the tax that they pay. In the Appendix 1, we present a simple 2 period economic model which shows that the optimal timing involves minimizing the effective marginal tax rate at the time of withdrawing. This means that only the marginal effective tax rates matter in deciding the optimal time to make an RRSP withdrawal and that preferences should not play any role. Depending on the difference between the METRs in the first and second periods, making a withdrawal at the wrong time can have considerable relative costs. Low-income households are especially vulnerable to this type of error and, depending on the size of the withdrawal, the relative cost can considerably impact their personal finances. This result is consistent with what is found by Boyer et al. (forthcoming) and Grenier (2019) in the context of choosing between an RRSP and a TFSA.

## 3. Data

#### 3.1 Data sources

Administrative data typically does not contain information on financial literacy. Similarly, survey data rarely contain longitudinal data on contributions and withdrawals to tax-preferred savings accounts. For this project, we have created the Financial Capability, Employment and Income Database (FCEID), which contains the Canadian Financial Capability Survey (CFCS) matched with T1 personal income tax records (tax forms). Administered by Statistics Canada in 2014, the CFCS collected information about Canadians' knowledge and skills regarding financial decision-making and about the respondents' financial planning behaviours. The purpose of the survey is to assess Canadians' financial situation and plans for the future along with how well they understand the financial services available to them<sup>8</sup>. The longitudinal data from the income tax return forms are provided by the central tax authority, Canada Revenue Agency, for the years 2000 to 2016. The data contain a wide range of information on income, tax credits, deductions, retirement savings, and taxes payable by the filers. A unique, anonymous identifier for each person is used

<sup>&</sup>lt;sup>8</sup> The exact list of questions used for this study can be consulted under the "Objective personal assessment" section of the following link:

https://www23.statcan.gc.ca/imdb/p3Instr.pl?Function=assembleInstr&a=1&&lang=en&Item Id=201522.

for matching these annual data with the individual's responses in the 2014 CFCS. This linkage procedure was done by a methodology team at Statistics Canada, not by ourselves, to further protect the confidentiality of the survey respondents.

### 3.2 Sampling and data linkage

The CFCS is a sample survey with a cross-sectional design whose target population is all Canadians aged 18 and older residing in one of the country's 10 provinces. The database contains analytical weights that enable the data to reflect the demographic composition of the overall Canadian population. The sample consists of a total of 6,685 individuals questioned in the context of the survey, 6,229 of whom were then able to be matched with their annual tax data. As such, a total of 105,893 observations are entered into the panel from 2000 to 2016. The match rate is high. We have checked that excluded respondents are not statistically different from those retained in the matched file. While respondents are slightly older, we cannot report meaningful differences across the two groups. Some restrictions are placed on the sample for the purpose of obtaining estimates representative of the population of interest. We restrict the sample to people who were between the ages of 30 and 70 when they participated in the CFCS in 2014 because we consider this to be the age group most likely to contribute to an RRSP. We keep tax records between the ages of 18 and 70 for each year observed. The final sample consists of a total of 4,476 respondents and 69,228 tax-year observations between the years 2000 and 2016.

## 3.3 Measuring financial literacy

Respondents in the CFCS were asked a set of 14 questions for the purpose of objectively assessing their knowledge level of financial concepts. The questions dealt with such concept as inflation, knowledge of various types of investment products, the stock market, interest rates, and a number of other topics pertaining to personal finance.<sup>10</sup> In what follows, we use the number of correct answers to those questions as an index of the respondents' financial literacy level. The scores are grouped into two categories based on the median observed in the sample, either low (score of 9/14

<sup>&</sup>lt;sup>9</sup> The age difference is likely the results of students not having to file taxes while studying and therefore not having tax records in the years preceding the CFCS.

<sup>&</sup>lt;sup>10</sup> The exact list of questions can be consulted under the "Objective personal assessment" section of the following link: <a href="https://www23.statcan.gc.ca/imdb/p3Instr.pl?Function=assembleInstr&a=1&&lang=en&Item\_Id=201522">https://www23.statcan.gc.ca/imdb/p3Instr.pl?Function=assembleInstr&a=1&&lang=en&Item\_Id=201522</a>. Appendix 2 shows the percentage of correct answers obtained for each question by financial literacy group.

or less) or high (score of 10/14 or more). Overall, 2,362 individuals fall into the low financial literacy category versus 1,855 individuals in the high financial literacy category. There are 259 individuals who did not answer any of the 14 questions, and therefore do not belong to any of the two groups. The distribution of scores is shown in Figure 1.

Financial literacy varies substantially across demographic and economic dimensions. Table 1 presents a set of descriptive statistics for each financial literacy group as well as the t-and-p-value statistics of a Student's t-test on the difference between the averages of the two groups. All the differences in averages are statistically significant at the 99% confidence level with one exception, which confirms that the differences between the two financial literacy groups are indeed significant.

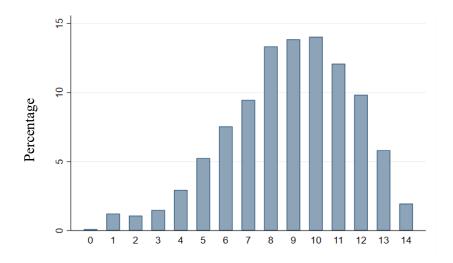


Figure 1 – Distribution of financial literacy score on 14 questions: The figure shows the percentage of the sample that obtained each number of good answers. The individuals who did not answer any questions are counted as missing.

By comparing the averages of the two groups presented in Table 1, we first see that the low financial literacy group consists of a larger proportion of women and individuals with a marital status other than married. Also, the averages of the education and income variables confirm that the group with better financial literacy generally reaches higher education and income levels. In particular, nearly double the proportion of the high financial literacy group completed a bachelor's

<sup>&</sup>lt;sup>11</sup> Generally, we have found our results to be relatively insensitive to the cut-off used for the two groups.

<sup>&</sup>lt;sup>12</sup> The individuals who did not answer any of the financial literacy questions are dropped from the sample. The average age of this group is 44.71 years old and 55.51% are women. On average, 32.73% have withdrawn from their RRSP at least once between 2000 and 2016 and the mean amount withdrawn is 274.12\$. The average METR for this group is 18.12%. When adding individuals from this group into the low financial literacy group, our main results do not change significantly.

degree, and they earn an employment income roughly \$15,000 higher on average. In addition, the average interest and dividend income for this group is nearly triple that of the low financial literacy group. Table 1 also shows us that the proportion of the low financial literacy group who made an RRSP withdrawal at least once between 2000 and 2016 is 43%, versus an average of 48% for the high financial literacy group. The average amounts of RRSP withdrawals are also significantly higher for the high financial literacy group.

		M	ean	t test	p value
		Low FL	High FL		
Demographics					
	Age	44.58	44.3	10	0
	Female (%)	55.65%	46.34%	20	0
	Married (%)	52.32%	60.44%	-18	0
Education					0
	High School Not Completed	15.28%	6.75%	51	0
	High School Diploma	22.56%	14.45%	15	0
	Some Postsecondary or University Certificate	44.19%	42.93%	2	0.06
	Bachelor Degree	12.83%	23.09%	-38	0
	Higher than Bachelor Degree	5.14%	12.78%	-35	0
<b>Total Income</b>					
	Employment Income	28,020	43,428	-39	0
	Self Employment Income	1,011	3,279	-13	0
	Investment/Dividend Income	1,069	2,969	-11	0
	Unemployment Insurance Income	762	564	17	0
Sources of Income	Employment Income	76.99%	80.68%	-16	0
	Self Employment Income	7.87%	12.32%	-16	0
	Investment/Dividend Income	25.87%	39.22%	-35	0
	Unemployment Insurance Income	13.6%	9.35%	20	0
Assets (2014)	Average estimated total value of financial assets	112,874	217,876	-3	0.01
,	Average estimated total value of RRSP	92,116	158,495	-6	0
RRSP Withdrawals	At Least One Withdrawal Between 2000-2016	43.1%	48.02%	-15	0
	Amount Withdrawn	387.42	504.63	-7	0
METR		20.42%	26.278%	-44	0

Table 1 – Descriptive statistics by level of financial literacy: Individuals with a score of 9/14 or less are in the low FL group, and those who scored 10/14 or more are in the high FL group. Income figures are in nominal dollars. The assets values are estimated in 2014 by the CFCS respondents,

#### 3.4 Who withdraws more from an RRSP?

Table 2 gives descriptive statistics about RRSP withdrawals for three different subsamples. First, when factoring in all the observations in our database, a negative relationship emerges between RRSP withdrawals and total income quartiles. In fact, taxpayers in the first income quartile are the ones who withdraw the most, on average, while the last quartile withdraws the least. When we eliminate from the sample all individuals who made no withdrawals between 2000 and 2016, we see larger withdrawals on average for the high financial literacy group compared with the other group. We can also see that the relationship between RRSP withdrawals and total income quartiles appears to be U shaped. The average amount withdrawn decreases with income as far as the third quartile, and then rises at the fourth quartile.

We also produce the same statistics for the group with positive RRSP withdrawals. This enables us to, among other things, isolate the impact of the size of the amounts withdrawn on the averages given above. We see that the difference in average amount withdrawn between the high and low financial literacy groups persist. Also interesting to note is that the proportion of observations of non-zero withdrawals among the individuals who had made at least one withdrawal is very similar between the two groups, namely 20.99% and 21.16% respectively for those with low and high financial literacy. The relationship between income quartiles and withdrawals is still a quadratic trend, with the first quartile's average amount withdrawn being significantly higher than that of the other quartiles. This could be explained by liquidity constraints: consumers also use their RRSP to finance consumption in periods where income is low.

Descriptive Statistics on RRSP Withdrawals						
All Observations	Mean	Median	90%	95%	Nb. obs.	
Financial Literacy Group						
Low	387.42	0	0	1,333	36,395	
High	504.63	0	11	1,646	29,038	

Income Quartile					
17,752\$ or less	504.12	0	0	1,333	17,307
17,754\$ to 35,460\$	406.18	0	0	1,560	17,308
35,461\$ to 57,939\$	399.58	0	128	1,657	17,306
57,940\$ or more	392.34	0	0	1,300	17,307
Total	427.57	0	0	1,333	69,228
At Least One RRSP Withdrawal Between 2000-2016					
Financial Literacy Group					
Low	899.88	0	1,988	5,181	15,479
High	1,053.27	0	1,898	6,652	14,044
Income Quartile					
17,752\$ or less	1,573.76	0	4,999	10,250	5,757
17,754\$ to 35,460\$	873.80	0	2,040	5,555	8,052
35,461\$ to 57,939\$	746.57	0	1,394	5,000	8,839
57,940\$ or more	839.47	0	1,333	4,973	8,483
Total	964.35	0	1,903	5,829	31,131
Observations with non-zero RRSP Withdrawals					
Financial Literacy Group					
Low	4,246.24	1,666	10,786	18,192	3,25
High	5,009.10	1,600	14,847	20,000	2,972
Income Quartile					
17,752\$ or less	6,827.88	3,467	19,706	26,666	1,365
17,754\$ to 35,460\$	4,532.53	2,298	11,440	16,455	1,625
35,461\$ to 57,939\$	3,527.13	1,333	9,686	14,267	1,858
57,940\$ or more	4,088.89	1,333	11,890	17,918	1,679
Total	4,616.51	1,666	13,152	19,515	6,527

**Table 2– Descriptive statistics on RRSP withdrawals:** The low and high financial literacy groups respectively scored 9/14 or less and 10/14 or more. Income quartiles are defined on the basis of total income, from which the RRSP withdrawal made in the year is subtracted, if any.

# 3.5 Estimating marginal effective tax rates

We estimate for each respondent and tax year, the annual marginal effective tax rate (METR) using the Canadian Tax and Credit Simulator (CTaCS), <sup>13</sup> which contains a large number of Canadian

<sup>&</sup>lt;sup>13</sup> See Milligan (2016). A comprehensive list of the CTaCS input and output variables is described in the user guide at the following link: <a href="https://faculty.arts.ubc.ca/kmilligan/ctacs/">https://faculty.arts.ubc.ca/kmilligan/ctacs/</a>. Among the input variables are sociodemographic

tax rules and parameters from the years 1962 to 2016. The METR calculated from the estimates provided by CTaCS account for the federal and provincial taxes, but also the impact of tax credits, deductions and benefits on disposable income. To compute the METR we simulate a change in RRSP withdrawals and observe the effect on after tax income. We first estimate disposable income  $\pi_n$  based on the vector  $Y_{i,t}$ , i.e., all annual tax variables entered into CTaCS, and the variable w, i.e., RRSP withdrawals, which is set at zero in the baseline. We then repeat the exercise, changing the value from w to \$1,000 for all observations. With these two disposable income estimates, we are now able to calculate the METR per year t for each individual i, which is defined as follows:

$$METR_{i,t} = 1 - \frac{\pi_n(Y_{i,t}, w) - \pi_n(Y_{i,t}, 0)}{w}$$

Figure 2 gives a general overview of the distribution of the estimated METR observed in our sample. We find a standard deviation of 0.158 and 0.152 respectively for the low and high financial literacy group. In other words, fluctuations in METR over time are similar on average for both financial literacy group. The within standard deviations are respectively 0.128 and 0.1204 for the low and high financial literacy group. The between standard deviations are equal to 0.093 for both financial literacy groups. Interestingly, these statistics indicate that tax rates vary more within individuals over time than they do across individuals at a point in time, on average. This likely arises because effective tax rates depend on both statutory rates and a plethora of tax and transfer programs that interact with each other and indirectly affect how much taxes people pay. Small changes in income from one year to the next move individuals around these discontinuity points and cause rates to vary meaningfully over time. Hence, timing likely matters a lot when making savings withdrawal for tax planning purposes.

-

variables such as sex, age and marital status, plus financial variables such as income from various sources and several expenses and contributions that can qualify the taxpayer for certain tax deductions or credits. Based on that, the simulator puts out such things as total tax payable, as well as the main government transfers and federal and provincial refundable and non-refundable credits.

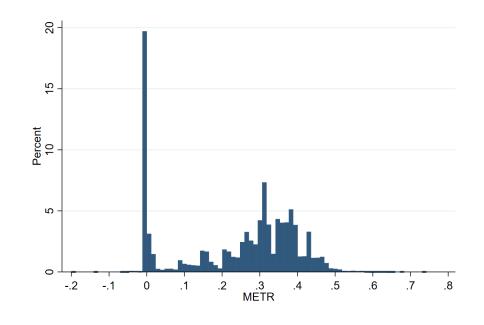


Figure 2 – Distribution of the estimated METR: The figure shows the distribution in our sample of the marginal effective tax rates on RRSP withdrawals estimated with CTaCS. The mean is 22.7% and the median is 27.3%.

# 4. Methodology

## 4.1 Linear regression models with fixed effects

The theoretical prediction is that withdrawals should be correlated negatively with the METR at the respondent level over time: i.e. respondents should time withdrawals when marginal tax rates are low. We implement a fixed effect regression framework that exploits this idea. If an individual makes a withdrawal during a period when their METR is above the average, they are potentially incurring a loss from a financial perspective. By estimating this regression separately by financial literacy group, we have a direct test of whether the timing of widthdrawals is more tax efficient among those with higher financial literacy. Precisely, we postulate the following fixed effect linear regression:

$$R_{i,t} = \alpha METR_{i,t} + X'_{i,t}\beta + \mu_i + \varepsilon_{i,t}$$
 (1)

where RRSP withdrawals  $R_{i,t}$ , the marginal effective tax rate  $METR_{i,t}$  and a set of explanatory variables  $X_{i,t}$  are observed for each individual i over several periods t and an unobservable term

 $\mu_i$  captures the time-invariant characteristics for each individual i. Lastly, the term  $\varepsilon_{i,t}$  represents the residuals. Standard errors are clustered at the respondent level. We first produce an initial specification of our model where our vector of explanatory variables  $X_{i,t}$  contains only demographic variables. We include a dummy variable for each age level between the ages of 18 and 70 and a dummy variable for each marital status. This specification enables us to measure the impact of the METR on the amount withdrawn from the RRSP without controlling for income. We than add two explanatory variables to the vector  $X_{i,t}$ , namely  $y_{i,t}$  and  $y_{i,t}^2$ , which represent, respectively, total income<sup>14</sup> and total income squared. The quadratic term is included to enable a non-linear relationship between our dependent variable and total income. This allows to account for mechanical changes in METR which are due to a change in circumstances that may prompt a withdrawal. By controlling for income, we are effectively comparing respondents with similar changes in income from one year to the next but for whom METRs vary for other reasons (because of variation induced by provincial differences or changes over time in tax rules). Lastly, the model's third specification contains a more varied set of controls for income including employment income, self-employment income, investment and dividend income, employment insurance, and social assistance. We also include a quadratic term for the first three income types mentioned because they can reach fairly large amounts and, therefore, their relationship to RRSP withdrawals could be non-linear. For each of the specifications, we estimate the coefficients twice: once for each financial literacy group.

Then we develop a fully interactive model in order to determine if the difference between the coefficients obtained for the high and low financial literacy groups is statistically significant. This model simply involves adding the interaction terms between the high financial literacy dummy variable  $FL_i$  and the set of explanatory variables to the previous regression. This gives us the following regression:

$$R_{i,t} = \alpha_0 METR_{i,t} + \alpha_1 METR_{i,t} \cdot FL_i + X'_{i,t}\beta_0 + X'_{i,t}\beta_1 \cdot FL_i + \mu_i + \varepsilon_{i,t}$$
 (2)

In this model, the coefficients  $\alpha_0$  and  $\beta_0$  of equation 2 are identical to the coefficients  $\alpha$  and  $\beta$  of equation 1 estimated for the low financial literacy group. The coefficients  $\alpha_1$  and  $\beta_1$  correspond

<sup>-</sup>

<sup>&</sup>lt;sup>14</sup> Total income is the sum of all sources of income before taxes, as defined by the Canada Revenue Agency (line 15000 of the income tax return), minus the income from an RRSP withdrawal.

to the difference between the coefficients derived from model 1 for the high financial literacy group and the low financial literacy group. In other words, model 2 captures the additional impact of having high financial literacy on all the explanatory variables. Specifically, using our estimate of the coefficient  $\alpha_1$ , we will be able to determine whether individuals with high financial literacy place greater importance on their METR when deciding how much to withdraw from their RRSP. In fact, if that is true, the coefficient  $\alpha_1$  is expected to be negative, meaning that a marginal increase in the METR would result in a greater reduction in the amount withdrawn from the RRSP in the high financial literacy group.

#### 4.2 Extensions

Our next model aims to estimate the impact of a change in the METR at the extensive margin, i.e. the decision to withdraw funds from RRSPs rather than the total amount. We change our RRSP withdrawals variable  $R_{i,t}$  into a dichotomous variable  $D_{i,t}$  that takes the value 1 if the individual made a withdrawal during the observed period or 0 if not. We then estimate the following linear probability model with fixed effects:

$$D_{i,t} = \alpha_0 METR_{i,t} + \alpha_1 METR_{i,t} \cdot FL_i + X'_{i,t}\beta_0 + X'_{i,t}\beta_1 \cdot FL_i + \mu_i + \varepsilon_{i,t}$$
 (3)

In addition, a significant proportion of individuals in the sample do not make any withdrawals in all the years observed. Although our database doesn't contain any variable indicating whether the individual has an RRSP, we observe at least one withdrawal between the years 2000 and 2016 for roughly 48% of the high financial literacy group versus roughly 43% for the low financial literacy group (see Table 1). In the likely event that individuals with low financial literacy account for a larger proportion of the group without RRSPs, the bias towards zero would be stronger for the coefficients pertaining to this subsample. This could artificially increase the difference between the coefficients for the high and low financial literacy groups. We therefore perform a robustness test for the purpose of checking the validity of our estimates in the absence of this bias. To do so, we remove from our sample all individuals who made no withdrawals between 2000 and 2016, and we estimate model (2) from that subsample.

Lastly, we seek to determine whether the results are robust across various demographic groups. To do so, we divide our sample by sex, province and mother tongue, and we estimate model 2 on each subsample. By limiting the sample to the observations of certain groups, we can check whether the results obtained hold up and how they differ by group.

## 5. Results

## **5.1 Main Regression Results**

To start, we present the estimates separately for each financial literacy groups (equation 2) in Table 3. For each subsample, the first column gives the estimates with explanatory variables only for age and marital status; the second column adds control variables for total income; and the third column includes explanatory variables for various types of income.

We find that the METR has a significant negative effect on RRSP withdrawals for both financial literacy groups and this effect persists with the addition of various control variables. In other words, the null hypothesis that a change in the METR has no impact on RRSP withdrawals is rejected with a 1% confidence level for individuals with low and high financial literacy. In the specification with only demographics controls, the METR elasticities of -0.96 and -1.57 were obtained for the low and high financial literacy groups respectively. This implies that a 10% increase in the METR reduces the amount withdrawn from the RRSP by 9.6% on average for the low financial literacy group and 15.7% on average for the high financial literacy group. When comparing the two subsamples, we also notice that the coefficient for the METR variable is larger in absolute value for those with high financial literacy. This result is also consistent with our original assumption that a high level of financial literacy makes people more responsive to their METR when deciding how much to withdraw from their RRSP. In the next section, we will test whether this difference is statistically significant.

Fixed Effect Regression on RRSP Withdrawals					
	Low Financial Literacy	High Financial Literacy			

\_

<sup>&</sup>lt;sup>15</sup> For the low FL group, the coefficient is -18.258, while the average of the METRs and withdrawals is 20.42 and 387.42, respectively (Table 1). For the high FL group, the coefficient is -30.096, and the average of the METRs and withdrawals is 26.28 and 504.63. As a result, the averaged elasticities are -0.96 and -1.57.

	(1)	(2)	(3)	(1)	(2)	(3)
METR	-18.258*** (2.484)	-9.911*** (2.101)	-7.119*** (2.085)	-30.096*** (3.624)	-27.249*** (3.560)	-23.108*** (3.489)
Income variables						
Total income		-115.007*** (23.031)			-24.9556*** (9.001)	
Total income ^2		1.339e-04*** (3.153e-05)			3.139e-6*** (1.259e-6)	
Employment income			-135.532*** (21.648)			-54.879*** (12.819)
Employment income ^2			1.77e-04** (7.67e-05)			1.54e-05*** ( 3.97e-6)
Investment income			-110.757*** (30.997)			-47.197*** (17.992)
Investment income ^2			1.25e-04*** ( 3.66e-05)			5.18e-05*** (1.99e-05)
Net self employment incom	e		-260.858*** (69.237)			-151.333*** (31.623)
Net self employment incom	ne ^2		3.06e-04*** ( 8.28e-05)			2.92e-04*** (7.12e-05)
Unemployment insurance in	ncome		-168.589*** (54.152)			13.741 (75.505)
Social assistance income			-287.090*** (60.743)			-401.917*** (89.873)
Marital Status						
Common law	88.128 (81.828)	69.113 (82.457)	54.968 (82.671)	-113.105 (91.805)	-120.25 (92.035)	-112.588 (91.857)
Widow	522.61** (224.454)	531.61** (223.22)	444.222** (222.405)	267.834 (299.416)	265.489 (299.05)	242.621 (298.269)
Divorced	303.288* (170.506)	300.464* (170.936)	289.933* (170.953)	-81.62 (166.553)	-87.799 (166.414)	-78.428 (166.329)
Separated	205.07* (115.41)	197.471* (116.234)	187.614 (115.384)	275.5* (147.862)	268.607* (148.493)	278.803* (149.073)
Single	166.515** (75.634)	131.202* (76.658)	114.598 (77.604)	144.682 (119.518)	124.667 (119.765)	130.261 (119.989)
(Married)	-	-	-	-	-	-
Age Fixed Effect	Included	Included	Included	Included	Included	Included
Intercept	333.698*** (88.513)	196.076*** (88.76)	157.415** (88.132)	711.813*** (127.64)	636.784*** (128.287)	560.015*** (127.276)
Number of observations	36,244	36,178	36,244	28,907	28,818	28,907

**Table 3 – Fixed effect regression on RRSP withdrawals**: Coefficients are estimated by OLS. Robust standard errors are shown in parentheses. The METR variable is expressed as a percentage multiplied by 100. Income variables are divided by 10 000. Total income refers to all sources of income, from which RRSP withdrawals are subtracted. Investment income is the combined dividend and interest income. A dummy variable is included for each age level between 18 and 70, but is not presented above. \*\*\*, \*\*\*, \* denote statistical significances at the 1, 5 et 10 percent levels, respectively.

By including only explanatory sociodemographic variables, we see that the METR coefficient is larger than the two specifications with income controls, especially for the subsample with low financial literacy. Without controlling for income, the METR coefficient partially captures the impact of a change in income on RRSP withdrawals due to the positive covariance between these two variables. Therefore, this bias accentuates the negative value of the METR coefficient. When adding two control variables for total income in the second column, the coefficient associated with the METR decreases by almost half for the low financial literacy group; whereas a similar effect of much smaller magnitude is seen for the high financial literacy group. In fact, this produces METR elasticities of -0.52 and -1.42 for the low and high financial literacy groups, respectively. In the third model, we replace the two control terms for total income with a number of explanatory variables for taxpayers' main sources of income. This specifically allows each type of income to have a distinct impact on RRSP withdrawals. All the coefficients associated with these explanatory variables are statistically significant at the 1% threshold, with one exception. For the low financial literacy group, an elasticity of -0.375 is estimated while the elasticity is -1.2 for the group with high financial literacy. These estimates are significant at the 1% threshold.

In Figure 2, we plot age dummies for both financial literacy groups at each distinct age between the ages of 26 and 70. These coefficients represent the marginal effect on RRSP withdrawals at each age level, using age 25 as the point of reference. All the coefficients presented in the chart are statistically significant at the 1% threshold. We see that the two trends are relatively similar until the individuals reach age 60. At this age, Canadians can choose to start receiving part of their Canada Pension Plan (CPP) benefit<sup>16</sup>. Withdrawals are more frequent for the high financial literacy group before the age of 65. At age 65, Canadians become eligible for the Guaranteed Income Supplement, a means-tested benefit, which substantially increases marginal effective tax rates for low earnings. One possibility is that respondents with higher financial literacy are better at avoiding withdrawing at those ages.

<sup>-</sup>

<sup>&</sup>lt;sup>16</sup> The Canada Pension Plan/Quebec Pension Plan is a contributory, earnings-based program that provides financial support during retirement. Normally, CPP/QPP benefits begin at age 65. However, Canadians can take up their benefit as early as age 60 and incur 0.6% reduction to the benefit for each month collected before turning 65.

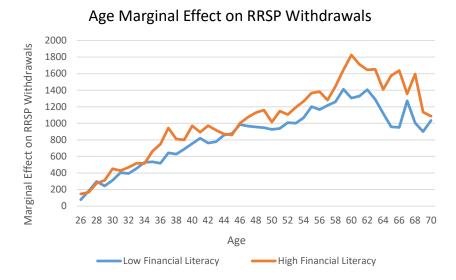


Figure 2 – Age marginal effect on RRSP withdrawals: The graph above shows the trend of the coefficients estimated by OLS for each age between 26 and 70 years old, with 25 years old as a reference point. The estimates come from the 3rd specification of model 1.

# 5.2 Testing the difference between financial literacy groups

We also estimate a fully interacted model to test the difference in the effect of METR on withdrawals between the two financial literacy groups. Results are presented in Table 4. The interaction term is always negative and statistically significant at the 1% confidence level. In other words, this means that individuals with high financial literacy are significantly more responsive to a change in the METR when they have to decide how much to withdraw from their RRSP each period.

Fully Interacted Regression with Fixed Effects on RRSP Withdrawals					
	(1)	(2)	(3)		
METR * FL	-11.838*** (4.393)	-17.338*** (4.133)	-15.989*** (4.064)		
Number of observations	65,131	65,131	65,131		

**Table 4 – Fully interacted regression with fixed effects on RRSP withdrawals**: Coefficients are estimated by OLS. Robust standard deviations are shown in parentheses. The METR variable is expressed as a percentage multiplied by 100. Income variables are divided by 10 000. Total income refers to all sources of income, from which RRSP withdrawals are subtracted. Investment

income is the combined dividend and interest income. Interaction terms of marital status and age and all non interacted variables are not presented. \*\*\*, \*\*, \* denote statistical significances at the 1, 5 et 10 percent levels, respectively.

#### **5.3 Extensions**

#### 5.3.1 Extensive margin

Assume that the decision-making process for making an RRSP withdrawal occurs in two stages. First, the individual decides whether to make an RRSP withdrawal during the period. Then they decide how much to withdraw, if applicable. Our approach so far factored in these two decisions together. However, the METR may have a separate marginal effect on each of these decisions.

The first three columns in Table 5 presents regression estimates using as a dependent variable a dichotomous variable taking the value of 1 if a withdrawal was made during the period, and 0 if not. First, we see that the difference between the marginal effects of the METR of the high and low financial literacy groups captured by the coefficient of the variable METR ·FL is negative and statistically significant at the 1% threshold in the second and third columns. In addition, the coefficient of the METR variable is not statistically significant when adding the diverse set of controls for income in the third column. In other words, we are unable to reject the null hypothesis that the marginal effect of the METR on the decision to withdraw from the RRSP is zero for the low financial literacy group. This implies that the individuals in this group may not factor their METR into the decision about whether to make an RRSP withdrawal during the period.

By adding controls for income, we see that the difference in marginal effect of the METR between the two groups expands, suggesting that the impact of income on the decision to withdraw is greater in the low financial literacy group. In fact, the coefficients of the income terms have a negative non-linear marginal effect that is statistically greater when financial literacy is low in the case of the total income and employment income variables.

Overall, the results estimated by this model show trends similar to those obtained using our basic models. The high financial literacy group places more importance on their METR not only when deciding how much to withdraw from their RRSP, but also when deciding whether to withdraw in the year observed.

#### 5.3.2 Model limited to individuals who made at least one RRSP withdrawal or contribution

This robustness test involves estimating the model with interactions on a limited sample consisting only of individuals who made at least one RRSP withdrawal or contribution between 2000 and 2016. The objective of this test is to check whether the conclusions we drew using the basic models still hold when excluding individuals who do not use their RRSP during the observed periods and, therefore, for whom the estimated marginal effect of the METR is inevitably zero. The last three columns of Table 5 give the results of that model.

Again, we obtain a coefficient for the interaction term METR · FL that is negative and statistically significant at the 10% threshold in the first column and at the 1% threshold in the second and third columns. We estimate that an increase in the METR of 10 percentage points over their average would encourage individuals with high financial literacy to withdraw on average \$180 less than individuals with low financial literacy. The estimated coefficient of the interaction term in the third column has a value relatively similar to the one obtained with the estimate across the entire sample given in Table 4. Although the marginal effect of the METR on withdrawals is significantly greater for each financial literacy group compared with the previous estimates, the difference between the effects remains relatively constant. We also see that the marginal effects of income on withdrawals align with what was previously obtained, and the coefficients increase in magnitude due to the restriction placed on the sample.

The results estimated by this model do not invalidate the previously drawn conclusions about the difference between the marginal effect of the METR on both financial literacy groups' RRSP withdrawals. On the contrary, obtaining a discernibly larger coefficient for the interaction term of the METR and financial literacy implies that individuals who made no withdrawals during the periods observed pull the coefficients towards zero slightly more for the high financial literacy group. This also explains why we estimate a statistically zero difference between the effects of the two groups' income. Since our database does not enable us to identify taxpayers who do not have an RRSP, we are unable to measure with any certainty the bias towards zero that they generate. It is also important to point out that the observations of individuals who make no withdrawals, but use the RRSP as a savings vehicle, are relevant to our estimates. Therefore, the restriction placed on the sample eliminates all observations that bias the coefficients towards zero, as well as a certain portion of valid observations that have some explanatory power over our question of interest. All told, the estimated results in this subsection enable us to assume with some degree of confidence

that the bias towards zero brought about by individuals who do not use an RRSP does not affect either of the two groups in a disproportionate way that would invalidate the previously obtained results.

Extensions						
	Extensive Margin			Conditional	on having used an 2000-2016	RRSP between
	(1)	(2)	(3)	(1)	(2)	(3)
METR * FL	-0.0005 (0.0003)	-0.0009*** (0.0003)	-0.001*** (0.0004)	-10.2861* (5.6276)	-18.4284*** (5.3017)	-18.1558*** (2.9894)
METR	0.0015*** (0.0002)	-0.0008*** (0.0002)	-0.0004 (0.0002)	-26.3354*** (3.5435)	-15.0832*** (3.0822)	-11.0228*** (2.9894)
Income variables * FL						
Total income * FL		0.007*** (0.002)			117.9298*** (29.8445)	
Total income ^2 * FL		-1.08e- 08*** (2.38e-09)			-0.0002*** (3.873e-05)	
Employment income * FL			0.010*** (0.003)			120.9681*** (29.2860)
Employment income ^2 * FL			-3.32e-08*** (1.09e-08)			-0.0003*** (0.0001)
Investment income * FL			0.004 (0.003)			98.8929** (43.1920)
Investment income ^2 * FL			-4.23e-09 (3.58e-09)			-0.0001** (4.59e-05)
Net self employment income * FL			0.003 (0.004)			227.7514** (92.2683)
Net self employment income ^2 * FL			7.22e-10 (7.36e-09)			-0.0001 (0.0001)
Unemployment insurance income * F	L		0.015 (0.012)			311.8123** (123.6051)
Social assistance income * FL			0.005 (0.018)			-207.0198 (364.9491)
Income variables						
Total income		-0.010*** (0.001)			-143.7199*** (28.2082)	
Total income ^2		1.11e-08*** ( 2.38e-09)			0.0002*** (3.871e-05)	
Employment income			-0.017*** (0.002)			-175.9818*** (26.0425)
Employment income ^2			3.49e-08*** (1.09e-08)			0.0003*** (0.0001)

Investment income			-0.006*** (0.002)			-149.2478*** (38.7093)
Investment income ^2			7.59e-09*** (2.77e-09)			0.0002*** (4.07e-05)
Net self-employment income			-0.016*** (0.003)			-386.7006*** (85.8200)
Net self-employment income ^2			1.91e-08*** (4.95e-09)			0.0004*** (0.0001)
Unemployment insurance income			-0.002 (0.008)			-278.9664*** (78.5345)
Social assistance income			-0.035*** (0.010)			-475.9386 (318.8319)
Intercept	0.026 (0.014)	0.019 (0.014)	0.013 (0.014)	513.6913*** (106.7107)	386.1202*** (107.3861)	328.8556*** (106.4518)
Number of observations	65,207	65,207	65,207	48,596	48,596	48,596

**Table 5 – Extensions:** Coefficients are estimated by OLS. Robust standard deviations are shown in parentheses. Income variables are divided by 10 000. Terms for marital status and age are included, but not presented above. \*\*\*, \*\*, \* denote statistical significances at the 1, 5 et 10 percent levels, respectively.

#### 5.3.3 Heterogeneity by demographic groups

We split the sample using nine separate subsamples: men, women, Atlantic, Quebec, Ontario, Prairies, BC, anglophones, and francophones. The estimates of the model with interactions for each of these groups help to determine whether a difference in financial literacy level has the same impact on the marginal effect of the METR, despite the distinct characteristics of the demographic groups. The results for each subsample are presented in appendix 3. For the sake of conciseness, only the third specification of the model is estimated (with detailed controls for income).

First, we see that the estimated coefficient of the term METR · FL for men is very similar to the one yielded for the entire sample, while the marginal effect among women is significantly lower. Also, for men with low financial literacy, the coefficient obtained for the METR is nearly twice the coefficient estimated for the entire sample (see Table 3). Conversely, the marginal effect of the METR on withdrawals for women with low financial literacy is not statistically significant. A plausible explanation for these results is that there may be a certain level of heterogeneity within the financial literacy groups. For example, the average financial literacy score for men in the high financial literacy group is statistically higher than the average score for women in the same group. Different savings behaviours may be observed when comparing individuals with one or two points of difference in the questionnaire, but who are in the same financial literacy group. Also, a larger

proportion of women do not make any RRSP withdrawals before retirement, which would partly explain why the estimated coefficients for this subsample are closer to zero compared with men.

In fact, this was the finding in the studies by Mawani and Paquette (2011) and Messacar (2017). Nevertheless, the difference in the marginal effect of the METR on RRSP withdrawals between the financial literacy groups for women is still statistically significant at the 10% threshold, which follows the same trend as the results reported throughout this chapter.

Next, the third and fourth columns of Table 6 show the estimates for the subsamples comprising individuals from each of the five main regions of Canada. The estimated coefficient for the term METR · FL is statistically significant at the 5% confidence level forOntario and the Prairies only. For four out of five regions, we cannot reject the null hypothesis that the marginal effect of the METR on withdrawals for the low financial literacy group is zero.

Lastly, the last two columns in Table 6 give the results for individuals whose first language is English and French respectively. We obtain estimates that are very similar to those obtained for the full sample with anglophones. However, the model estimates that the difference in marginal effect of the METR on withdrawals between the two financial literacy groups is not statistically significant for francophones. As previously mentioned, these results may in part be due to some level of heterogeneity in savings behaviours within the financial literacy groups. It is also likely that these results are attributable to systematic differences in savings behaviour between Quebec, the province of residence of the majority of the country's francophones, and the rest of Canada. It has been established in the literature that the savings behaviours of Quebeckers are considerably different from the rest of Canada, including the study by Lalime and Michaud (2014). For example, this study finds that Quebeckers save considerably less than the rest of Canada considering their income level, and are less likely to participate in the stock and bond markets. In the likely event that this same trend applies to our sample regardless of financial literacy level, this would mean that a larger proportion of this group does not have an RRSP or even much money in other types of savings accounts to fall back on in the event of financial difficulties.

### **5.3.4 RRSP contributions**

While this paper focuses exclusively on savings withdrawals, a similar argument for how individuals should respond to changes in their marginal effective tax rates in terms of savings

contributions also holds. In particular, one would expect contributions to *increase* as the tax rate increases since the benefit of deferring the tax rises, and this response may also depend at least in part on the level of financial literacy.

There are several reasons why we focus exclusively on withdrawals. First, contributions have received considerably more attention in the related literature. For example, Milligan (2002) estimates how RRSPs respond to changes in tax rates in the Canadian context. Second, if savers are behavioural, then decisions about whether and how much to contribute versus withdraw are not symmetric problems. Decisions about contributing may be affected by such issues as inertia, procrastination, passive saving, or bounded rationality. In contrast, pre-retirement withdrawals may be driven by problems of present-bias and temptation. While financial literacy may affect all of these behavioural issues, focusing on withdrawals on its own is warranted. This informs how to help people keep their savings intact until retirement *after* overcoming the initial hurdle of making a contribution.

Nevertheless, we briefly consider how our main results look using RRSP contributions rather than withdrawals, as shown in Table 6. We estimate our fully interacted model on RRSP contributions to test the difference in the effect of METR on contributions between the two financial literacy groups. As expected, the response to the marginal effective tax rate is generally positive and it is much larger for those with high financial literacy. This suggests financial literacy is a relevant determinant of how much people save both in terms of initially contributing and keeping their savings intact.

Fully Interacted Regression with Fixed Effects on RRSP Contributions

	(1)	(2)	(3)
	29.3331***	52.6108***	16.9657**
METR * FL	(4.3888)	(8.9895)	(6.7637)
METR	31.4404*** (2.3810)	-14.1471*** ( 3.7202 )	0.6388 (2.9375)
Income variables * FL			
Total income * FL		-417.2161*** (86.3543)	

Total income ^2 * FL		0.0005*** (0.0001)	
Employment income * FL			5.3562 (64.9908)
Employment income ^2 * FL			-0.0004 (0.0003)
Investment income * FL			-194.9509** (95.0946)
Investment income^2 * FL			0.0001 (0.0001)
Net self employment inc * FL			303.8558** (118.1408)
Net self employment inc ^2 * FL			-0.0006** (0.0003)
Employment insurance inc * FL			87.1575 (142.3148)
Social assistance inc * FL			1.4044 (187.2592)
Income variables			
Total income		611.2459*** (54.7077)	
Total income ^2		-0.0006*** (0.0001)	
Employment income			343.0697*** (38.8210)
Employment income ^2			0.0003 (0.0003)
Investment income			223.0832*** (72.2446)
Investment income^2			-0.0002** (0.0001)
Net self employment inc			343.2754*** (83.2831)
Net self employment inc ^2			-0.0003*** (0.0001)
Employment insurance inc			245.2167*** (81.4836)
Social assistance inc			179.2742*** (50.6880)
Intercept	-464.53845*** (162.6079)	55.1705 (159.6232)	169.72837 (142.4089)
Nb of observations	65,047	65,047	65,047

**Table 6 – Fully interacted regression with fixed effects on RRSP contributions**: Coefficients are estimated by OLS. Robust standard deviations are shown in parentheses. The METR variable is expressed as a percentage multiplied by 100. Income variables are divided by 10 000. Total income refers to all sources of income, from which RRSP withdrawals are subtracted. Investment income is the combined dividend and interest income. Interaction terms of marital status and age and all non-interacted variables are not presented. \*\*\*, \*\*\*, \*\* denote statistical significances at the 1, 5 et 10 percent levels, respectively.

#### 5.3.5 Subjective financial literacy

In addition to measuring Canadians' *objective* financial literacy using a 14-question survey, the dataset also contains information about respondents' *subjective* (or self-reported) financial literacy. As a result, it is possible to compare how our results by financial literacy differ using this self-reported measure. While many respondents likely have a good sense of their true financial literacy so the two measures are closely related, it is also possible for the two measures to differ. Over-confidence with savings and investment decisions can have costly long-term implications for financial well-being in retirement.

In Table 7, we report estimates of how savings withdrawals respond to changes in the effective marginal tax rate separately for individuals with low and high objective and subjective financial literacy. Two findings are apparent. First, the sample sizes for the regressions with low (high) objective literacy but high (low) subjective literacy are large, indicating there are many people who either over- or under-estimate their true knowledge. This may to some extent occur because our high objective literacy indicator is based on a comparison to the median, rather than relative to some absolute threshold as one might use for assessing their own knowledge.

Second, we see that savings responses are much larger among those with high objective financial literacy *irrespective* of their subjective literacy. Similarly, those with low objective literacy underrespond to a change in the tax rate irrespective of their subjective literacy. These finding suggest that self-reported literacy scores do not adequately measure a person's actual financial capacity.

Fixed Effect Regression on RRSP Withdrawals Comparing Objective and Subjective Financial Literacy					
	Low Objective Financial Literacy		High Objective Financial Literacy		
<del>-</del>	Low subjective FL	High subjective FL	Low subjective FL	High subjective FL	
METR	-5.9266** (2.4678)	-6.6695* (3.6806)	-23.5871*** (4.8725)	-19.2432*** (4.6392)	
Nb of observations	21,834	13,847	16,733	12,150	

Table 7 – Fully interacted regression on RRSP withdrawals comparing objective and subjective financial literacy: Coefficients are estimated by OLS. Robust standard deviations are shown in parentheses. The METR variable is expressed as a percentage multiplied by 100. Controls for income, marital status and age are not presented. \*\*\*, \*\*, \* denote statistical significances at the 1, 5 et 10 percent levels, respectively.

#### **5.4 Limitations**

In this section, we briefly discuss some of the limitations that our methodology places on the explanatory power of our models.

First, our database does not contain information enabling us to determine whether individuals have an RRSP with enough money to be able to make a large withdrawal.<sup>17</sup> Therefore, it is expected that many taxpayers considered in our estimates bias the parameters towards zero, which is true for both financial literacy groups. This is a concern if individuals with no RRSPs account for a larger proportion of individuals in the low financial literacy group than in the high financial literacy group so the magnitude of bias is larger when literacy is low.

Second, the fact that the explanatory variable of interest in our model, the METR, is an estimate can potentially limit the robustness of our results. The tax database does not include information about spouses, so net family income cannot be calculated. This variable can determine the amount received from a number of tax credits when the individual's income puts them within the clawback area for those credits. In other words, the METR estimates may contain measurement errors. That said, in the likely event that those measurement errors are not more prevalent in either of the two financial literacy groups, we assume that the bias caused by those errors would have a minimal impact on our results.

\_

<sup>&</sup>lt;sup>17</sup> The CFCS database contains information on the total combined assets in the entire family's RRSPs. No information is provided about individuals. Also, since the CFCS is a cross-sectional survey that was conducted in 2014 and the panel data date back to 2000, the information collected by the survey may not be consistent with the individual's reality for all the years observed.

# 6. Conclusion

In this study, we examined the effect of the METR on RRSP withdrawals and compared how this effect varies for tax filers with high versus low financial literacy. We first showed that most taxpayers recognize tax costs when they make decisions about how much to withdraw. The estimated relationship between RRSP withdrawals and the METR is negative and statistically significant for both financial literacy groups, on average. Second, using a fully interactive linear regression model with fixed effects, we find that individuals with high financial literacy are more responsive to changes in their METRs when making withdrawal decisions. This finding suggests financial literacy helps improve savers' understanding of the tax system and the effect of the METR on their net-of-tax benefit of withdrawals.

In addition, our analysis indicates that there is a significant connection between withdrawals and income shocks, and some demographic shocks. We find that, on average, income drops increase the amount withdrawn from RRSPs regardless of the METR in the current year. Death of a spouse and divorce have significant positive marginal effects on withdrawals for those with low financial literacy. The estimated effects of demographic and income shocks on withdrawals are often greater for the low financial literacy group, suggesting that this group may have less savings in other types of accounts that they can use in the event of financial difficulty. That said, our results show that the RRSP is used as a contingency fund before retirement by both financial literacy groups, which is consistent with Mawani and Paquette (2011).

This study paves the way for future research into the impact of financial literacy on individuals' responses to various tax incentives put in place by governments. There is a variety of government programs that aim to encourage behaviours other than retirement saving (such as saving for a child's education or making charitable donations), so exploring whether the effectiveness of those programs depend on financial literacy is a promising topic for future research.

Amid Canada's aging population, the importance of deepening our understanding of retirement savings behaviours is increasingly relevant. This is also true in many industrialized countries with similar demographic trajectories. The RRSP is a savings vehicle governed by complex rules, and contributors face many decisions that can be critical to maintaining their standard of living in

retirement. It is important for effective policy-making to identify factors that help savers make better decisions and to take these factors into account when designing new policies. The findings of this study have implications for public policy insofar as they indicate efforts to increase financial literacy among the population can help many taxfilers improve their financial situations during working years and into retirement, most notably for those with low incomes.

# **Appendix 1: Optimal Timing of RRSP Withdrawals**

Consider a model over two t periods: the first one during the individual's working life and the second one when they retire. We're looking at an individual who had previously contributed to an RRSP and is considering making a withdrawal. They receive an after-tax income  $y_1$  at t=1 and a pension benefit  $\psi y_1$  at t=2, where  $\psi$  represents the income replacement rate during their working life. The marginal effective tax rates are at  $\tau_1$  during the first period and  $\tau_2$  during the second period. At t=1, the individual can choose to allocate an amount S to a TEE-type savings account to finance their consumption at t=2. The rest of their income at t=1 is allocated to consumption  $c_1$ . The model includes a factor  $\beta$  that describes the individual's patience and depends on their preferences. We also introduce an endogenous variable  $\phi \in [0.1]$ , which is the portion of the retirement savings that the individual can choose to withdraw from their RRSP at t=1, knowing that the remaining portion  $(1-\phi)$  must be withdrawn at t=2. In our context, we are especially interested in the optimal value of  $\phi$  because that is what we will be able to use to identify the variables that affect the timing of the withdrawal.

The individual's problem is solved by maximizing the amount of the utility from their consumption at t = 1 and the expected utility from their consumption at t = 2. We use a utility function that is concave up in relation to consumption. For simplicity, we use a logarithmic utility function:

$$u(c) = \ln(c)$$

Our maximization problem is represented as:

$$max_{s,\phi} u(c_1) + \beta u(c_2)$$

s.c. 
$$c_1 \le y_1 + (1 - \tau_1)\phi - S$$
 (1)

$$c_2 \le \psi y_1 + (1+r)[(1-\tau_2)(1-\phi) + S]$$
 (2)

$$\phi \in [0,1] \tag{3}$$

The Kuhn-Tucker conditions for this problem are the following:

$$(1 - \tau_1)u'_1 - \beta(1+r)(1-\tau_2)u'_2 + \lambda_1 - \lambda_2 = 0 \tag{4}$$

$$-u'_1 + \beta(1+r)u'_2 = 0 \tag{5}$$

$$\lambda_1 \phi = 0 \tag{6}$$

$$\lambda_2(\phi - 1) = 0 \tag{7}$$

where  $\lambda_1$  and  $\lambda_2$  are the Kuhn-Tucker multipliers for  $\phi \ge 0$  and  $\phi \le 1$  respectively, and  $u'_1$  and  $u'_2$  represent marginal utility of consumption for each period. We can derive solution  $\phi$  simply by substituting the Euler equation (5) in equation (4). This manipulation gives us the following equation:

$$\beta(1+r)(\tau_2 - \tau_1)u'_2 + \lambda_1 - \lambda_2 = 0 \tag{8}$$

Equation (8) tells us that the solution for  $\phi$  is a corner solution, in other words that  $\phi = 0$  or  $\phi = 1$ . In fact, the Kuhn-Tucker conditions (6) and (7) require  $\lambda_1$  and  $\lambda_2$  to be greater than or equal to 0. Also, the variables  $\lambda$  cannot both be greater than 0 because, in that case, it would require  $\phi$  to be equal to both 0 and 1 in order to meet conditions (6) and (7). The case where  $\lambda_1 = 0$  and  $\lambda_2 = 0$  is the only exception to the corner solution; however, that implies that  $\tau_1 = \tau_2$  for equation (8) holds, which is highly unlikely in our context. If  $\tau_1 < \tau_2$ , this implies that  $\tau_2 - \tau_1 > 0$  and so necessarily that  $\lambda_1 - \lambda_2 < 0$  given equation (8). In this case, Kuhn-Tucker conditions (6) and (7) enable us to infer that  $\lambda_1 = 0$  and  $\lambda_2 > 0$ , and therefore that  $\phi = 1$ . The reverse logic applies if  $\tau_1 > \tau_2$ , and it then follows that  $\phi = 0$ . This means that the individual withdraws all the savings in the RRSP either during the first period ( $\phi = 1$ ) or the second period ( $\phi = 0$ ).

# **Appendix 2: Percentage of Correct Answers to the CFCS Questions Used to Measure Financial Literacy**

## Percentage of Correct Answers by Financial Literacy Group

Questions	Topic	Low FL	High FL
1	Inflation, interest rate	52.6%	90.2%
2	Credit report	28.0%	68.8%
3	Stock market	11.7%	48.6%
4	Unit pricing	64.0%	87.6%
5	Life insurance	69.1%	93.0%
6	Savings account	42.8%	81.2%
7	Inflation	41.9%	80.6%
8	Types of savings account	46.2%	85.6%
9	Investment, inflation	29.7%	65.3%
10	Loans	12.4%	49.5%
11	Debit cards	61.9%	92.9%
12	Credit rating	86.2%	99.6%
13	Loans, interest	62.1%	88.7%
14	Mortgages	82.4%	98.8%

**Table A1 – Percentage of correct answers to CFCS questions by financial literacy groups**: The low and high financial literacy groups respectively scored 9/14 or less and 10/14 or more. The exact list of questions can be consulted under the "Objective personal assessment" section of the following link:

 $https://www23.statcan.gc.ca/imdb/p3Instr.pl?Function=assembleInstr\&a=1\&\&lang=en\&Item\_Id=201522.$ 

# **Appendix 3: Extension on Demographic Groups**

Extensions - Demographic Groups									
	Males	Females	Atlantic	Quebec	Ontario	Prairies	ВС	Native Language - English	Native Language - French
METR * FL	-16.865**	-8.868*	-10.6594	-11.931	-22.9043**	-20.4027**	-7.6337	-16.481***	-8.953
	(6.935)	( 5.024)	(9.4713)	(10.253)	(9.3532)	(7.9285)	(7.5190)	(4.837)	(8.983)
METR	-13.639***	-3.104	-9.5952***	-7.033	0.6046	-5.0776	-3.5387	-7.247***	-9.588**
	(4.343)	(2.336)	(3.6785)	(4.996)	(5.0528)	(4.4843)	(4.6067)	(2.534)	(4.834)
Income variables * FL									
Employment income * FL	79.514***	5.360	-81.3525	52.353	91.3397	115.9461**	161.1263*	65.683**	48.188
	(29.456)	(54.805)	(63.3522)	(64.395)	(74.5257)	(51.1485)	(83.4080)	(32.435)	(61.757)
Employment income ^2 * FL	-1.5e-04**	2.03e-04	3.97e-05	6.43e-05	-1.61e-05	-0.0002	-0.0008	-1.3e-04	2.1e-04
	(6.87e-05)	(4.42e-04)	(0.0002)	(1.45e-04)	(0.0002)	(0.0002)	(0.0005)	(8.98e-05)	(1.71e-04)
Investment income * FL	115.552*	-24.179	-12.3902	440.160*	165.1471	41.4341	209.5098	-4.003	520.817***
	(59.464)	(39.779)	(77.7505)	(237.919)	(115.5784)	(43.3528)	(182.3445)	(31.876)	(188.511)
Investment income ^2 * FL	-1.45e-04*	9.1e-06	1.7e-05	-0.001*	-0.0006**	-3.55e-05	-0.0013	6.56e-06	-0.002***
	(8.2e-05)	(3.54e-05)	(0.0001)	(0.001)	(0.0003)	(4.29e-05)	(0.0011)	(3.58e-05)	(0.001)
Net self employment income * FL	122.677	115.251	-117.3361	55.149	260.6708	109.0274	683.9343*	45.198	310.801
	(97.350)	(124.997)	(129.8850)	(77.209)	(224.6430)	(172.3070)	(397.5223)	(93.221)	(308.753)
Net self employment income ^2 * FL	-3.4e-05	-0.001	0.0004	0.001**	0.0016	0.0002	-0.0023*	1.65e-04	-0.002
	(1.29e-04)	(4.51e-04)	(0.0003)	(0.001)	(0.0039)	(0.0009)	(0.0012)	(1.69e-04)	(0.003)
Unemployment insurance income * FL	206.300	91.452	21.2099	244.108	561.1483*	-177.5452	177.2305	97.189	275.785
	(163.001)	(112.035)	(161.1776)	(177.628)	(298.9979)	(173.1702)	(359.2106)	(118.086)	(169.796)
Social assistance income * FL	-179.133	-106.676	41.4476	-1123.2078	-259.1887	-55.1100	-274.3101	2.456	-747.064
	(193.824)	(126.561)	(278.6828)	(1266.7269)	(172.8632)	(179.9444)	(864.9244)	(112.637)	(666.534)

Income variables									
Employment income	-125.157***	-138.877***	-75.1250*	-201.835***	182.8057***	165.9284***	-224.3422***	-117.718***	-211.846***
	(25.766)	(45.987)	(40.2294)	(45.649)	(66.5635)	(46.3195)	(73.9559)	(29.185)	(37.815)
Employment income ^2	1.636e-04**	2.19e-05	0.0002	3.33e-04***	0.0002	0.0003	0.0009*	1.44e-04	2.54e-04***
	(6.85e-05)	(4.36e-04)	(0.0002)	(8.09e-05)	(0.0002)	(0.0002)	(0.0005)	(8.97e-05)	(9.15e-05)
Investment income	-153.173***	-65.393***	-62.8010	-510.359**	334.7831***	-70.1353*	-290.0339**	-41.799	-550.468***
	(55.828)	(23.837)	(57.4364)	(233.633)	(97.1440)	(36.4606)	(142.9104)	(24.544)	(186.423)
Investment income ^2	2.02e-04***	6.39e-05***	0.0001	0.001**	0.0008***	0.0001*	0.0016	4.78e-05	0.002***
	(7.74e-05)	(2.21e-05)	(0.0001)	(8.84e-04)	(0.0003)	(3.64e-05)	(0.0010)	(2.65e-05)	( 0.001)
Net self employment income	-268.402***	-320.494***	-104.6935	-178.112***	-432.7126**	-297.5818*	-819.9957**	-227.766***	-469.861
	(89.269)	(110.854)	(85.6064)	(62.442)	(218.3925)	(157.5554)	(393.1576)	(80.849)	(304.761)
Net self employment income ^2	3.09e-04***	0.001***	0.0001	-0.001**	-0.0013	0.0002	0.0026**	2.61e-04***	0.002
	(1.01e-04)	(4.01e-04)	(0.0001)	(6.35e-04)	(0.0039)	(0.0009)	(0.0012)	(8.92e-05)	(0.003)
Unemployment insurance income	-177.964	-152.752**	-115.4254	-325.540***	-152.3663	-161.9802	-167.1711	-104.205	-292.005***
	(96.848)	(62.214)	(81.0969)	(90.054)	(209.5986)	(128.2994)	(220.7557)	(74.743)	(82.573)
Social assistance income	-492.332***	-175.187***	360.8930***	-327.022***	-244.0772*	378.2419***	311.4174	-321.939***	-375.034***
	(84.993)	(81.219)	(96.0101)	(111.059)	(138.0542)	(92.2786)	(820.6494)	(65.439)	(121.070)
Intercept	595.418***	-78.999	399.9148***	-55.774	146.2971	333.1686**	34.3727	294.715**	60.001
	(130.796)	(94.61)	(146.9948)	(130.017)	(216.1748)	(144.3264)	(250.1262)	(95.66)	(118.511)

**Table A2** – Extension on demographic groups: Coefficients are estimated by OLS. Robust standard deviations are shown in parentheses. Income variables are divided by 10 000. Terms for marital status and age are included, but not presented above. \*\*\*, \*\*, \* denote statistical significances at the 1, 5 et 10 percent levels, respectively.

17,020

11,852

11,855

17,797

6,155

44,192

14,856

Number of observations

30,971

34,152

# References

Blancquaert, A., N.-J. Clavet, J.-Y. Duclos, B. Fortin et S. Marchand (2017). « Taux marginaux effectifs d'imposition : une comparaison Québec Ontario », *L'Actualité économique*, vol. 93, no 4, p. 531–558.

Boisclair, D., A. Lusardi et P.C. Michaud (2017). « Financial literacy and retirement planning in Canada », *Journal of Pension Economics and Finance*, vol. 16, no 3, p. 277–296.

Boyer, M., P. d'Astous et P.C. Michaud (forthcoming). « Tax-Sheltered Retirement Accounts: Can Financial Education Improve Asset Location Decisions? », *Review of Economics and Statistics*.

Chan, S., et A. H. Stevens (2008). « What You Don't Know Can't Help You: Pension Knowledge and Retirement Decision-Making », *Review of Economics and Statistics*, vol. 90, no 2, p. 253–66.

Clark, R., A. Lusardi, et O. S. Mitchell (2017). « Financial knowledge and 401(k) investment performance: a case study », *Journal of Pension Economics and Finance*, vol. 16, no 3, p. 324–347.

Duclos, J., B. Fortin et A. Fournier (2008). « Une analyse des taux marginaux effectifs d'imposition au Québec », *L'Actualité économique*, vol. 84, no 1, p. 5–46.

Feldman, N., P. Katuščák et L. Kawano (2016). « Taxpayer Confusion: Evidence from the Child Tax Credit », *American Economic Review*, vol. 106, no. 3, p. 807-35.

Gideon, M. (2017). « Do Individuals Perceive Income Tax Rates Correctly? », *Public Finance Review*, vol. 45, no 1, p. 97–117.

Giles, P. et K. Maser (2004). « Using RRSPs Before Retirement », *Perspectives on Labour and Income*, vol. 5, no 12, p. 14-22.

Godbout, L. et M. Angers (2012). « L'imposition implicite des revenus additionnels : comment se compare le Québec », Working Paper no. 2012-04, Chaire en fiscalité et en finances publiques.

Grenier, É. (2019): « Inattention rationnelle et décisions d'épargne-retraite » Masters Thesis, HEC Montréal.

Hastings, J. S., O. S. Mitchell, and E. Chyn. (2011). « Fees, Framing, and Financial Literacy in the Choice of Pension Manager » dans *Financial Literacy: Implications for Retirement Security and the Financial Marketplace*, dir. O. S. Mitchell et A. Lusardi, Oxford, Oxford University Press, p. 101-115.

Lalime, T. et P.C. Michaud (2014). « Littératie financière et préparation à la retraite au Québec et dans le reste du Canada », *L'Actualité économique*, vol. 90, no 1, p. 23-45.

Laurin, A. et F. Poschmann (2010). « Saver's Choice : Comparing the Marginal Effective Tax Burens on RRSPs and TFSAs, Working Paper, C.D. Howe Institute.

Lusardi, A. et O. Mitchell (2007). « Financial Literacy and Retirement Preparedness: Evidence and Implications for Financial Education », *Business Economics*, vol. 42, no 1, p. 35–44.

Lusardi, A. et O. Mitchell (2011). « Financial literacy and planning: implications for retirement wellbeing », in *Financial Literacy: Implications for Retirement Security and the Financial Marketplace*, dir. A. Lusardi and O. Mitchell, Oxford, Oxford University Press, p. 17–39.

Lusardi, A. et O. Mitchell (2014). «The Economic Importance of Financial Literacy», *Journal of Economic Literature*, vol. 52, no 1, p. 5-44.

Marchand, S. (2018). « Who Benefits from Tax-Preferred Savings Accounts? », Working Paper 18-12, Chaire de recherche sur les enjeux économiques intergénérationnels.

Mawani, A. et S. Paquette (2011). « Pre-Retirement RRSP Withdrawals », *Canadian Tax Journal*, vol. 59, no 2, p. 183-219.

Messacar, D. (2017): « Trends in RRSP Contributions and Pre-retirement Withdrawals, 2000 to 2013 », *Economic Insights*, no. 064, p. 1-9.

Miller, B. et K. Mumford (2015). « The Salience of Complex Tax Changes: Evidence from the Child and Dependent Care Credit Expansion », *National Tax Journal*, vol. 68, no 3, p. 477–510.

Milligan, K. (2002). "Tax-preferred savings accounts and marginal tax rates: Evidence on RRSP participation." *Canadian Journal of Economics* 35 (3): 436-456.

Milligan, K. (2016). *Canadian Tax and Credit Simulator* (CTaCS). Accessed September 2019 from https://faculty.arts.ubc.ca/kmilligan/ctacs/.

Mullock, K. et J. Turcotte (2012). « Financial Literacy and Retirement Saving », Working Paper, Finance Department.

Rees-Jones, A. et D. Taubinsky (2019). « Measuring "Schmeduling" », *Review of Economic Studies*, vol. 87, no 5, p. 2399-2438.

Shillington, R. (2003). « New Poverty Traps: Means-Testing and Modest-Income Seniors», *CD Howe Institute Backgrounder*, n. 65, p.1-9.

van Rooij, M., A. Lusardi et R. Alessie (2012). « Financial Literacy, Retirement Planning and Household Wealth », *The Economic Journal*, vol. 122, no 560, p. 449–478.

Veall, M. R. (2014). « Estimates of the Number of Guaranteed Income Supplement Recipients Who Receive Income from Registered Retirement Savings Plans », *Canadian Tax Journal*, vol. 62, no 2, p. 383–399.

Yoo, K.-Y. et A. de Serres (2004). « Tax Treatment of Private Pension Savings in OECD Countries and the Net Tax Cost Per Unit of Contribution to Tax-Favoured Schemes », Working Paper, OCDE Economics Department.