

2002s-08

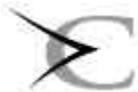
# **The Rate of Risk Aversion May Be Lower Than You Think**

*Kris Jacobs*

---

**Série Scientifique**  
*Scientific Series*

---



**CIRANO**  
Centre interuniversitaire de recherche  
en analyse des organisations

Montréal  
Janvier 2002

## CIRANO

Le CIRANO est un organisme sans but lucratif constitué en vertu de la Loi des compagnies du Québec. Le financement de son infrastructure et de ses activités de recherche provient des cotisations de ses organisations-membres, d'une subvention d'infrastructure du ministère de la Recherche, de la Science et de la Technologie, de même que des subventions et mandats obtenus par ses équipes de recherche.

*CIRANO is a private non-profit organization incorporated under the Québec Companies Act. Its infrastructure and research activities are funded through fees paid by member organizations, an infrastructure grant from the Ministère de la Recherche, de la Science et de la Technologie, and grants and research mandates obtained by its research teams.*

### Les organisations-partenaires / The Partner Organizations

- École des Hautes Études Commerciales
- École Polytechnique de Montréal
- Université Concordia
- Université de Montréal
- Université du Québec à Montréal
- Université Laval
- Université McGill
- Ministère des Finances du Québec
- MRST
- Alcan inc.
- AXA Canada
- Banque du Canada
- Banque Laurentienne du Canada
- Banque Nationale du Canada
- Banque Royale du Canada
- Bell Canada
- Bombardier
- Bourse de Montréal
- Développement des ressources humaines Canada (DRHC)
- Fédération des caisses Desjardins du Québec
- Hydro-Québec
- Industrie Canada
- Pratt & Whitney Canada Inc.
- Raymond Chabot Grant Thornton
- Ville de Montréal

© 2002 Kris Jacobs. Tous droits réservés. *All rights reserved.* Reproduction partielle permise avec citation du document source, incluant la notice ©.

*Short sections may be quoted without explicit permission, if full credit, including © notice, is given to the source.*

Les cahiers de la série scientifique (CS) visent à rendre accessibles des résultats de recherche effectuée au CIRANO afin de susciter échanges et commentaires. Ces cahiers sont écrits dans le style des publications scientifiques. Les idées et les opinions émises sont sous l'unique responsabilité des auteurs et ne représentent pas nécessairement les positions du CIRANO ou de ses partenaires.

*This paper presents research carried out at CIRANO and aims at encouraging discussion and comment. The observations and viewpoints expressed are the sole responsibility of the authors. They do not necessarily represent positions of CIRANO or its partners.*

# The Rate of Risk Aversion May Be Lower Than You Think\*

*Kris Jacobs*<sup>†</sup>

## Résumé / Abstract

A l'aide de données sur la consommation des ménages et d'équations d'Euler, cet article estime le taux d'aversion au risque relatif. Ces équations d'Euler sont les implications de structures de marché qui ne permettent pas toujours aux agents de s'assurer parfaitement. Cet article porte plus particulièrement sur des tests de l'équation d'Euler inconditionnelle. Dans le cadre d'un agent représentatif, ce type de test mène aux rejets les plus intuitivement convaincants des modèles d'évaluation d'actifs, comme l'énigme de la prime de risque et l'énigme du taux sans risque. Lorsque l'on ignore les erreurs de mesure de la consommation, les erreurs de l'équation d'Euler ne sont pas statistiquement différents de zéro pour les valeurs du taux d'aversion au risque relatif comprises entre 1 et 3. Lorsque l'on tient compte de la présence des erreurs de mesure, les estimations conservatrices du taux d'aversion au risque relatif pour les participants à un marché d'actifs indiquent une valeur entre 2 et 8. Ces résultats suggèrent que le taux d'aversion au risque pourrait être plus bas que ce qui est couramment perçu. Par conséquent, l'imperfection des marchés pourrait servir à résoudre les énigmes d'évaluation d'actifs.

This paper estimates the rate of relative risk aversion using Euler equations based on household-level consumption data. These Euler equations are implications of market structures that do not necessarily allow agents to perfectly insure themselves. The paper focuses on tests of the unconditional Euler equation. In representative-agent frameworks, this type of test leads to the most intuitively convincing rejections of asset-pricing models, such as the equity premium puzzle and the riskfree rate puzzle. When measurement error in consumption is ignored, Euler equation errors are not statistically different from zero for values of the rate of relative risk aversion between 1 and 3. When allowing for the presence of measurement error, conservative estimates of the rate of risk aversion for asset market participants indicate a value between 2 and 8. These findings suggest that the rate of risk aversion may be much lower than commonly thought. Consequently, market incompleteness is likely to be part of a resolution of asset pricing puzzles.

**Mots clés :** Aversion au risque, Énigme de la prime de risque, Erreurs de mesure, Cohortes synthétiques, Équations d'Euler, Participation au marché des actifs

**Keywords:** Risk Aversion; Equity Premium Puzzle; Measurement Error; Synthetic Cohorts; Euler Equations; Asset Market Participation

**JEL :** G12

---

\* Faculty of Management, McGill University and CIRANO. I would like to thank FCAR of Québec and SSHRC of Canada for financial support.

<sup>†</sup> Correspondence to: Kris Jacobs, Faculty of Management, McGill University, 1001 Sherbrooke Street West, Montreal, Canada H3A 1G5; Tel: (514) 398-4025; Fax: (514) 398-3876; E-mail: jacobs@management.mcgill.ca.

# 1 Introduction

Asset pricing puzzles such as the equity premium puzzle and the riskfree rate puzzle illustrate the failure of the representative-agent consumption-based asset pricing models of Lucas (1978) and Breeden (1979) to explain the data. The reason for the widespread interest in these puzzles is probably that they summarize the failure of the model in a way that is intuitively appealing, by referring to mean returns or the first two moments of returns (e.g. see Mehra and Prescott (1985) and Hansen and Jagannathan (1991)). This failure of the model is easier to grasp intuitively than a statistical rejection.

The focus of this paper is on the importance of incomplete markets structures, and therefore on the measurement of the pricing kernel, for the interpretation of these asset pricing puzzles. In the literature on incomplete markets, the focus has traditionally been on asset-pricing puzzles and less on statistical tests of the model. Early papers (for instance Telmer (1993) and Lucas (1994)) conclude, using a simulation framework, that the impact of incomplete markets on the equity premium puzzle is minimal, because through self-insurance agents end up with consumption allocations that are very similar to those in the complete markets case. Later work by Heaton and Lucas (1996), Constantinides and Duffie (1996), Telmer, Storesletten and Yaron (1997) and Constantinides, Donaldson and Mehra (1998) points out that the impact of incomplete markets depends on a number of issues, such as the persistence of idiosyncratic shocks, the relation of idiosyncratic shocks to aggregate shocks, the net supply of bonds and the specification of market frictions. The consensus from this literature, which is mostly based on simulation, seems to be that we do not yet have a model that can explain the stylized facts on asset returns, but market incompleteness brings model predictions more in line with the data and therefore holds promise to explain asset pricing puzzles, possibly by incorporating additional rigidities or alternative utility functions.

In a closely related literature, papers by Jacobs (1999), Brav, Constantinides and Geczy (1999), Cogley (1999) and Vissing-Jorgensen (1999) provide additional insight into incomplete markets structures by testing the incomplete markets framework using data on household consumption. While these tests are of great interest for asset pricing, their interpretation is sometimes problematic because of data limitations and because simple implications of dynamic asset markets such as Euler equations can in principle be tested in many different ways. This paper contributes to this literature by investigating the importance of these two potential problems (data limitations and the type of testing technique), which are to some extent interrelated. First, consider the type of test. Jacobs (1999) uses a standard econometric technique to conclude that certain implications of incomplete market models are statistically rejected by the data. Whereas these statistical tests are instructive, they suffer from the same problem as the Hansen and Singleton (1982) tests in the representative agent

literature, because they are intuitively less appealing than the formulation of the model’s performance in terms of asset pricing puzzles. This paper attempts to provide additional insight by using household consumption data to construct tests that can be more readily related to the equity premium puzzle and the risk free rate puzzle.

Second, one of the critical problems with the interpretation of tests of Euler equations that use household data is the impact of measurement error in consumption. One of the ways in which such measurement error can be addressed is by loglinearizing the Euler equation, as in Runkle (1991) and Vissing-Jorgensen (1999). This paper investigates a different, more direct approach to the measurement error problem. Instead of using individual household data to construct the tests, I use data for a consumer that is “representative” for a certain group, such as an age group. For several plausible types of measurement error, this approach will reduce the consequences of the presence of measurement error.

I find that it is important to take measurement error into account. However, even after allowing for it, the data still indicate moderate degrees of risk aversion. This finding contrasts with findings in the representative agent literature. Studies of representative agent models that exploit the unconditional Euler equation such as Mehra and Prescott (1985) and Hansen and Jagannathan (1991) typically find very large values of the rate of relative risk aversion. The paper therefore concludes that the incomplete markets framework holds great promise to explain the equity premium puzzle and the riskfree rate puzzle. The paper proceeds as follows. Section 2 briefly discusses the evidence on asset pricing puzzles in representative agent models. Section 3 discusses the existing evidence on incomplete markets, outlines the testing methodology and discusses the strategy to deal with measurement error. Section 4 presents the empirical results. Section 5 concludes by discussing the results and outlining future research.

## 2 Asset Pricing Puzzles in Representative Agent Models

This paper investigates a set of joint restrictions on security returns and consumption. Consider an agent with a time-separable constant relative risk aversion (TS-CRRA) utility function. If the agent maximizes lifetime utility, if she has a riskless and a risky asset at her disposal, and if she does not end up at a corner solution, the following intertemporal Euler equations will hold

$$1 = \beta E_t \left( \frac{c_{i,t+1}}{c_{i,t}} \right)^{\gamma-1} R_{rl,t+1} \tag{1}$$

$$1 = \beta E_t \left( \frac{c_{i,t+1}}{c_{i,t}} \right)^{\gamma-1} R_{ri,t+1} \tag{2}$$

where  $c_{i,t}$  is the consumption of agent  $i$  in period  $t$ ,  $R_{r,t,t+1}$  and  $R_{ri,t,t+1}$  are the riskless and the risky return between periods  $t$  and  $t + 1$ ,  $\beta$  is the discount factor,  $E_t$  is the mathematical expectation conditional on information available at  $t$ , and  $1 - \gamma$  is the rate of relative risk aversion. The theory underlying (1) and (2), coupled with a complete markets assumption, is the basis for much of modern asset pricing theory, and consequently there is substantial interest in testing whether these intertemporal restrictions are supported by the data. If a complete set of markets is available, then individual marginal utility is perfectly correlated with the marginal utility of the representative agent, who consumes aggregate consumption, and one can test the underlying theory by testing (1) and (2), with aggregate consumption taking the place of individual household consumption.

The resulting literature on tests of the representative-agent consumption-based asset-pricing model has shown that the model is at odds with the data in a number of dimensions. The early literature (Hansen and Singleton (1982,1983), Grossman, Melino and Shiller (1987)) mainly focuses on the statistical investigation of a number of restrictions implied by the model, and concludes that the model is statistically rejected by the data. Starting with Mehra and Prescott (1985), attention has shifted to well-defined puzzles, aspects of the model which are easily understood intuitively and are clearly at odds with the data. Even though a considerable research effort has unearthed a large number of these puzzles, the ones that are most often cited are the equity premium puzzle and the riskfree rate puzzle. The equity premium puzzle, originally pointed out by Mehra and Prescott (1985), states that to explain the difference in the return of the risky and the riskless asset, we need risk aversion of the representative agent to be higher than 10, a number which Mehra and Prescott believe to be implausibly high. The riskfree rate puzzle, first pointed out by Weil (1989), observes that if risk aversion is that high, the discount factor  $\beta$  required to match the historically observed riskfree rate is larger than one.

It must be noted that the equity premium puzzle and the riskfree rate puzzle are not entirely without controversy. To understand this, note that as Kocherlakota (1996) points out, these puzzles are quantitative but not qualitative puzzles. The comovements between the variables of interest in the data conform to the theory, but to match the magnitudes observed in the data, the unobservable parameters of interest  $\beta$  and  $1 - \gamma$  have to assume values that are considered implausible. A potential reaction to the puzzles is therefore that judging behavioral parameters as implausible, presumably based on introspection, is not very convincing. In this context, note that Kandel and Stambaugh (1991) argue that fairly high levels of risk aversion may not be implausible, while Kocherlakota (1990) has demonstrated that values of  $\beta$  larger than one can be supported in equilibrium. Whereas this argument is interesting, this paper takes it as given that a large number of economists regard the values of the behavioral parameters as implausible.

These intuitively appealing summaries of the failure of the model have given rise to a very extensive literature. A large part of this literature has focused on changes to the TS-CRRA utility specification that is used in Mehra and Prescott (1985). As pointed out in a number

of studies (e.g. see Hansen and Jagannathan (1991)), the finding that risk aversion has to be high can be attributed to the fact that the marginal utility of the representative agent is not volatile enough. Consequently, a solution to the puzzle is to make marginal utility more volatile by changing the utility function. Kocherlakota (1996) argues that proposed utility specifications such as habit formation (see Abel (1990), Campbell and Cochrane (1999), Constantinides (1990), Detemple and Zapatero (1991), Ferson and Constantinides (1991), Heaton (1995) and Sundaresan (1989)) and nonexpected utility (see Epstein and Zin (1989,1991) and Weil (1990)) are able to explain the risk free rate puzzle but not the equity premium puzzle. This paper does not discuss empirical evidence on these alternative preference structures, not because they do not have merit, but because the issue of interest is best illustrated using the popular TS-CRRA specification.

## 3 Household Consumption Data and Asset Pricing Puzzles

### 3.1 Empirical Methodology

It has long been recognized that incomplete markets structures are a potential explanation of the equity premium puzzle (e.g. see Mankiw (1986) and Mehra and Prescott (1985)). However, the analysis of incomplete markets structures has been conducted using different empirical methodologies. First, a number of studies use simulation techniques. Whereas early quantitative evidence (Aiyagari and Gertler (1991), Huggett (1993), Telmer (1993) and Lucas (1994)) indicates that incomplete markets cannot resolve asset pricing puzzles, it is now recognized that the occurrence of uninsurable risk in a model economy brings predicted asset returns more in line with the data. For instance, Telmer, Storesletten, and Yaron (1997) and Constantinides, Donaldson and Mehra (1999) achieve some degree of success in matching the data using an overlapping-generations framework.

However, the predictions of model economies with incomplete markets depend critically on a number of issues that have to be measured prior to setting up the model to explain asset returns. For instance, the precise measurement of transaction costs is an important issue (see Aiyagari and Gertler (1991), Constantinides (1986), He and Modest (1995), Heaton and Lucas (1995,1996) and Luttmer (1996)). Another important issue is the persistence of idiosyncratic shocks to income (e.g. see Constantinides and Duffie (1996) and Heaton and Lucas (1995, 1996)).<sup>1</sup> It must be noted that any future resolution of issues such as

---

<sup>1</sup>See also the discussion in Kocherlakota (1996). The issue of persistence is important because it determines the extent to which an individual can insure against income shocks. Given sufficient potential for self-insurance, the outcome of an incomplete markets economy will be similar to that of a complete markets economy (Telmer (1993)). It must be noted that a model economy with high shock persistence has implications similar to a model economy where the number of shocks is small in comparison to the individual's life

the persistence of income shocks requires a careful analysis of household-level data, and therefore simulation-based studies are subject to similar criticism about data limitations as direct statistical analyses of household-level data.

A second approach to analyzing market incompleteness focuses on restrictions implied by the model under a weaker set of assumptions, such as restrictions (1) and (2). This is the approach followed in this paper. Besides providing a valuable complement to simulation-based studies, parameter estimates obtained through this analysis are also of interest as inputs in simulations or calibrations. To motivate the empirical analysis in this paper and to see the link with existing studies, consider the following framework. Consider the errors  $u_{i,t}^1$  and  $u_{i,t}^2$  associated with the Euler equations (1) and (2), where  $t$  is the time index and  $i$  is the household index. Theory specifies  $E_{t-1}u_{i,t}^1 = E_{t-1}u_{i,t}^2 = 0$ . The equity premium puzzle and the riskfree rate puzzle can then be understood using the expressions  $E u_{i,t}^1 = E u_{i,t}^2 = 0$ , which are readily obtained using the law of iterated expectations. First consider the corresponding restrictions in the representative-agent case. In this case  $E u_t^1 = E u_t^2 = 0$ . There is no  $i$  subscript: the theory specifies only two orthogonality conditions. Therefore, in the representative-agent case, we have a just identified econometric problem: two equations in the two unknowns  $\beta$  and  $1 - \gamma$ .

The incomplete markets case we analyze in this paper can be cast in a similar setup. In this case we have a maximum of  $T$  observations on  $N$  households. In contrast to the representative agent problem, the optimization problem using household data is overidentified: we have  $2 \times N$  restrictions to determine 2 parameters. However, these  $2 \times N$  orthogonality restrictions can conveniently be reduced to 2 orthogonality conditions in an intuitively appealing way. Noting that the panel is unbalanced, denote the number of households in the sample at time  $t$  as  $N_t$ . We can simply use the two orthogonality conditions  $E \overline{u}_t^1$  and  $E \overline{u}_t^2$  where  $\overline{u}_t^j = (1/N_t) \sum_{i=1}^{N_t} u_{i,t}^j$ , which gives us two equations in two unknowns.

It is possible to provide more intuition by further relating this setup to the existing literature using a generalized method of moments (GMM) framework. See Campbell, Lo and MacKinlay (1997), Cochrane (2001) and Ferson (1995) for a more in-depth treatment. In the representative agent case one can think of the errors  $u_t^1$  and  $u_t^2$  associated with the Euler equations as pricing errors. In the incomplete markets case  $\overline{u}_t^1$  and  $\overline{u}_t^2$  are average pricing errors. The type of analysis that is directly based on the pricing errors is often referred to as the analysis of the unconditional Euler equation. Now consider the same sample with a maximum of  $T$  observations on  $N$  households and consider an instrument  $z_{i,t-1}$  (notice that for simplicity the instrument does not carry a superscript, to indicate that we can use the same instrument for both equations). Then consider  $\varepsilon_{i,t}^1 = u_{i,t}^1 z_{i,t-1}$  and  $\varepsilon_{i,t}^2 = u_{i,t}^2 z_{i,t-1}$ . Using the law of iterated expectations we know that  $E \varepsilon_{i,t}^1 = E \varepsilon_{i,t}^2 = 0$ , for all  $i$ . We therefore again have  $2 \times N$  orthogonality conditions, which can be reduced to

---

span. This partly explains the success of the overlapping generations models in Telmer, Storesletten, and Yaron (1997) and Constantinides, Donaldson and Mehra (1999).

two orthogonality conditions  $E \bar{\varepsilon}_t^1$  and  $E \bar{\varepsilon}_t^2$  where  $\bar{\varepsilon}_t^j = (1/N_t) \sum_{i=1}^{N_t} \varepsilon_{i,t}^j$ . If the instrument  $z_{i,t-1}$  depends on time (such as asset returns or consumption growth), the resulting analysis is referred to as exploiting the conditional Euler equation. Another way this is referred to in the literature is as an analysis of “scaled” pricing errors or as a payoff to a “managed portfolio” (Cochrane (2001)). In contrast, the analysis of the unconditional Euler equation described above analyzes the “unscaled” pricing errors. Another way to interpret this is as an analysis of scaled pricing errors where the “scaling” instrument  $z_{i,t-1}$  is a constant.

The GMM framework looks at two theoretical restrictions such as  $E u_t^1 = E u_t^2 = 0$  or  $E \bar{\varepsilon}_t^1 = E \bar{\varepsilon}_t^2 = 0$  as two orthogonality conditions, regardless of whether they originate from representative agent problems or incomplete markets problems and regardless of whether the instrument  $z_{i,t-1}$  is time-varying. The GMM approach exploits the fact that if the theory can explain the data, the sample equivalent of these orthogonality conditions should be zero. At its most general, consider the restrictions implied by an incomplete markets economy with a time-varying instrument  $z_{i,t-1}$ . Define the  $2 \times 1$  vector  $\bar{\varepsilon}_t = (\bar{\varepsilon}_t^1, \bar{\varepsilon}_t^2)'$  and consider the minimization problem

$$J_T = \left[ \frac{1}{T} \sum_{t=1}^T \bar{\varepsilon}_t(\phi)' \right] W \left[ \frac{1}{T} \sum_{t=1}^T \bar{\varepsilon}_t(\phi) \right]. \quad (3)$$

where  $\phi$  denotes the  $(2 \times 1)$  parameter vector and  $W$  is a weighting matrix. The equity premium puzzle can be cast in this statistical framework by observing that the corresponding minimization problem (3), when using aggregate data and when using a constant as an instrument, yields large levels of risk aversion and a value of  $\beta$  larger than one. The question addressed by this paper asks if the corresponding minimization problem using household data yields more plausible parameter values. It must be noted that to obtain consistent estimates of the parameters  $\beta$  and  $1 - \gamma$ , one can actually analyze  $\bar{\varepsilon}_t = (\bar{\varepsilon}_t^1, \bar{\varepsilon}_t^2)'$  instead of minimizing (3), because the problem is just identified. Therefore, to find the parameter values, we can solve  $\bar{\varepsilon}_t = (\bar{\varepsilon}_t^1, \bar{\varepsilon}_t^2)'$  using a nonlinear solution algorithm, and we do not necessarily have to analyze (3). A valid solution to the optimization problem (3) is one for which  $\bar{\varepsilon}_t = (\bar{\varepsilon}_t^1, \bar{\varepsilon}_t^2)'$  is exactly zero.<sup>2</sup> However, for analyses of the conditional Euler equation that use more instruments  $z_{i,t-1}$ , this is not the case. For instance, if we consider the Euler equations (1) and (2) and use two lagged instruments for each Euler equation, we end up with four orthogonality conditions to determine the two parameters. The problem is then overidentified and can only be analyzed by minimizing (3). In this case the choice of the weighting matrix  $W$  is critical, while it is of no consequence for a just-identified problem. See Hansen and Singleton (1982) and Jacobs (1999) for econometric analyses of overidentified problems in the representative agent and incomplete markets cases respectively.

---

<sup>2</sup>This statement of course has to be understood to mean that the Euler equation errors are arbitrarily close to zero. The same is true for the objective function in (3).

## 3.2 Interpreting Asset Pricing Puzzles

A few more observations about the interpretation of the equity premium puzzle and the risk-free rate puzzle are in order which are most conveniently illustrated using a GMM framework. First, in the case of the equity premium puzzle, since we are dealing with a just identified GMM problem it is clear that we cannot test the theory in a formal econometric sense. We have just enough orthogonality conditions available to estimate parameters and compute standard errors. It is therefore clear that in the context of this GMM optimization problem the equity premium puzzle has to be seen as a statement about the plausibility of estimated parameters. Second, when considering the risk free rate puzzle from the perspective of some well-defined GMM problem, since we have one equation with two parameters, the problem is underidentified. So not only can we not test the theory econometrically, in principle we cannot obtain estimates of the behavioral parameters either. The caveat, clearly expressed in the literature, is of course that when discussed in this context, the risk free rate puzzle is a puzzle conditional on the existence of the equity premium puzzle (e.g. see Campbell, Lo and MacKinlay (1997), Kocherlakota (1996)). Conditional on large risk aversion, the Euler equation associated with the riskless asset implies an implausible discount rate. So in the context of the GMM problem above, once we have an estimate of  $1 - \gamma$ , we have one equation in one remaining unknown, a just identified problem.

However, the asset-pricing literature also contains a number of techniques and analyses which provide a slightly different formulation of asset-pricing puzzles. In the GMM setup described above, the Euler equations for the riskless and the risky asset yield two Euler equations in the representative agent context. The reason why we are not able to test the model statistically is that we choose to use these two orthogonality conditions to obtain inference on the two parameters. Alternatively, conditional on knowledge of the parameter values, we have two overidentifying conditions available to test the model. Even the risk-free rate puzzle becomes testable econometrically, since it becomes a problem with one overidentifying condition. Therefore, we can construct tests of these asset pricing models conditional on a wide range of parameter values. Asset pricing puzzles are then slightly restated but intuitively similar: instead of using the data to infer parameter values, and then judge whether these parameter values are plausible, we construct econometric tests of Euler equation errors conditional on a range of parameter values, and then ask ourselves if those parameter values that imply Euler errors that are statistically indistinguishable from zero are plausible. The simplest possible implementation of this methodology is illustrated in Kocherlakota (1996). Whereas that paper constructs univariate  $t$ -tests, in principle multivariate chi-square tests can be constructed. Indeed, to use multivariate information in an optimal way, GMM-type testing conditional on the behavioral parameters can be used (see Burnside (1994), Cecchetti, Lam and Mark (1994) and Hansen, Heaton and Luttmer (1995)). To some extent, this can be interpreted as an application of incorporating prior information, with the prior either resulting from introspection, or from other (related) information. Although they have a dif-

ferent focus than the analysis in Kocherlakota (1996), the techniques advocated by Hansen and Jagannathan (1991) use a similar setup. These graphical techniques focus on values of behavioral parameters needed to match the first and second moments of asset returns. For instance, using only the riskfree asset we find that for a “plausible” value of  $\beta$  of 0.95, “implausible” levels of risk aversion are needed to match the riskfree rate. This paper follows a similar approach to analyze the incomplete markets case using panel data. The tests are based on overidentified GMM problems such as (3), which lead to tests of the models using  $\chi^2$  statistics.

### 3.3 Measurement error

The estimation and analysis of (1) and (2) using data on individual household consumption assumes that the available data are measured without error. We know that for panel data this is unlikely to be the case (see Altonji (1986) and Altonji and Siow (1987)). While it is by definition impossible to remove this measurement error from the data, there are a number of alternatives available to reduce its consequences. One potential solution is to loglinearize the Euler equations, as in Runkle (1991) and Vissing-Jorgensen (1999). Such a loglinearization can be justified by relying on a Taylor series approximation or by assuming that the distribution of the data is jointly lognormal, and it can resolve measurement error problems given appropriate instrumentation. While this is an interesting approach, the disadvantage is that if the assumptions underlying the loglinearization are incorrect, the resulting test statistics and parameter estimates will be biased. Another potential technique to assess the implications of measurement error is to assume a functional form for the measurement error and to use simulation techniques to investigate its impact on the test statistics. This approach is followed by Vissing-Jorgensen (1999) and Brav, Constantinides and Geczy (1999). While this is an interesting approach, its validity is predicated on assuming the correct type of measurement error.

In this paper I propose to use synthetic cohorts to reduce the negative impact of measurement error. Synthetic cohorts have been used in the economics literature for a long time.<sup>3</sup> The approach requires the construction of cohorts, groups of individuals with comparable characteristics. For instance, one often used approach is to form cohorts based on age. For each age group, we construct a “representative” individual who has data associated with him/her that are the average of the data in that age group. It is clear that this is also an imperfect solution to the measurement error problem. If the consumption growth of one of the individuals in a certain age group is measured with error, the measurement error is not eliminated. Instead, the relative importance of the measurement error is reduced by averaging it out using other observations. Moreover, it must be noted that the formation of

---

<sup>3</sup>The traditional motivation for using the synthetic cohorts technique is somewhat different. It has mainly been used because many available panels are too short for the analysis of rational expectations models (see Browning, Deaton and Irish (1985) and Attanasio and Weber (1995)).

synthetic cohorts can also adversely affect the analysis. The central thesis of the market incompleteness literature is that the study of representative agent data is to some extent uninformative because variability in individual marginal rates of substitution is averaged out when studying the marginal rate of substitution of the representative consumer. When constructing synthetic cohorts, one runs into the same problem that besides averaging out the effects of measurement error, one also averages out genuine volatility in marginal rates of substitution. In fact, the construction of a representative agent can be seen as the extreme of synthetic cohorts construction, with just one cohort. As a result, one partially reduces the consequences of measurement error, but at the cost of introducing a new problem. The synthetic cohort approach can therefore be interpreted as finding a middle way between the problems associated with using individual data (measurement error) and the problems associated with the representative agent approach (eliminating MRS variability).

In the empirical work, I use synthetic cohorts based on relatively large groups, which presumably eliminate most of the measurement error. Because such a construction also eliminates some genuine MRS variability, the resulting parameter estimates can usefully be interpreted as conservative estimates of the rate of risk aversion. I first construct synthetic cohorts by sorting on the age of the household head. I use households where the household head is aged between 25 and 60, resulting in 36 cohorts. I provide a robustness check by also constructing synthetic cohorts using the household head's education as well as his age, using 12 age categories and 2 educational categories. This results in 24 cohorts. In principle, it is possible to conduct further robustness checks by using interesting variables such as household wealth and household income as sorting variables. The analysis in this paper is limited to the two sorts described above to keep the presentation of the evidence manageable.

There is one remaining issue regarding the cohort construction. The main variable of interest in this study is consumption growth. For each cohort group, we therefore construct an average consumption growth variable. However, this can be done in (at least) two different ways. First one can construct the cohort consumption level and then use the cohort consumption level to compute cohort consumption growth. More precisely, given individual consumption  $c_{i,t}$  the consumption of a representative cohort  $j$  at time  $t$  is given by  $cco_{j,t} = (1/N_{j,t}) \sum_{i=1}^{N_{j,t}} c_{i,t}$ , where  $N_{j,t}$  is the number of observations in this cohort at time  $t$ . Consumption growth for the cohort is then computed as  $ccog_{1,j,t} = (cco_{j,t}/cco_{j,t-1})$ . Alternatively, one can compute cohort consumption growth by averaging over individual consumption growth of the cohort members. More precisely,  $ccog_{2,j,t} = (1/N_{j,t}) \sum_{i=1}^{N_{j,t}} (cg_{i,t})$  where  $cg_{i,t} = c_{i,t}/c_{i,t-1}$ .

### 3.4 Related Literature

Besides the literature on incomplete markets referred to in section 3.1, there are a number of studies that are related to the empirical work in this paper. First, the complete markets assumption has been strongly statistically rejected in a number of recent studies (see Attana-

sio and Davis (1997), Cochrane (1991), Mace (1991), Hayashi, Altonji and Kotlikoff (1996); see Altug and Miller (1990) and Miller and Sieg (1997) for conflicting evidence). Second, Mankiw and Zeldes (1991) investigate the impact of asset market participation on the equity premium puzzle. They are motivated by aggregation results that hold only for asset market participants (see Grossman and Shiller (1982)). Therefore, they use panel data to eliminate non-assetholders from the sample and investigate if the resulting per capita consumption of assetholders differs from per capita consumption for all households. Mankiw and Zeldes (1991) also obtain inference on parameters by analyzing the unconditional Euler equation. However, because of their setup, they obtain inference on the rate of risk aversion but not on the discount rate  $\beta$ . Also, the analysis focuses on per capita (average) consumption and not on the properties of individual household consumption; however, the properties of per capita consumption are dramatically different from the properties of household consumption or cohort consumption, suggesting that the assumptions necessary for aggregation of those households at interior conditions are not validated by the data. While the use of per capita consumption based on panel data presumably avoids most of the measurement error problems associated with individual consumption, it seems desirable to investigate the dispersion in marginal rates of substitution in the cross-section.

Third, by focusing on individual household consumption, some of the results in this paper are also intimately related to certain findings in the life-cycle literature, for instance in the papers by Attanasio and Browning (1995), Attanasio and Weber (1993,1995), Hall and Mishkin (1982), Runkle (1991) and Zeldes (1989). These papers typically investigate linearizations of the conditional Euler equation associated with the riskless asset. Altug and Miller (1990) use panel data to investigate nonlinear Euler equations under the maintained hypothesis of market completeness. Finally, Jacobs (1999) investigates Euler equations (1) and (2) without making a market completeness assumption. That study focuses on overidentified estimation problems which allow statistical tests of the model. The overidentifying conditions are obtained from the conditional Euler equation. Jacobs (1999) finds that the consumption-based asset pricing model is statistically rejected by the data along several dimensions. Estimates of behavioral parameters are robust across estimation exercises, with fairly plausible inference on the discount rate as well as the rate of relative risk aversion.

Whereas the results in the life-cycle literature and in Jacobs (1999) are insightful, more evidence on asset pricing puzzles is desirable by focusing exclusively on the unconditional Euler equation. To motivate this, a reference to the empirical evidence in the representative agent literature is instructive. The results in Jacobs (1999) are obtained using panel data and an incomplete markets assumption using similar information as exploited by Hansen and Singleton (1982) in a representative agent framework. Hansen and Singleton (1982,1984) reject the model statistically and obtain estimates of the rate of relative risk aversion that are in most cases intuitively plausible.<sup>4</sup> Estimates of the discount rate are very precise and

---

<sup>4</sup>However, some of the estimates of the rate of relative risk aversion in Hansen and Singleton (1982, 1984)

in most cases intuitively plausible. However, the work of Hansen and Singleton (1983), Mehra and Prescott (1985), Grossman, Melino and Shiller (1987) shows that if only the unconditional Euler equation is used in the representative agent framework, inference on the rate of relative risk aversion changes dramatically. Indeed, by using the unconditional Euler equation one obtains the equity premium puzzle and the implied risk aversion is very high. This finding is confirmed using the tools developed by Hansen and Jagannathan (1991). So obviously the type of information used is critical in representative agent studies, and this could very well be the case when using household level data. The focus of this paper is therefore to verify if the intuitively plausible estimates of the rate of relative risk aversion obtained in Jacobs (1999) are confirmed when only unconditional information is used and when measurement error is taken into account. As explained above, this is of substantial interest because implicitly or explicitly, the validity of asset pricing theories has been judged by these parameter values. It must be noted that the studies by Brav, Constantinides and Geczy (1999), Cogley (1999) and Vissing-Jorgensen (1999) give partial answers to some of these questions. These studies seem to indicate that measurement error has potentially serious implications and that risk aversion is lower than in representative agent studies, although the estimates differ considerably between papers.

## 4 Empirical Results

### 4.1 Data

The data used in the analysis are consumption data from the Panel Study of Income Dynamics (PSID) that are also used in Jacobs (1999). I refer the reader to Jacobs (1999) for a detailed discussion of the data; here I provide a brief summary. The data are taken from the 1987 PSID files and allow construction of consumption growth rates for the period 1974-1985. The empirical analysis is repeated for four different samples. The first sample consists of all households who fulfill a set of sample selection criteria. This dataset contains 3555 households and 18813 total observations (note that the panel is unbalanced: a household can be in the sample one year and not in another year). This sample is referred to as sample 1, but a problem with this sample is that some of the households do not hold assets. Including such households in the sample is likely to bias the results. I therefore use three smaller samples that contain only assetholders. A second sample (referred to as sample 2) includes only households that report nonzero holdings of the riskless and the risky asset. It contains 740 households and 5029 total observations. A third and fourth sample (referred to as samples 3 and 4) contain only households that report holdings of the riskless and risky assets larger than \$1,000 and \$10,000, respectively. These datasets contain 413 and 103 households and

---

are in the nonconcave region of the parameter space and the estimates are imprecise for several estimation exercises.

2990 and 769 total observations, respectively. Table I provides descriptive statistics for these samples (in panel A) as well as for the samples synthetic cohorts based on them (in panels B and C).<sup>5</sup> Table IA indicates that sample 1 (the one containing all households) contains more extreme outliers compared to the other samples. However, this may be partly due to the fact that it is the largest sample. Another observation is that the cohort construction significantly alters the properties of the different datasets, as can be seen from a comparison of table IA with tables IB and IC.

## 4.2 Asset Pricing Puzzles Under Incomplete Markets

We obtain inference on the equity premium in incomplete markets for the four samples described above by analyzing  $\bar{u}_t = (\bar{u}_t^1, \bar{u}_t^2)'$ , which is a system of two equations in two unknowns, for each of these samples. Intuitively this analysis indicates whether we can find values of the parameters  $\beta$  and  $1 - \gamma$  that can explain the average (across households) Euler equation pricing errors for the riskless and the risky asset. However, there is an additional empirical problem with household level consumption data in this context which has not yet been discussed. Whereas the Euler equations (1) and (2) are defined for an individual agent, consumption data are available at the household level. In principle this issue is easily addressed by including a function of household size in the Euler equation. We can estimate

$$1 = \beta E_t \left( \frac{c_{i,t+1}}{c_{i,t}} \right)^{\gamma-1} R_{r,t+1} \exp(f_1 f s_{i,t+1} + f_2 f s_{i,t}) \quad (4)$$

$$1 = \beta E_t \left( \frac{c_{i,t+1}}{c_{i,t}} \right)^{\gamma-1} R_{ri,t+1} \exp(f_1 f s_{i,t+1} + f_2 f s_{i,t}) \quad (5)$$

where  $f s_{i,t}$  stands for family size in period  $t$  and  $f_1$  and  $f_2$  are scalar parameters. However, this means four parameters to estimate and the Euler equations for the riskless and the risky asset give us only two orthogonality conditions, when using a constant as the sole instrument. The problem is underidentified. We can solve this problem in two ways. First, we can include current and lagged household size in the instrument set, to yield four equations in four unknowns. The approach I follow instead is to plug in estimates of  $f_1$  and  $f_2$  that are obtained from the estimation of overidentified or just identified estimation problems implied by (4) and (5). This results in a just-identified problem of two equations in two unknowns. It was verified that estimates of  $f_1$  and  $f_2$  are robust across estimation exercises, and do not influence the empirical results.

Results on the equity premium puzzle are presented using a graphical tool that is in the same spirit as the tables in Kocherlakota (1996). Instead of conducting an econometric

---

<sup>5</sup>Panels B and C indicate that some cohorts do not contain observations in some cases.

analysis and evaluating the parameters using the results of econometric estimation, we evaluate the orthogonality restrictions for a number of combinations of the parameter values, and subsequently ask if any of the parameter values for which the theory is not rejected is intuitively plausible. To facilitate the interpretation of the results on incomplete markets obtained with this alternative technique, figure 1 analyzes the Euler equation errors for the more familiar representative agent case, using the Mehra and Prescott (1985) annual data for 1889-1979. Figure 1C summarizes the information in the sample equivalents to the Euler equations in the way it is going to be used in the rest of the paper. This figure summarizes the information contained in figures 1A and 1B. To save space the counterparts to figures 1A and 1B are omitted in the figures that analyze household data.

Figure 1A presents a grid where the inputs are the discount rate and the rate of relative risk aversion. For the discount rate  $\beta$  values between 0.5 and 1.1 are analyzed. For  $\gamma - 1$  (the negative of the rate of relative risk aversion) we limit our attention to values between -50 (in the concave region of the parameter space) and +20 (in the nonconcave region of the parameter space). The vertical axis represents the negative of the absolute value of the average Euler equation error for the riskless asset. Therefore, this grid reaches a maximum value of zero for parameter combinations that solve the Euler equation for the riskless asset. We can therefore conduct the following informal exercise: pick an a priori reasonable value of  $\beta$  and verify what level of risk aversion is necessary to solve the Euler equation. Many economists would argue that values of beta between 0.8 and 1 are intuitively plausible. Given these values and the data, a solution to the Euler equation is given for relative risk aversion of approximately 25 to 35, which many economists consider implausibly high. Figure 1B presents the same analysis for the Euler equation associated with the risky asset. Again, given plausible discount rates we need very high rates of risk aversion to solve the Euler equation. Note that for a given plausible discount rate, the risk aversion needed to match the riskless rate is slightly lower than the one needed to match the risky rate.

Figure 1C presents the information contained in figures 1A and 1B in a more parsimonious way. This figure has the discount rate on the horizontal axis and risk aversion on the vertical axis. The triangles and squares present combinations of the parameters that make the Euler equations for the risky and the riskless asset respectively equal to zero.<sup>6</sup> For all values of  $\beta$  smaller than one, the squares are above the circles. In other words, it is not possible to find a parameter combination that solves both Euler equations as long as the discount rate

---

<sup>6</sup>Because figures 1A and 1B are in fact created using a grid, the “ridge” in these figures that suggests zero Euler equation errors is in fact a locus of very small Euler equation errors. It was verified that actual solutions to the equations exist in the neighbourhood of the parameter values in the grid. Presenting a finer grid does not provide any additional insight. With respect to figure 1C, note that the “solutions” pulled of the grid are therefore “near” solutions. This figure could be made more precise by presenting a locus of exact solutions, but again this does not provide any additional insight. After all, whether a relative risk aversion of 30.32 or 30.36 is needed to match the restrictions of the model is not likely to influence our assessment of the model’s performance.

is smaller than one. When  $\beta$  is allowed to be larger than one, a solution to the system of equations can be found, namely for a discount rate approximately equal to 1.08 and a rate of relative risk aversion approximately equal to 23. As extensively discussed above, these values are judged to be implausible. What is perhaps not clearly indicated in the representative agent literature is the fact that, for values of the discount rate only slightly higher than 1.08, it is impossible to find a rate of risk aversion that solves the Euler equations for the risky and riskless asset separately (let alone jointly).<sup>7</sup>

We now turn to similar graphical analyses of the restrictions imposed by the incomplete markets models on the data for individual households. Figure 2 presents information for the four individual household samples similar to that in figure 1C. Figure 2A reports on the sample of households that report nonzero holdings of the riskless and risky assets, and figures 2B and 2C report on those households with holdings of the riskless and risky asset in excess of \$1,000 and \$10,000 respectively. Finally, figure 2D presents results for the complete sample of 18813 observations.<sup>8</sup>

Figure 2 clearly indicates that parameter combinations needed to solve each Euler equation separately are much more plausible than in the representative agent model. Most importantly, compared to the representative agent model, the data point towards lower risk aversion. For most plausible values of the discount rate, the rate of relative risk aversion implied by the data is between 0.5 and 3. With respect to discount rates, whereas in the representative agent case solutions to separate Euler equations do not exist for discount rates larger than 1.1, for the household level data we cannot obtain solutions to the Euler equation for the risky asset for  $\beta$  larger than 0.9, and to the Euler equations for the riskless asset for

---

<sup>7</sup>In a sense, the representation of asset pricing puzzles in figure 1C is superior to simply reporting the results of an econometric estimation exercise. The figure has an advantage similar to an often used representation of the Hansen-Jagannathan (1991) bounds, because in essence the test information is repeated for many different parameter combinations, and not exclusively for the parameter combination that minimizes a given objective function. In some sense, the information in figure 1C is also superior to Hansen-Jagannathan (1991) bounds, because those bounds only use the first two moments of the IMRS. However, all these perceived advantages over other testing methods have to be put in perspective. The representation of asset pricing puzzles in figure 1C is convenient because in the case of a CRRA specification we are dealing with a two-parameter problem. Tools such as the Hansen-Jagannathan bounds and GMM estimation are designed to deal with richer parameterizations of the utility function (see Heaton (1995) and Cochrane and Hansen (1992)). Therefore, whereas the graphical tool used in figure 1C may be useful in the context of this paper, it is of limited use more generally.

<sup>8</sup>To provide some additional insight into the properties of the four samples, table II provides descriptive statistics on a year-by-year basis. This complements the descriptive statistics in table IA, which are for the entire sample. Table II indicates that there is quite a bit of variation in the risky rate of return over the sample period. When comparing the four different samples, average consumption growth rates on a year-by-year basis are fairly similar. The minima and maxima for consumption growth are more extreme for the sample that contains all households, but it has to be kept in mind that this sample is larger, and therefore more likely to contain extremely high and low growth rates of consumption. In terms of measurement error, it must be noted that there seem to be very few extreme outliers, even though no consumption growth criterion was used to screen observations.

$\beta$  larger than 0.95.

It is somewhat surprising that for the sample that includes non-asset holders (figure 2D) implied risk aversion is not much different from figures 2A, 2B and 2C. Studies by Constantinides, Brav and Geczy (1999), Vissing-Jorgensen (1999), and Mankiw and Zeldes (1991) suggest that the distinction between asset holders and non-asset holders is important. The evidence presented in figure 2 suggests that the issue of aggregation is much more important than the inclusion of non-asset holders in the sample. If the presence of non-asset holders in the sample is an important issue, we would expect to see an implied rate of risk aversion in figure 2D that is more similar to the ones implied by representative agent models, and therefore much higher than the ones in figures 2A, 2B and 2C.

There is another important stylized fact in figure 2 which requires elaboration. While there is a large number of intuitively plausible combinations of  $\beta$  and  $1 - \gamma$  for which each of the Euler equations can be solved, there is no combination of  $\beta$  and  $1 - \gamma$  that solves both Euler equations simultaneously. In the figures, such a parameter combination would be represented by an intersection of the lines formed by the triangles and the squares. For instance, in the representative agent case in figure 1C there is such a case, but only for an implausible parameter combination with  $\beta$  larger than 1 and a large value for the rate of relative risk aversion.<sup>9</sup>

What does this finding mean? Asset pricing puzzles in the representative agent model are often summarized by the finding in figure 1C that is summarized above: a “calibration” of the model using both Euler equations yields an exact solution to the system of equations for a discount rate approximately equal to 1.08 and a rate of relative risk aversion approximately equal to 23. Of course, the puzzle is much deeper. Figure 1 clearly indicates that for almost all reasonable discount rates, the rate of risk aversion implied by either Euler equation is very large. The relevant question is not whether the particular values “calibrated” on the system of two equations in two unknowns are plausible, but whether there are any parameter combinations for which both Euler equation errors are statistically indistinguishable from zero. This is the approach followed in most of the literature (see for instance Kocherlakota (1996) and also (in a different context) Hansen and Jagannathan (1991) and Cochrane and Hansen (1992)). We therefore turn to a similar analysis for incomplete markets involving the PSID household-level data.

---

<sup>9</sup>As mentioned above, figure 2 represents systems of two equations in two unknowns. In principle therefore, to find the parameter values implied by the data, we can solve  $\bar{\varepsilon}_t = (\bar{\varepsilon}_t^1, \bar{\varepsilon}_t^2)'$  using a nonlinear solution algorithm. Alternatively, we can analyze (3) which should have a minimum of zero if there is an exact solution. Figure 2 indicates that whereas (3) has a minimum, at this minimum  $\bar{\varepsilon}_t = (\bar{\varepsilon}_t^1, \bar{\varepsilon}_t^2)'$  is not zero. This finding is corroborated by an extensive grid search and a search using a nonlinear solution algorithm, which indicates that there are no exact solutions.

### 4.3 Statistically significant Euler equation errors

Figure 3 extends the graphical representation of the Euler equation errors in figure 2 to illustrate all parameter combinations for which the Euler equations are not statistically different from zero. Figure 3A presents results for sample 2, figure 3B for sample 3, figure 3C for sample 4, and figure 3D for sample 1. Each figure contains two solid lines and two broken lines. The area between the two solid lines is the area for which the Euler equation for the riskless asset is statistically indistinguishable from zero.<sup>10</sup> The area between the broken lines indicates the area for which the Euler equation for the risky asset is statistically indistinguishable from zero. Therefore, any overlap between these two areas can be interpreted as a set of parameter combinations for which the equity premium is resolved.<sup>11</sup>

Each of the figures contains such an overlap between the statistically significant solution for the riskless asset and the risky asset. In all samples, the number of parameter combinations that solves the Euler equation is much larger for the risky asset. As expected, the areas that solve the Euler equations are largest in figure 3C, the figure for the smallest sample, because this sample contains the smallest amount of information.<sup>12</sup> Once again it is the case that figure 3D, which also includes non-assetholders, looks very similar to figures 3A, 3B and 3C.

For the sake of comparison, figure 1D provides a similar analysis for the representative agent framework, using the Mehra-Prescott (1985) data. Note that this figure contains only 1 solid line and 1 broken line. For a given discount rate, if the null hypothesis of a zero Euler equation error cannot be rejected for a certain rate of relative risk aversion, it can also not be rejected for any larger rate of relative risk aversion (this was verified only for rates of risk aversion smaller than 70). Therefore, figure 1D indicates that even though there is only one parameter combination that solves the system of two equations in two unknowns (see figure 1C), there are many parameter combinations for which the hypothesis that either Euler equation is zero cannot be rejected (tested separately). However, these parameter combinations always involve a rate of relative risk aversion which is considered implausibly high. The equity premium puzzle is very robust in representative agent models because, even

---

<sup>10</sup>As explained above, the statistical tests are based on the GMM framework. The orthogonality condition  $E u_t^1$  can be interpreted as 1 equation in two unknowns,  $\beta$  and  $1 - \gamma$ , and is therefore underidentified. However, conditional on given values of  $\beta$  and  $1 - \gamma$ , we have 1 overidentifying condition, which can be used to construct a  $\chi_1^2$  statistic. To construct this statistic, it is assumed that the Euler equation errors for a household are uncorrelated over time. The correlation between different households at the same point in time is determined by the size of the aggregate shock hitting the economy at that time, and therefore it is allowed to be nonzero.

<sup>11</sup>Strictly speaking, one should construct a joint test for the hypothesis that both Euler equation errors are nonzero. I present univariate tests because they facilitate the graphical presentation of the results.

<sup>12</sup>This of course also illustrates the conventional problem that test statistics that are statistically not distinguishable from zero may result from a lack of sufficient data and therefore from a lack of power of the test.

though we can find a wide range of parameter values for which the Euler equation errors are statistically indistinguishable from zero (not just the one parameter combination indicated by GMM), all these parameter combinations are judged implausible.

#### 4.4 Analysis of Cohort Data

Figures 4 through 7 present an analysis of the Euler equation errors for cohort data. Figures 4 and 5 present results for cohorts formed on the basis of age and education. Figures 6 and 7 present results for cohorts formed on the basis of age. Figures 4 and 6 use cohort consumption growth that is obtained by first obtaining the consumption levels of a representative cohort, and subsequently computing the cohort's consumption growth. Figures 5 and 7 obtain cohort consumption growth by averaging over individual consumption growth. Each figure contains four panels, corresponding to samples 2, 3, 4 and 1. To further motivate the analysis of the cohort data, it is instructive to compare the descriptive statistics for the cohort data in tables III and IV with the descriptive statistics for the individual data in table II. It is clear that there are important differences in the properties of the consumption growth data. Regardless of the method used to construct cohort consumption growth, the cohort consumption growth data contain far fewer outliers than the individual data in table II. A few other stylized facts should be mentioned. First, average growth rates are always higher when computed by averaging over individual growth rates as opposed to averaging over the consumption levels. Second, because the cohort statistics in table IV are obtained using more cohorts, the distribution of consumption growth is more spread out than in table III. These stylized facts are also clear from comparing tables IB and IC.

Inspection of figures 4 through 7 yields three important general conclusions. First, implied rates of risk aversion are higher than for the individual data in figure 2, but dramatically lower than for the representative agent data in figure 1. Second, the implied rate of risk aversion is lower for assetholders than for non-assetholders. In panel D, the rate of risk aversion is implausibly large for plausible values of the rate of time preference. I therefore conclude that the inclusion of non-assetholders likely biases estimates of the behavioral parameters. This finding is in accordance with existing findings in Mankiw and Zeldes (1991), Constantinides, Brav and Geczy (1991) and Vissing-Jorgensen (1999). This conclusion differs from the findings in figure 2, which were obtained using individual data. In figure 2, different samples indicate comparable rates of risk aversion, regardless of whether non-assetholders were included in the sample or not. It is therefore reasonable to conclude that there is substantial measurement error in individual consumption data, and that this measurement error is (partially) eliminated by the representative cohort construction.<sup>13</sup>

---

<sup>13</sup>There is one important caveat in interpreting these figures, which is the differences in the sizes of the cohorts used to generate the figures. Because the samples get smaller as we consider households that hold larger asset portfolios (see tables I and II), the average cohorts used to generate panel C in figures 4 through 7 are always smaller than those used to generate panel A in figures 4 through 7. To investigate the impact

Third, different samples of assetholders imply different rates of risk aversion. In each of the four figures, we see that panel A has the highest rates of risk aversion, with risk aversion getting smaller as we consider households that hold larger amounts of assets.<sup>14</sup> Again, this finding is in accordance with existing research (see Mankiw and Zeldes (1991), Constantinides, Brav and Geczy (1991) and Vissing-Jorgensen (1999)). What estimate of risk aversion should we then remember from these figures? Not all rates of time preference investigated in these pictures are intuitively plausible: many economists would argue that  $\beta$  should be at least 0.85 or even 0.9. For those values of  $\beta$ , the figures indicate a rate of risk aversion of no higher than 8 and in several cases as low as 2.

## 4.5 Analysis of Representative Agent Data

Figure 8 relates the analysis of this paper to Mankiw and Zeldes (1991). Mankiw and Zeldes (1991) were the first to emphasize the importance of asset market participation for the equity premium puzzle and for estimates of the rate of risk aversion. They also use PSID data but instead of using individual data or cohort data, they create a representative consumer for different samples of consumers based on assetholdings. The representative consumer in their study is formed by first computing the average consumption level, and subsequently computing consumption growth. They find that the implied rate of risk aversion is smaller when considering assetholders and decreases when considering consumers that hold larger amounts of assets. While the analysis used in Figure 8 is slightly different from Mankiw and Zeldes' approach<sup>15</sup>, the empirical results are quite similar.<sup>16</sup> When considering samples of assetholders, the implied rate of risk aversion is smaller compared to samples that contain assetholders as well as non-assetholders. This is especially the case for figure 8C as compared

---

of this data issue on the results, the analysis was repeated for the larger samples by repeatedly drawing subsamples with a number of observations equal to that of the smaller samples. By using these subsamples to generate similar figures, I investigated how much of the difference in results between different panels is due to differences in cohort size. On average, 10 to 15 % of the differences in risk aversion found are due to these sample size effects, while the remaining 85 to 90% is due to genuine differences in the characteristics of the different types of assetholders.

<sup>14</sup>The empirical results in panels A, B and C for figures 4 through 7 indicate that cohort construction influences estimates of the rate of risk aversion. This is not surprising given the differences in the descriptive statistics evidenced in tables III and IV. First, when averaging over consumption growth (in figures 5 and 7), we obtain larger estimates of risk aversion compared to when averaging over the consumption level (in figures 4 and 6). Also, risk aversion is a bit higher when using cohorts based on age and education (in figures 4 and 5) compared to the cohorts based on age in figures 6 and 7. Altogether, the findings in panels A, B and C are surprisingly robust to cohort construction.

<sup>15</sup>Mankiw and Zeldes (1991) use a local normality argument based on the continuous time model. As a result, the only determinant of the equity premium is the covariance between consumption growth and the risky return. In this study all moments of the distribution are taken into consideration.

<sup>16</sup>See table V for descriptive statistics on the representative agent data.

to figure 8D, and to a lesser extent for figure 8B as compared to figure 8D. Note also that the use of panel data in itself does not lead to lower estimates of the rate of risk aversion. The estimates of risk aversion in figure 8D are in some cases higher than in Figure 1. This finding is also similar to the one in Mankiw and Zeldes (1991).

## 4.6 Conditional Information

Figures 9, 10 and 11 discuss the role of conditional information. Whereas this is not the main focus of this paper, the graphical analysis conducted here provides some interesting perspectives on this issue. In the representative agent literature, the analysis of the unconditional Euler equation points towards large risk aversion but econometric analysis of the conditional Euler equation indicates small risk aversion (see Hansen and Singleton (1982)).<sup>17</sup> However, for the individual household data the evidence seems much more robust, because the unconditional Euler equation analyzed here points towards intuitively plausible parameter values that are very similar to those obtained in Jacobs (1999).

To resolve this issue, figures 9, 10 and 11 further investigate the importance of conditional information for asset pricing in a perhaps unconventional fashion. These figures investigate orthogonality conditions two at a time, where each pair is formed by using the Euler equation for the riskless and risky asset and one single instrument. This allows us in principle to determine both parameters of interest. So we proceed in exactly the same fashion as with the unconditional information. Figure 9 presents these results for the individual data, figure 10 for the cohort data and figure 11 for the representative agent data used by Mehra and Prescott (1985). Figures 9 and 10 both use same sample 3 (asset holdings > \$1000) and the same instruments: lagged rates of return in panels A and B, and lagged unemployment rates in panel C. The unemployment rates are better instruments than lagged consumption because they perform better in the presence of measurement error. Panel D uses the lagged unemployment rates interacted with the age of the household head. For the representative agent data in figure 11, figure 11A uses the lagged risky return as an instrument, figure 11B the lagged riskless return, figure 11C lagged consumption growth and figure 11D consumption growth lagged twice. Inspection of these figures makes it quite clear that the impact of conditional information is minimal. The corresponding orthogonality conditions are solved for parameter combinations that are very similar to the case of unconditional information.

---

<sup>17</sup>The relationship between the estimates of the rate of risk aversion and the type of information used in the test was pointed out by Hansen and Singleton (1983) and Grossman, Melino and Shiller (1987). The issue is addressed in more detail in Singleton (1990), Ferson and Harvey (1992,1993), Ferson and Constantinides (1991) and Cochrane (2001).

## 4.7 Identification Problems

Finally, figure 12 presents more evidence that is useful to reconcile the results in this paper with existing results in the literature. Specifically, the results in figures 2 and 9 indicate that individual household data always imply fairly low values of the rate of relative risk aversion. When inspecting figures 1 and 11, it seems that representative agent data always lead to high estimates of the rate of risk aversion. However, several studies in the representative agent literature starting with Hansen and Singleton (1982) have obtained low estimates of the rate of risk aversion when using conditional information.

This finding is due to an identification problem that is illustrated in figure 12. Figure 12A is a more complete representation of the information in figure 1C. It indicates that for many plausible values of  $\beta$  (for instance  $\beta = 0.85$ ), the Euler equation is solved by a rate of risk aversion in the concave region of the parameter space but also by another rate of relative risk aversion in the nonconcave region of the parameter space. The solution in the nonconcave region of the parameter space is not represented in figure 1C. However, figure 12A also illustrates that for very large values of the discount rate, we can have two solutions to each Euler equation that are in the concave region of the parameter space, the implausibly large one that is presented in figure 1C and another value that is much smaller. Note that these smaller risk aversion values do not exactly solve the equity premium puzzle (the curve formed by joining the plus signs stays above the curve formed by joining the stars). However, those solutions ensure that the representative agent literature does not have an “equity return puzzle”, which would exist using exclusively the evidence in figure 1C. In figure 12A, we can find discount rates smaller than 1 and low rates of risk aversion that explain the return on the risky asset by itself.

Moreover, figure 12A also indicates that if we do not take into account the equity premium puzzle, the risk free rate puzzle does not exist. There are some combinations of discount rates less than one and low risk aversion that explain the riskfree rate. This is also noted for instance by Kocherlakota (1996), in his discussion of table 2 in that paper. However, the number of parameter combinations that solves the riskfree rate puzzle is much smaller than those that explain the rate of return on equity. What is perhaps not well articulated in the representative agent literature is that these parameter values that solve individual Euler equations are in a loose sense “closer” in the parameter space to parameter combinations in the nonconcave region of the parameter space (squares and triangles) than to those parameter values that are more commonly associated with asset pricing puzzles (plus signs and stars).

Figure 12B illustrates that when incorporating conditional information, the evidence looks very similar to figure 12A. For many plausible values of the rate of time preference, there are two rates of risk aversion in the concave region that solve the Euler equation. This is not the case when considering individual household data (or cohort data). As in the representative agent case, there is a range of discount rates for which there are two rates of risk aversion that solve the Euler equation. However, using household data it is always the

case that one rate of risk aversion is in the concave region of the parameter space and the other is not. Therefore, since the nonconcave region of the parameter space is thought of as entirely implausible, the nonconcave solution is of no interest for econometric inference.<sup>18</sup> For the representative agent case, it is possible that certain combinations of instruments indicate a large rate of risk aversion and other combinations of instruments indicate a small rate of risk aversion, as is the case in Hansen and Singleton (1982).<sup>19</sup>

## 5 Conclusion

This paper demonstrates that the incomplete markets setting holds great promise to resolve a number of important issues in asset pricing, such as the equity premium puzzle and the riskfree rate puzzle. Even though no parameter combination can be found for which the Euler equation error for the riskless asset and the Euler equation error for the risky asset are both exactly zero, there exist a number of intuitively plausible combinations of the discount rate and the rate of relative risk aversion for which Euler equation errors are not statistically different from zero. In the representative agent literature, we can also find parameter combinations that yield Euler equation errors that are statistically indistinguishable from zero. However, these parameter combinations all imply high rates of relative risk aversion, and are therefore judged implausible.

One of the potential problems with tests of incomplete markets is the presence of substantial measurement error. This study investigates if the findings on the rate of risk aversion are robust to the measurement error problem by repeating the empirical tests for synthetic cohorts. While the rate of risk aversion implied by the Euler equations is affected, indicating that measurement error is important, the level of risk aversion is still intuitively plausible for assetholders.

These findings complement a number of findings in the existing literature on incomplete markets that are typically obtained using simulation techniques. The conclusion from this literature is that the existence of uninsurable risk can bring predicted asset returns more in line with the data, but often at the cost of employing a model economy which may be at odds with other stylized facts. For instance, the degree of persistence in shocks to

---

<sup>18</sup>Some authors advocate to exclude the nonconcave region of the parameter space from consideration (see Braun, Constantinides and Ferson (1993)).

<sup>19</sup>This discussion is meant to illustrate an identification problem which is not always elaborated upon in the literature. It does not claim to provide intuition for the underlying economic determinants of this identification problem. Ferson and Constantinides (1991) provide an intuitive explanation of the conflicting evidence on risk aversion in representative agent models, by associating conditional information with intertemporal substitution and the high rate of risk aversion based on unconditional information with the agent's attitude towards risk.

individual income needed in these models may be at odds with the persistence inferred from the dataset used in this study, the PSID. This debate is far from closed and it is likely that additional modifications to the model economy will be found that bring the model's predictions even more in line with the data. However, the strength of the evidence contained in this paper is that few assumptions have to be made in order to analyze the Euler equations for individuals. The finding that the Euler equation errors are statistically indistinguishable from zero is therefore a robust indication that the presence of uninsurable idiosyncratic risk will be an important component of any future resolution of asset pricing puzzles. These findings indicate that we have to come up with models that do not allow agents to self-insure, leading to large differences in the marginal rate of substitution ratios across consumers. It must also be noted that the findings in this paper are very useful for the simulation-based literature (and for several other areas in economics and finance) in a different respect, because we obtain inference on behavioral parameters that can be used as inputs in those studies.

## References

- [1] Abel, Andrew B., 1990, Asset prices under habit formation and catching up with the Joneses, *American Economic Review Papers and Proceedings* 80, 38-42.
- [2] Aiyagari, S. Rao and Mark Gertler, 1991, Asset returns with transaction costs and uninsured individual risk, *Journal of Monetary Economics* 27, 311-331.
- [3] Altonji, Joseph G., 1986, Intertemporal substitution in labor supply: evidence from micro data, *Journal of Political Economy* 94, S176-S215.
- [4] Altonji, Joseph G. and Aloysius Siow, 1987, Testing the response of consumption to income changes with (noisy) panel data, *Quarterly Journal of Economics* 102, 293-328.
- [5] Altug, Sumru and Robert A. Miller, 1990, Household choices in equilibrium, *Econometrica* 58, 543-570.
- [6] Attanasio, Orazio P. and Martin Browning, 1995, Consumption over the life cycle and over the business cycle, *American Economic Review* 85, 1118-1137.
- [7] Attanasio, Orazio P. and Steven J. Davis, 1997, Relative wage movements and the distribution of consumption, *Journal of Political Economy* 104, 1227-1262.
- [8] Attanasio, Orazio P. and Guglielmo Weber, 1993, Consumption growth, the interest rate and aggregation, *Review of Economic Studies* 60, 631-649.

- [9] Attanasio, Orazio P. and Guglielmo Weber, 1995, Is consumption growth consistent with intertemporal optimization? Evidence from the consumer expenditure survey, *Journal of Political Economy* 103, 1121-1157.
- [10] Braun, Phillip, George Constantinides and Wayne Ferson, 1993, Time nonseparability in aggregate consumption, *European Economic Review* 37, 897-920.
- [11] Brav, Alon, George Constantinides and Christopher Geczy, 1999, Asset Pricing with Heterogeneous Consumers and Limited Participation: Empirical Evidence, working paper, University of Chicago.
- [12] Breeden, Douglas T., 1979, An intertemporal asset pricing model with stochastic consumption and investment opportunities, *Journal of Financial Economics* 7, 265-296.
- [13] Browning, Martin, Angus S. Deaton and Margaret Irish, 1985, A profitable approach to labor supply and commodity demands over the life-cycle, *Econometrica* 53, 503-544
- [14] Burnside, A. Craig, 1994, Hansen-Jagannathan bounds as classical tests of asset-pricing models, *Journal of Business and Economics Statistics* 12, 57-79.
- [15] Campbell, John Y. and John Cochrane, 1999, By force of habit: A consumption-based explanation of aggregate stock market behaviour, *Journal of Political Economy* 107, 205-251.
- [16] Campbell, John Y., Andrew W. Lo and A. Craig MacKinlay, 1997, *The econometrics of financial markets* (Princeton University Press, Princeton, New Jersey).
- [17] Cecchetti, Stephen G., Lam, Pok-Sang and Nelson C. Mark, 1994, The equity premium and the riskfree rate: Matching the moments, *Journal of Finance* 49, 123-152.
- [18] Cochrane, John, 1991, A simple test of consumption insurance, *Journal of Political Economy* 99, 957-976.
- [19] Cochrane, John, 2001, *Asset Pricing* (Princeton University Press, Princeton, New Jersey).
- [20] Cochrane, John and Lars Peter Hansen, 1992, Asset pricing lessons for macroeconomics, in O. Blanchard and S. Fischer, eds.: 1992 NBER Macroeconomics Annual (National Bureau of Economic Research, Cambridge, MA).
- [21] Cogley, Timothy, 1999, Idiosyncratic Risk and the Equity Premium: Evidence from the Consumer Expenditure Survey, working paper, Arizona State University.

- [22] Constantinides, George M., 1982, Intertemporal asset pricing with heterogeneous consumers and without demand aggregation, *Journal of Business* 55, 253-267.
- [23] Constantinides, George M., 1986, Capital market equilibrium with transaction costs, *Journal of Political Economy* 94, 842-862.
- [24] Constantinides, George M., 1990, Habit formation: A resolution of the equity premium puzzle, *Journal of Political Economy* 98, 519-543.
- [25] Constantinides, George M. and Darrell Duffie, 1996, Asset pricing with heterogeneous consumers, *Journal of Political Economy* 104, 219-487.
- [26] Constantinides, George M., Donaldson, John B. and Rajnish Mehra, 1998, Junior can't borrow: a new perspective on the equity premium puzzle, mimeo, University of Chicago, January.
- [27] Detemple, Jerome and Fernando Zapatero, 1991, Asset prices in an exchange economy with habit formation, *Econometrica* 59, 1633-1657.
- [28] Epstein, Larry G. and Stanley E. Zin, 1989, Substitution, risk aversion, and the temporal behavior of asset returns: A theoretical framework, *Econometrica* 57, 937-969.
- [29] Epstein, Larry G. and Stanley E. Zin, 1991, Substitution, risk aversion, and the temporal behavior of asset returns: An empirical analysis, *Journal of Political Economy* 99, 263-287.
- [30] Ferson, Wayne E., 1995, Theory and Empirical Testing of Asset Pricing Models, in R.A. Jarrow, V. Maksimovic and W.T. Ziemba, eds.: *Handbooks in Operations Research and Management Science* (Elsevier Science B.V, Amsterdam).
- [31] Ferson, Wayne E. and George M. Constantinides, 1991, Habit persistence and durability in aggregate consumption: empirical tests, *Journal of Financial Economics* 29, 199-240.
- [32] Ferson, Wayne E. and Campbell R. Harvey, 1992, Seasonality and Consumption-Based Asset Pricing, *Journal of Finance* 47, 511-552.
- [33] Ferson, Wayne E. and Campbell R. Harvey, 1993, Seasonality and Heteroskedasticity in Consumption-Based Asset Pricing: An Analysis of Linear Models, *Research in Finance* 11, 1-35.
- [34] Grossman, Sanford J. and Robert J. Shiller, 1982, Consumption correlatedness and risk measurement in economies with non-traded assets and heterogeneous information, *Journal of Financial Economics* 10, 195-210.

- [35] Grossman, Sanford J., Melino, Angelo and Robert J. Shiller, 1987, Estimating the continuous-time consumption-based asset pricing model, *Journal of Business and Economics Statistics* 5, 315-327.
- [36] Hall, Robert E., and Frederic S. Mishkin, 1982, The sensitivity of consumption to transitory income: Estimates from panel data on households, *Econometrica* 50, 461-482.
- [37] Hansen, Lars Peter, 1982, Large sample properties of generalized method of moments estimators, *Econometrica* 50, 1029-1054.
- [38] Hansen, Lars Peter and Ravi Jagannathan, 1991, Implications of security market data for models of dynamic economies, *Journal of Political Economy* 99, 225-262.
- [39] Hansen, Lars Peter and Kenneth Singleton, 1982, Generalized instrumental variables estimation of nonlinear rational expectations models, *Econometrica* 50, 1269-1286.
- [40] Hansen, Lars Peter and Kenneth Singleton, 1983, Stochastic consumption, risk aversion and the temporal behavior of asset returns, *Journal of Political Economy* 91, 249-265.
- [41] Hansen, Lars Peter and Kenneth Singleton, 1984, Errata, *Econometrica* 52, 267-268.
- [42] Hansen, Lars Peter, Heaton, J. and Erzo G.J. Luttmer, *Econometric evaluation of asset pricing models*, *Review of Financial Studies* 8, 237-274.
- [43] Hayashi, Fumio, Joseph Altonji and Laurence Kotlikoff, 1996, Risk sharing across and within families, *Econometrica* 64, 261-294.
- [44] He, Hua and David M. Modest, 1995, Market frictions and consumption-based asset pricing, *Journal of Political Economy* 103, 94-117.
- [45] Heaton, John, 1995, An empirical investigation of asset pricing with temporally dependent preference specifications, *Econometrica* 63, 618-717.
- [46] Heaton, John and Deborah Lucas, 1995, The importance of investor heterogeneity and financial market imperfections for the behavior of asset prices, *Carnegie-Rochester Series on Public Policy* 42,1-32.
- [47] Heaton, John and Deborah Lucas, 1996, Evaluating the effects of incomplete markets for risk sharing and asset pricing, *Journal of Political Economy* 104, 443-487.
- [48] Huggett, Mark, 1993, The risk-free rate in heterogeneous-agent incomplete-insurance economies, *Journal of Economic Dynamics and Control* 17, 953-969.

- [49] Jacobs, Kris, 1999, Incomplete markets and security prices: do asset-pricing puzzles result from aggregation problems?, *Journal of Finance* 54, 123-164.
- [50] Kandel, Shmuel and Robert F. Stambaugh, 1991, Asset returns and intertemporal preferences, *Journal of Monetary Economics* 27, 39-71.
- [51] Kocherlakota, Narayana, 1990, On the 'discount' factor in growth economies, *Journal of Monetary Economics* 25, 43-47.
- [52] Kocherlakota, Narayana, 1996, The equity premium: It's still a puzzle, *Journal of Economic Literature* 34, 42-71.
- [53] Lucas, Robert E. Jr., 1978, Asset prices in an exchange economy, *Econometrica* 46, 1429-1446.
- [54] Lucas, Deborah, 1994, Asset pricing with undiversifiable income risk and short sales constraints: Deepening the equity premium puzzle, *Journal of Monetary Economics* 34, 325-341.
- [55] Luttmer, Erzo, 1996, Asset pricing in economies with frictions, *Econometrica* 64, 1439-1467.
- [56] Mace, Barbara, 1991, Full insurance in the presence of aggregate uncertainty, *Journal of Political Economy* 99, 928-956.
- [57] Mankiw, N. Gregory, 1986, The equity premium and the concentration of aggregate shocks, *Journal of Financial Economics* 17, 211-219.
- [58] Mankiw, N. Gregory, and Stephen P. Zeldes, 1991, The consumption of stockholders and nonstockholders, *Journal of Financial Economics* 29, 113-135.
- [59] Mehra, Rajnish and Edward C. Prescott, 1985, The equity premium: A puzzle, *Journal of Monetary Economics* 15, 145-161.
- [60] Miller, Robert. and Holger Sieg, 1997, A microeconomic comparison of household behavior between countries, *Journal of Business and Economics Statistics* 15, 237-253.
- [61] Runkle, David E., 1991, Liquidity constraints and the permanent-income hypothesis: Evidence from panel data, *Journal of Monetary Economics* 27, 73-98.
- [62] Singleton, Kenneth, 1990, Specification and estimation of intertemporal asset pricing models, in B. Friedman and F. Hahn, eds.: *Handbook of Monetary Economics* (North-Holland, Amsterdam).

- [63] Sundaresan, Suresh, 1989, Intertemporally dependent preferences and the volatility of consumption and wealth, *Review of Financial Studies* 2, 73-88.
- [64] Telmer, Chris I., 1993, Asset pricing puzzles and incomplete markets, *Journal of Finance* 48, 1803-1832.
- [65] Telmer, Chris I., Kjetil Storesletten and Amir Yaron, 1997, Asset pricing with idiosyncratic risk and overlapping generations, GSIA working paper no. 1997-E226, Carnegie Mellon University.
- [66] Vissing-Jorgensen, Annette, 1999, Limited Stock Market Participation and the Equity Premium Puzzle, working paper, University of Chicago.
- [67] Weil, Philippe, 1989, The equity premium puzzle and the riskfree rate puzzle, *Journal of Monetary Economics* 24, 401-421.
- [68] Weil, Philippe, 1990, Nonexpected utility in macroeconomics, *Quarterly Journal of Economics* 105, 28-42.
- [69] Zeldes, Stephen P., 1989, Consumption and liquidity constraints, *Journal of Political Economy* 97, 305-346

**Table I: Descriptive Statistics for Household and Cohort Consumption Growth**

This table presents mean consumption growth for individual households and cohorts as well as the minimum and maximum consumption growth for individual households and cohorts. Panel A contains data on individual consumption growth. Panel B contains data on cohort consumption growth, with cohorts sorted by age and education. Panel C contains data on cohort consumption growth, with cohorts sorted by age. In Panel A, average consumption growth is indicated by  $cg_{av}$ , minimum consumption growth by  $cg_{\min}$  and maximum consumption growth by  $cg_{\max}$ . In panels B and C, we have two types of consumption growth. The first type, with average consumption growth  $cg_{1,av}$ , minimum consumption growth  $cg_{1,\min}$  and maximum consumption growth  $cg_{1,\max}$  is obtained by first averaging over the level of consumption in the cohort to get average cohort consumption, and subsequently computing consumption growth. The second type of consumption growth, with average consumption growth  $cg_{2,av}$ , minimum consumption growth  $cg_{2,\min}$  and maximum consumption growth  $cg_{2,\max}$  is obtained by averaging over individual household's consumption growth in the cohort. Every table also indicates the number of observations  $\#obs$ , average current and lagged family size  $fs_t$  and  $fs_{t-1}$ , and the average percentage lagged occupational unemployment rate  $occ_{t-1}$ . Panel A also contains the averages of the current and lagged (gross) real return on the risky asset  $R_{ri,t}$  and  $R_{ri,t-1}$  and the current and lagged (gross) real return on the riskless asset  $R_{rl,t}$  and  $R_{rl,t-1}$ .

Panel A: Data on Individual Households

sample	#obs	$cg_{av}$	$cg_{\min}$	$cg_{\max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$	$R_{ri,t}$	$R_{ri,t-1}$	$R_{rl,t}$	$R_{rl,t-1}$
>0	5029	1.061	0.224	4.387	3.180	3.154	5.026	1.083	1.019	1.016	1.010
>1000	2990	1.060	0.224	4.177	3.143	3.136	4.478	1.079	1.034	1.015	1.008
>10000	769	1.048	0.238	3.513	3.002	3.029	4.076	1.071	1.022	1.011	1.004
all	18813	1.069	0.116	9.741	3.285	3.241	6.510	1.081	1.037	1.016	1.009

Panel B: Data on cohorts sorted using age and education

sample	#obs	$cg_{1,av}$	$cg_{1,\min}$	$cg_{1,\max}$	$cg_{2,av}$	$cg_{2,\min}$	$cg_{2,\max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$
>0	288	1.005	0.653	1.348	1.046	0.653	1.344	3.281	3.297	5.225
>1000	287	1.013	0.577	1.542	1.052	0.584	1.693	3.206	3.218	4.985
>10000	249	1.046	0.321	3.513	1.064	0.394	3.513	3.117	3.158	4.537
all	288	1.000	0.778	1.191	1.048	0.855	1.259	3.399	3.417	5.970

Panel C: Data on cohorts sorted using age

sample	#obs	$cg_{1,av}$	$cg_{1,\min}$	$cg_{1,\max}$	$cg_{2,av}$	$cg_{2,\min}$	$cg_{2,\max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$
>0	432	1.010	0.544	1.408	1.050	0.648	1.605	3.260	3.274	4.931
>1000	432	1.020	0.443	2.770	1.057	0.443	2.770	3.174	3.182	4.503
>10000	336	1.040	0.321	3.513	1.056	0.489	3.513	3.056	3.080	4.020
all	432	0.999	0.726	1.279	1.047	0.803	1.418	3.429	3.450	6.394

**Table II: Descriptive Statistics for Household Consumption Growth (year by year)**

This table presents mean consumption growth for individual households as well as the minimum and maximum consumption growth for individual households (all on a year-by-year basis). Average consumption growth is indicated by  $cg_{av}$ , minimum consumption growth by  $cg_{min}$  and maximum consumption growth by  $cg_{max}$ . The table further indicates the number of observations in a given year  $\#obs$ , average current and lagged family size  $fs_t$  and  $fs_{t-1}$ , the average percentage lagged occupational unemployment rate  $occ_{t-1}$ , as well as the current and lagged (gross) real return on the risky asset  $R_{ri,t}$  and  $R_{ri,t-1}$  and the current and lagged (gross) real return on the riskless asset  $R_{rl,t}$  and  $R_{rl,t-1}$ .

Panel A: Asset Holdings > 0											
year	$\#obs$	$cg_{av}$	$cg_{min}$	$cg_{max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$	$R_{ri,t}$	$R_{ri,t-1}$	$R_{rl,t}$	$R_{rl,t-1}$
1974	368	1.026	0.332	3.379	3.315	3.331	3.651	1.018	0.723	0.981	0.969
1975	381	1.109	0.315	3.357	3.314	3.251	5.926	1.166	1.018	0.998	0.981
1976	423	1.078	0.342	4.387	3.241	3.245	5.474	0.946	1.166	0.985	0.998
1977	419	1.059	0.224	4.177	3.238	3.210	4.636	0.957	0.946	0.985	0.985
1978	421	1.026	0.354	4.141	3.216	3.180	3.928	1.017	0.957	0.975	0.985
1979	421	1.064	0.323	3.682	3.140	3.118	3.655	1.066	1.017	0.976	0.975
1980	408	1.037	0.347	2.853	3.127	3.085	4.753	1.037	1.066	1.021	0.976
1981	427	1.052	0.276	3.185	3.114	3.084	5.371	0.936	1.037	1.065	1.021
1982	442	1.089	0.301	2.928	3.101	3.070	6.756	1.363	0.936	1.058	1.065
1983	435	1.078	0.350	3.753	3.119	3.094	6.535	1.006	1.363	1.046	1.058
1984	434	1.047	0.352	3.191	3.149	3.108	4.868	1.172	1.006	1.050	1.046
1985	450	1.065	0.238	3.513	3.126	3.113	4.553	1.285	1.172	1.049	1.050
Panel B: Asset Holdings > 1000											
year	$\#obs$	$cg_{av}$	$cg_{min}$	$cg_{max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$	$R_{ri,t}$	$R_{ri,t-1}$	$R_{rl,t}$	$R_{rl,t-1}$
1974	237	1.030	0.363	3.379	3.299	3.295	3.280	1.018	0.723	0.981	0.969
1975	246	1.113	0.387	2.424	3.321	3.260	5.168	1.166	1.018	0.998	0.981
1976	263	1.070	0.342	3.232	3.235	3.250	4.847	0.946	1.166	0.985	0.998
1977	259	1.061	0.224	4.177	3.239	3.220	4.226	0.957	0.946	0.985	0.985
1978	256	1.035	0.354	4.141	3.183	3.167	3.529	1.017	0.957	0.975	0.985
1979	257	1.066	0.476	3.682	3.132	3.136	3.319	1.066	1.017	0.976	0.975
1980	249	1.036	0.401	2.853	3.080	3.048	4.189	1.037	1.066	1.021	0.976
1981	248	1.035	0.380	2.471	3.060	3.084	4.682	0.936	1.037	1.065	1.021
1982	247	1.081	0.301	2.928	3.016	3.016	6.008	1.363	0.936	1.058	1.065
1983	242	1.069	0.394	3.753	3.074	3.041	5.810	1.006	1.363	1.046	1.058
1984	236	1.065	0.378	3.191	3.067	3.097	4.484	1.172	1.006	1.050	1.046
1985	250	1.064	0.238	3.513	3.008	3.012	4.245	1.285	1.172	1.049	1.050

**Table II (continued)**

Panel C: Asset Holdings > 10000											
year	#obs	$cg_{av}$	$cg_{min}$	$cg_{max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$	$R_{ri,t}$	$R_{ri,t-1}$	$R_{rl,t}$	$R_{rl,t-1}$
1974	79	1.019	0.368	2.088	3.177	3.215	2.970	1.018	0.723	0.981	0.969
1975	81	1.084	0.580	2.424	3.222	3.172	4.820	1.166	1.018	0.998	0.981
1976	77	1.068	0.342	2.593	3.168	3.246	4.545	0.946	1.166	0.985	0.998
1977	71	0.960	0.415	1.836	3.197	3.239	4.044	0.957	0.946	0.985	0.985
1978	65	1.096	0.595	2.556	2.907	3.046	3.190	1.017	0.957	0.975	0.985
1979	63	1.053	0.516	2.367	2.920	2.873	2.924	1.066	1.017	0.976	0.975
1980	63	1.015	0.506	2.208	2.984	2.936	3.773	1.037	1.066	1.021	0.976
1981	61	1.029	0.432	2.471	2.868	2.934	4.303	0.936	1.037	1.065	1.021
1982	59	1.013	0.652	1.838	2.915	2.898	5.311	1.363	0.936	1.058	1.065
1983	54	1.045	0.394	2.250	2.833	2.907	5.118	1.006	1.363	1.046	1.058
1984	49	1.157	0.669	3.191	2.857	2.877	3.974	1.172	1.006	1.050	1.046
1985	47	1.061	0.238	3.513	2.659	2.702	4.168	1.285	1.172	1.049	1.050
Panel D: All Households											
year	#obs	$cg_{av}$	$cg_{min}$	$cg_{max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$	$R_{ri,t}$	$R_{ri,t-1}$	$R_{rl,t}$	$R_{rl,t-1}$
1974	1385	1.042	0.282	4.373	3.627	3.606	4.752	1.018	0.723	0.981	0.969
1975	1434	1.099	0.308	5.945	3.543	3.500	7.637	1.166	1.018	0.998	0.981
1976	1552	1.084	0.301	5.561	3.440	3.416	6.679	0.946	1.166	0.985	0.998
1977	1627	1.087	0.209	5.230	3.386	3.342	5.864	0.957	0.946	0.985	0.985
1978	1599	1.036	0.218	4.883	3.312	3.280	4.950	1.017	0.957	0.975	0.985
1979	1643	1.064	0.314	9.741	3.265	3.193	4.820	1.066	1.017	0.976	0.975
1980	1556	1.037	0.231	5.207	3.185	3.151	6.419	1.037	1.066	1.021	0.976
1981	1553	1.054	0.276	5.127	3.175	3.121	6.882	0.936	1.037	1.065	1.021
1982	1569	1.107	0.258	4.072	3.163	3.107	8.976	1.363	0.936	1.058	1.065
1983	1599	1.082	0.169	6.916	3.149	3.093	8.673	1.006	1.363	1.046	1.058
1984	1631	1.065	0.116	7.643	3.152	3.094	6.372	1.172	1.006	1.050	1.046
1985	1665	1.075	0.238	5.018	3.104	3.066	6.113	1.285	1.172	1.049	1.050

**Table III: Descriptive Statistics for Cohort Consumption Growth  
(year by year, cohorts formed using age and education)**

This table presents mean consumption growth for cohorts as well as the minimum and maximum consumption growth for cohorts (all on a year-by-year basis). The cohorts are sorted using age and education. The table reports on two types of cohort consumption. The first type, with average consumption growth  $cg_{1,av}$ , minimum consumption growth  $cg_{1,min}$  and maximum consumption growth  $cg_{1,max}$  is obtained by averaging over the level of consumption in the cohort. The second type, with average consumption growth  $cg_{2,av}$ , minimum consumption growth  $cg_{2,min}$  and maximum consumption growth  $cg_{2,max}$  is obtained by averaging over consumption growth in the cohort. The table further indicates the number of observations in a given year  $\#obs$ , average current and lagged family size  $fs_t$  and  $fs_{t-1}$  and the average percentage lagged occupational unemployment rate  $occ_{t-1}$ .

Panel A: Asset Holdings > 0										
year	#obs	$cg_{1,av}$	$cg_{1,min}$	$cg_{1,max}$	$cg_{2,av}$	$cg_{2,min}$	$cg_{2,max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$
1974	24	0.986	0.768	1.165	1.019	0.795	1.217	3.491	3.554	3.580
1975	24	1.065	0.851	1.348	1.098	0.870	1.344	3.483	3.459	6.131
1976	24	1.030	0.778	1.139	1.058	0.771	1.242	3.424	3.460	5.510
1977	24	0.989	0.762	1.152	1.037	0.821	1.193	3.436	3.466	4.707
1978	24	0.989	0.790	1.290	1.035	0.865	1.300	3.348	3.356	3.904
1979	24	1.000	0.734	1.187	1.040	0.824	1.185	3.277	3.303	3.690
1980	24	0.981	0.800	1.238	1.027	0.841	1.294	3.263	3.274	4.944
1981	24	0.991	0.852	1.185	1.031	0.865	1.224	3.197	3.221	5.584
1982	24	1.023	0.691	1.196	1.072	0.764	1.262	3.149	3.149	7.163
1983	24	1.011	0.662	1.268	1.054	0.653	1.341	3.136	3.184	7.261
1984	24	0.994	0.759	1.184	1.029	0.809	1.225	3.089	3.067	5.237
1985	24	0.999	0.653	1.195	1.050	0.722	1.334	3.077	3.070	4.985

Panel B: Asset Holdings > 1000										
year	#obs	$cg_{1,av}$	$cg_{1,min}$	$cg_{1,max}$	$cg_{2,av}$	$cg_{2,min}$	$cg_{2,max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$
1974	24	0.990	0.758	1.140	1.030	0.785	1.498	3.389	3.412	3.330
1975	24	1.076	0.898	1.341	1.107	0.918	1.330	3.370	3.344	5.666
1976	24	1.027	0.807	1.268	1.059	0.806	1.459	3.362	3.377	5.148
1977	24	1.017	0.707	1.304	1.062	0.754	1.338	3.353	3.390	4.500
1978	24	0.983	0.775	1.337	1.036	0.872	1.358	3.254	3.252	3.736
1979	24	1.005	0.577	1.262	1.038	0.584	1.387	3.212	3.241	3.618
1980	24	0.994	0.803	1.542	1.036	0.848	1.546	3.127	3.170	4.685
1981	24	0.994	0.756	1.229	1.022	0.765	1.297	3.105	3.141	5.421
1982	24	1.007	0.655	1.268	1.060	0.728	1.454	3.016	3.002	6.913
1983	24	1.027	0.687	1.453	1.053	0.692	1.693	3.089	3.095	6.596
1984	24	1.014	0.729	1.234	1.051	0.729	1.294	3.128	3.129	5.203
1985	24	1.028	0.671	1.504	1.073	0.761	1.662	3.060	3.052	5.066

**Table III (continued)**

Panel C: Asset Holdings &gt; 10000

year	#obs	cg1, <i>av</i>	cg1, <i>min</i>	cg1, <i>max</i>	cg2, <i>av</i>	cg2, <i>min</i>	cg2, <i>max</i>	fs <sub><i>t</i></sub>	fs <sub><i>t-1</i></sub>	occ <sub><i>t-1</i></sub>
1974	23	0.994	0.655	1.212	0.999	0.655	1.224	3.480	3.562	2.947
1975	24	1.098	0.727	2.273	1.117	0.733	2.273	3.399	3.467	5.477
1976	24	1.039	0.748	1.532	1.065	0.793	1.523	3.376	3.401	5.424
1977	24	0.933	0.555	1.525	0.956	0.555	1.525	3.340	3.459	4.446
1978	22	1.051	0.595	1.436	1.067	0.595	1.510	2.972	3.075	3.536
1979	20	1.127	0.703	2.367	1.143	0.778	2.367	3.053	3.043	3.150
1980	23	1.018	0.588	1.751	1.031	0.588	1.751	3.125	3.104	4.325
1981	20	1.015	0.578	2.471	1.035	0.688	2.471	2.990	3.009	5.203
1982	18	1.032	0.772	1.838	1.042	0.772	1.838	2.932	2.906	5.799
1983	16	1.021	0.394	1.296	1.034	0.394	1.453	2.880	2.990	5.514
1984	16	1.169	0.735	2.461	1.173	0.726	2.461	2.856	2.835	4.324
1985	19	1.104	0.321	3.513	1.149	0.526	3.513	2.663	2.678	4.605

Panel D: All Households

year	#obs	cg1, <i>av</i>	cg1, <i>min</i>	cg1, <i>max</i>	cg2, <i>av</i>	cg2, <i>min</i>	cg2, <i>max</i>	fs <sub><i>t</i></sub>	fs <sub><i>t-1</i></sub>	occ <sub><i>t-1</i></sub>
1974	24	0.980	0.834	1.092	1.017	0.855	1.154	3.732	3.768	4.250
1975	24	1.037	0.854	1.172	1.089	0.902	1.259	3.694	3.698	6.849
1976	24	1.021	0.895	1.089	1.064	0.959	1.167	3.630	3.666	6.092
1977	24	1.005	0.821	1.085	1.055	0.869	1.190	3.565	3.574	5.373
1978	24	0.971	0.872	1.191	1.014	0.912	1.219	3.456	3.506	4.454
1979	24	0.992	0.862	1.091	1.039	0.900	1.136	3.427	3.438	4.401
1980	24	0.967	0.778	1.118	1.021	0.881	1.167	3.321	3.344	5.806
1981	24	0.992	0.839	1.082	1.034	0.873	1.114	3.273	3.284	6.277
1982	24	1.020	0.904	1.110	1.077	0.937	1.190	3.223	3.226	8.223
1983	24	1.022	0.956	1.144	1.071	0.983	1.178	3.196	3.224	8.061
1984	24	1.002	0.832	1.108	1.053	0.875	1.185	3.157	3.154	6.012
1985	24	0.991	0.781	1.087	1.047	0.883	1.141	3.116	3.119	5.837

**Table IV: Descriptive Statistics for Cohort Consumption Growth  
(year by year, cohorts formed using age)**

This table presents mean consumption growth for cohorts as well as the minimum and maximum consumption growth for cohorts (all on a year-by-year basis). The cohorts are sorted using age. The table reports on two types of cohort consumption. The first type, with average consumption growth  $cg_{1,av}$ , minimum consumption growth  $cg_{1,min}$  and maximum consumption growth  $cg_{1,max}$  is obtained by averaging over the level of consumption in the cohort. The second type, with average consumption growth  $cg_{2,av}$ , minimum consumption growth  $cg_{2,min}$  and maximum consumption growth  $cg_{2,max}$  is obtained by averaging over consumption growth in the cohort. The table further indicates the number of observations in a given year  $\#obs$ , average current and lagged family size  $fs_t$  and  $fs_{t-1}$  and the average percentage lagged occupational unemployment rate  $occ_{t-1}$ .

Panel A: Asset Holdings > 0

year	#obs	$cg_{1,av}$	$cg_{1,min}$	$cg_{1,max}$	$cg_{2,av}$	$cg_{2,min}$	$cg_{2,max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$
1974	36	0.996	0.762	1.408	1.024	0.800	1.418	3.441	3.494	3.575
1975	36	1.052	0.731	1.333	1.090	0.842	1.605	3.490	3.464	5.895
1976	36	1.039	0.830	1.377	1.065	0.842	1.284	3.409	3.468	5.361
1977	36	0.997	0.544	1.228	1.046	0.648	1.299	3.445	3.458	4.522
1978	36	0.982	0.684	1.265	1.029	0.776	1.293	3.313	3.318	3.735
1979	36	1.012	0.749	1.289	1.047	0.819	1.328	3.253	3.269	3.558
1980	36	0.979	0.703	1.344	1.029	0.739	1.423	3.198	3.198	4.613
1981	36	0.994	0.742	1.229	1.033	0.830	1.328	3.195	3.224	5.254
1982	36	1.031	0.721	1.349	1.077	0.788	1.498	3.129	3.139	6.643
1983	36	1.036	0.913	1.279	1.076	0.874	1.415	3.085	3.122	6.543
1984	36	1.002	0.761	1.159	1.040	0.810	1.303	3.087	3.054	4.939
1985	36	1.000	0.757	1.255	1.050	0.822	1.379	3.083	3.078	4.533

Panel B: Asset Holdings > 1000

year	#obs	$cg_{1,av}$	$cg_{1,min}$	$cg_{1,max}$	$cg_{2,av}$	$cg_{2,min}$	$cg_{2,max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$
1974	36	0.995	0.760	1.388	1.021	0.811	1.424	3.315	3.344	3.274
1975	36	1.058	0.443	1.548	1.090	0.443	1.852	3.377	3.338	5.400
1976	36	1.037	0.793	1.230	1.069	0.793	1.398	3.301	3.342	4.876
1977	36	1.018	0.652	1.371	1.062	0.790	1.348	3.411	3.412	4.189
1978	36	0.986	0.653	1.408	1.037	0.714	1.450	3.194	3.180	3.392
1979	36	1.023	0.725	1.665	1.050	0.783	1.846	3.131	3.116	3.294
1980	36	0.993	0.703	1.493	1.036	0.739	1.516	3.080	3.083	4.132
1981	36	0.990	0.615	1.219	1.021	0.658	1.378	3.105	3.153	4.806
1982	36	1.024	0.607	1.697	1.073	0.643	1.842	3.054	3.083	6.184
1983	36	1.059	0.883	1.301	1.089	0.854	1.440	3.075	3.077	5.818
1984	36	1.077	0.718	2.770	1.109	0.759	2.770	3.050	3.071	4.420
1985	36	0.985	0.678	1.310	1.027	0.662	1.541	2.992	2.988	4.248

**Table IV (continued)**

Panel C: Asset Holdings &gt; 10000

year	#obs	cg1, <i>av</i>	cg1, <i>min</i>	cg1, <i>max</i>	cg2, <i>av</i>	cg2, <i>min</i>	cg2, <i>max</i>	fs <sub><i>t</i></sub>	fs <sub><i>t-1</i></sub>	occ <sub><i>t-1</i></sub>
1974	30	1.001	0.584	1.368	1.006	0.584	1.358	3.470	3.501	2.837
1975	31	1.036	0.643	2.334	1.058	0.643	2.348	3.389	3.365	4.716
1976	32	1.097	0.748	1.882	1.115	0.793	1.882	3.187	3.252	4.325
1977	31	0.913	0.489	1.212	0.940	0.489	1.212	3.208	3.318	4.072
1978	28	1.094	0.737	1.874	1.106	0.737	1.874	2.913	3.020	3.317
1979	29	1.036	0.521	2.367	1.047	0.521	2.367	2.985	2.978	2.952
1980	28	1.006	0.588	1.751	1.015	0.588	1.751	3.110	3.038	3.608
1981	26	1.027	0.578	2.471	1.046	0.688	2.471	2.872	2.960	4.264
1982	25	1.018	0.732	1.592	1.029	0.737	1.592	2.920	2.926	5.197
1983	25	1.040	0.813	1.264	1.057	0.820	1.403	2.840	2.900	5.055
1984	26	1.148	0.669	2.679	1.152	0.669	2.282	2.851	2.801	3.949
1985	25	1.082	0.321	3.513	1.122	0.526	3.513	2.730	2.690	4.211

Panel D: All Households

year	#obs	cg1, <i>av</i>	cg1, <i>min</i>	cg1, <i>max</i>	cg2, <i>av</i>	cg2, <i>min</i>	cg2, <i>max</i>	fs <sub><i>t</i></sub>	fs <sub><i>t-1</i></sub>	occ <sub><i>t-1</i></sub>
1974	36	0.986	0.842	1.117	1.020	0.888	1.163	3.825	3.865	4.652
1975	36	1.033	0.871	1.279	1.084	0.939	1.418	3.754	3.775	7.533
1976	36	1.013	0.877	1.095	1.058	0.952	1.184	3.661	3.708	6.601
1977	36	1.011	0.880	1.148	1.063	0.940	1.230	3.589	3.606	5.763
1978	36	0.963	0.800	1.098	1.005	0.861	1.142	3.482	3.538	4.770
1979	36	0.992	0.838	1.175	1.038	0.874	1.291	3.420	3.431	4.690
1980	36	0.968	0.808	1.223	1.019	0.861	1.226	3.323	3.356	6.283
1981	36	0.992	0.853	1.182	1.035	0.897	1.207	3.295	3.312	6.696
1982	36	1.021	0.904	1.142	1.080	0.936	1.256	3.259	3.255	8.769
1983	36	1.019	0.890	1.139	1.068	0.912	1.194	3.219	3.241	8.567
1984	36	0.995	0.726	1.100	1.045	0.803	1.219	3.178	3.182	6.357
1985	36	0.996	0.782	1.108	1.051	0.917	1.157	3.136	3.135	6.041

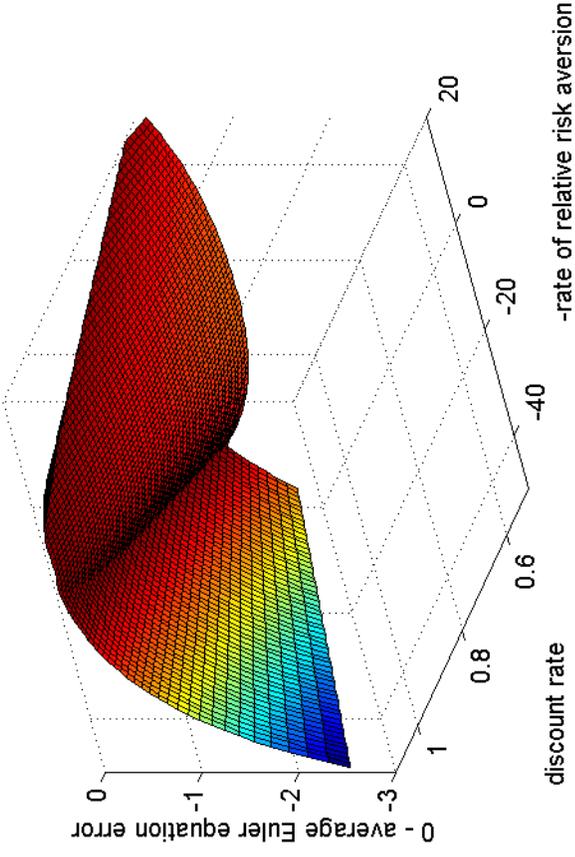
**Table V: Descriptive Statistics for Representative Agent Consumption Growth**

This table presents mean consumption growth as well as the minimum and maximum consumption growth for the representative agent constructed using the individual consumption data. Average consumption growth is indicated by  $cg_{av}$ , minimum consumption growth by  $cg_{min}$  and maximum consumption growth by  $cg_{max}$ . Consumption growth for the representative agent is obtained by first averaging over the household consumption level in a given year to get average consumption, and subsequently computing consumption growth using the average consumption levels in two consecutive years. Every table also indicates the number of observations  $\#obs$  (12 years available in all cases), average current and lagged family size  $fs_t$  and  $fs_{t-1}$ , and the average percentage lagged occupational unemployment rate  $occ_{t-1}$ .

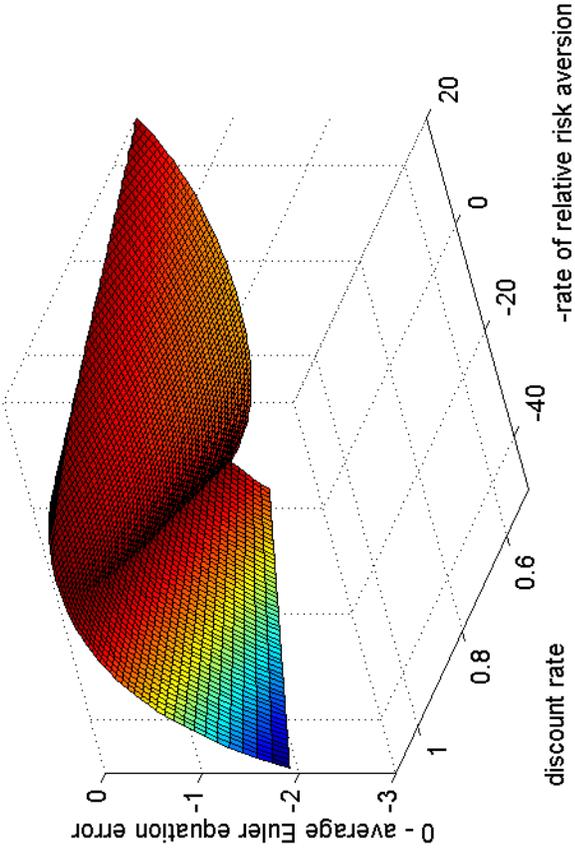
sample	$\#obs$	$cg_{av}$	$cg_{min}$	$cg_{max}$	$fs_t$	$fs_{t-1}$	$occ_{t-1}$
>0	12	1.010	0.966	1.059	3.229	3.208	4.951
>1000	12	1.011	0.960	1.069	3.170	3.166	4.452
>10000	12	1.007	0.905	1.103	2.982	3.011	4.078
all	12	1.007	0.971	1.032	3.359	3.326	6.430

Figure 1: Euler equation error analysis for the Mehra-Prescott data (yearly data 1889-1979; aggregate consumption data).

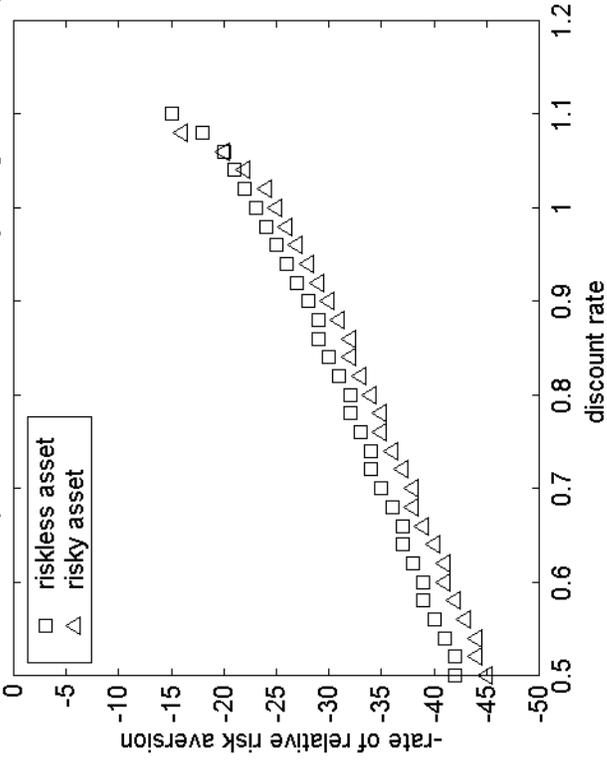
Panel A: Zero - absolute value of Euler equation error (riskless asset)



Panel B: Zero - absolute value of Euler equation error (risky asset)



Panel C: Risk aversion and time preference combinations yielding zero Euler equation errors



Panel D: Statistical Significance of Euler equation errors

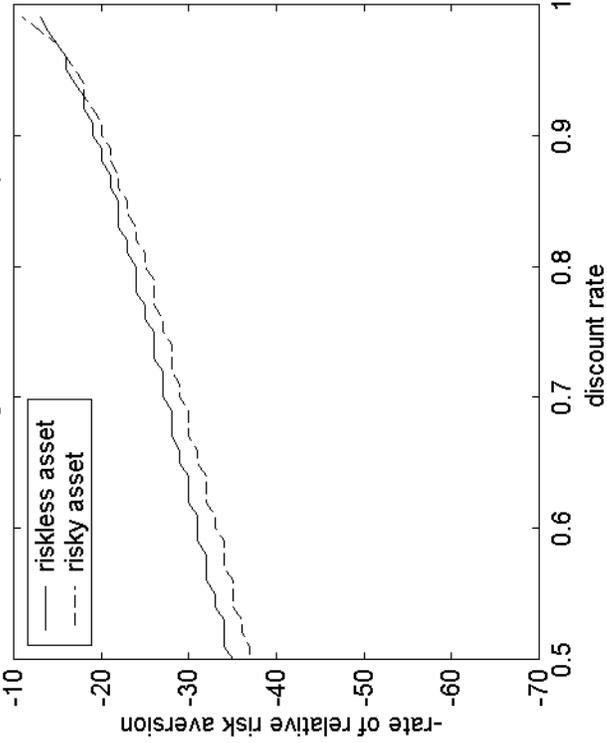


Figure 2: Euler equation error analysis for individual PSID data

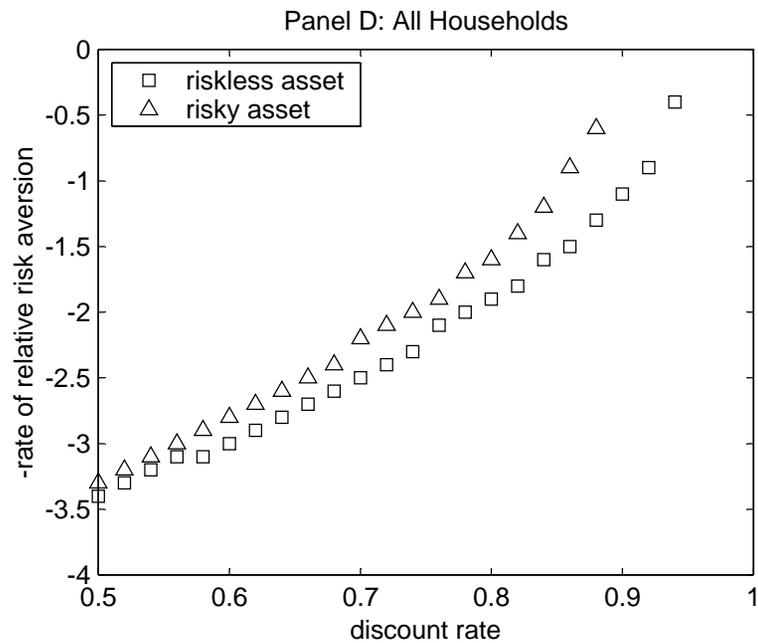
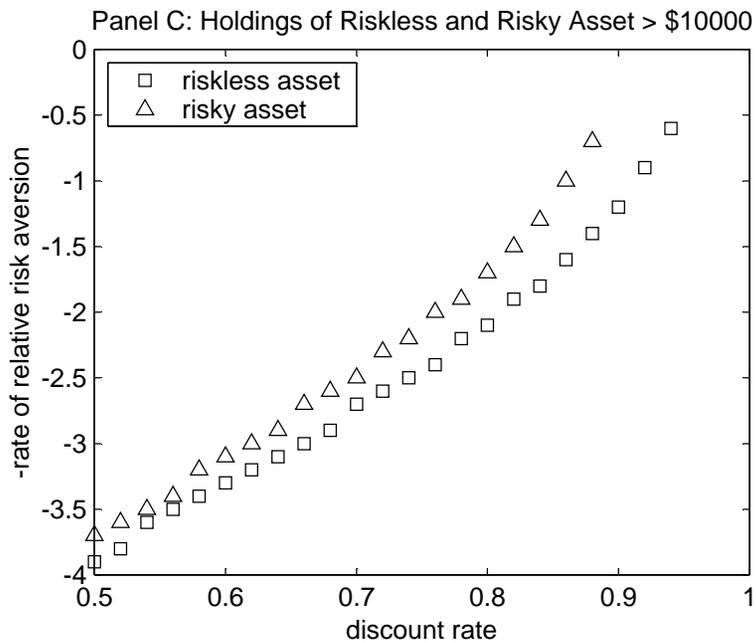
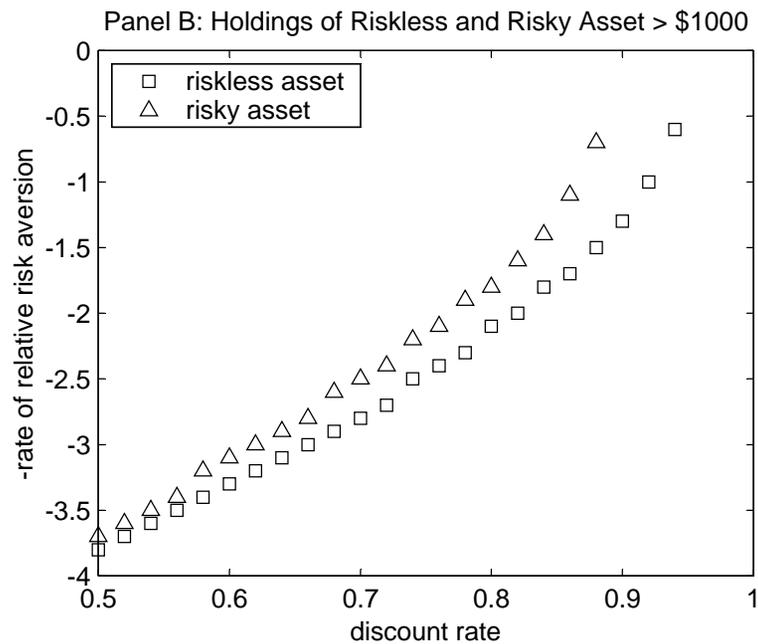
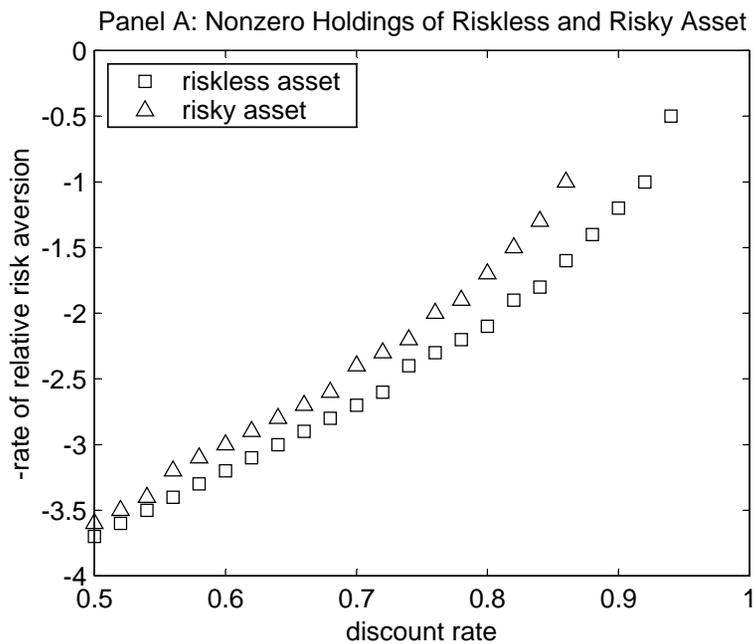


Figure 3: Euler equation error analysis for individual PSID data. The area between the solid (broken) lines indicates insignificant test statistics for the null hypothesis that the Euler error for the riskless (risky) asset is zero.

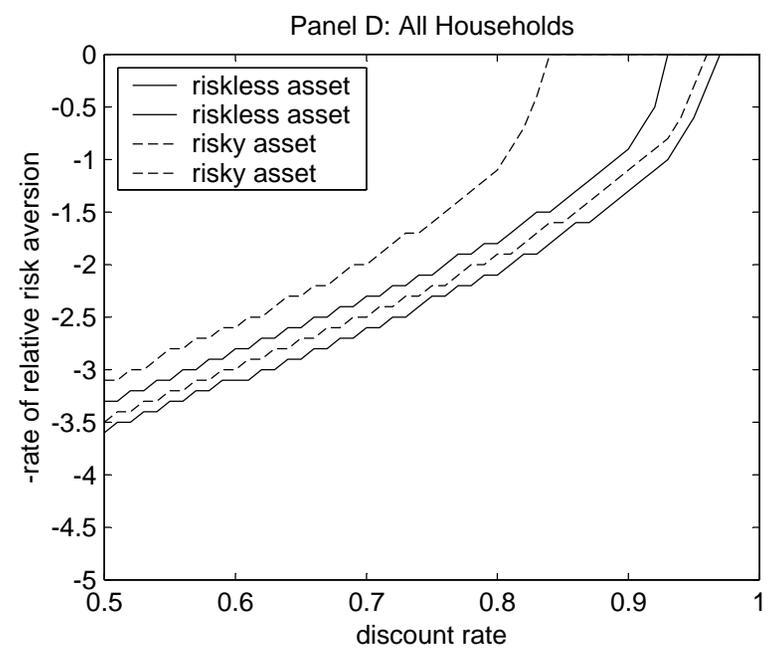
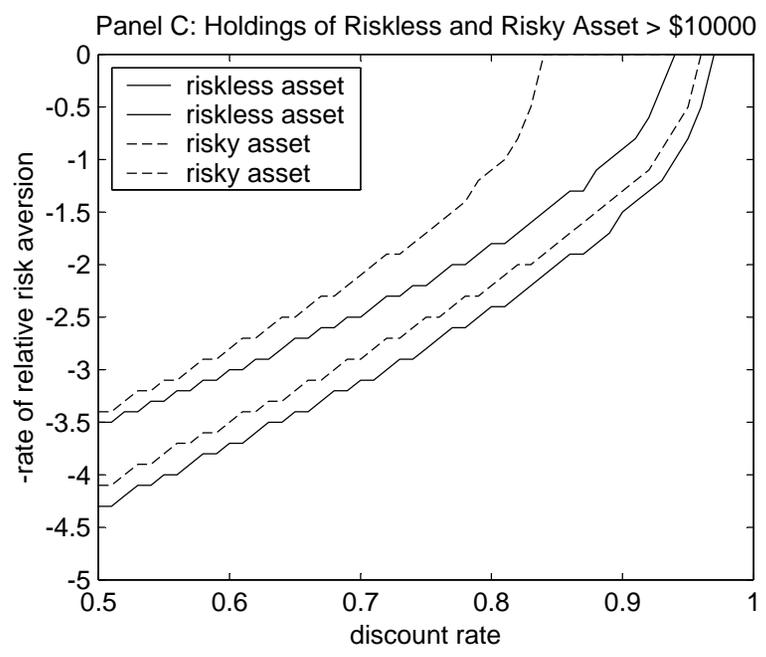
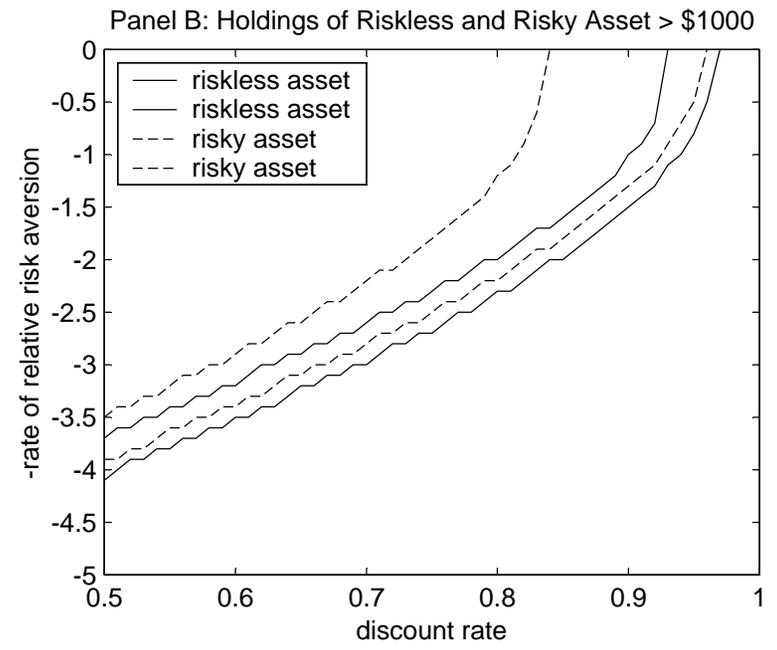
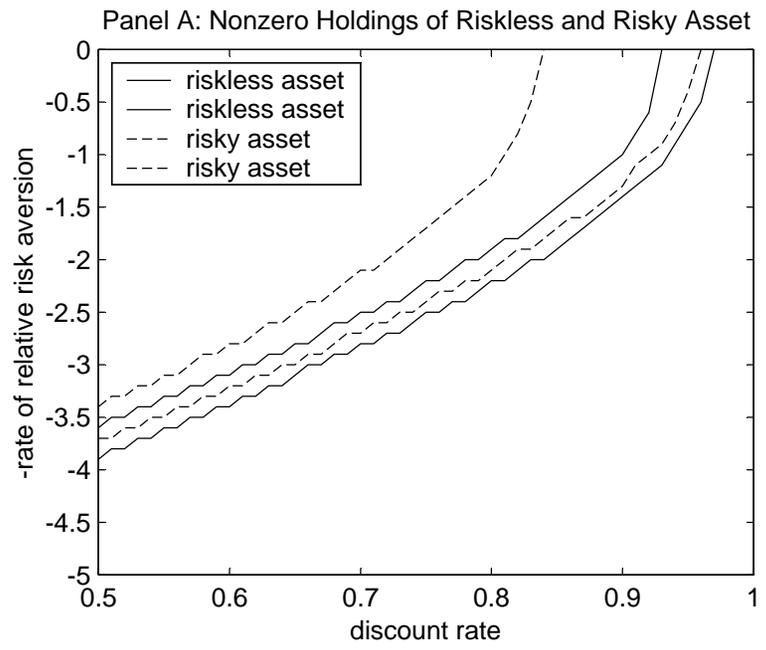


Figure 4: Euler equation error analysis for cohort PSID data. Cohorts are formed using age and education. Consumption growth is obtained by averaging over individual consumption. Squares and triangles indicate discount rate-risk aversion combinations that equate the Euler equation to zero in sample.

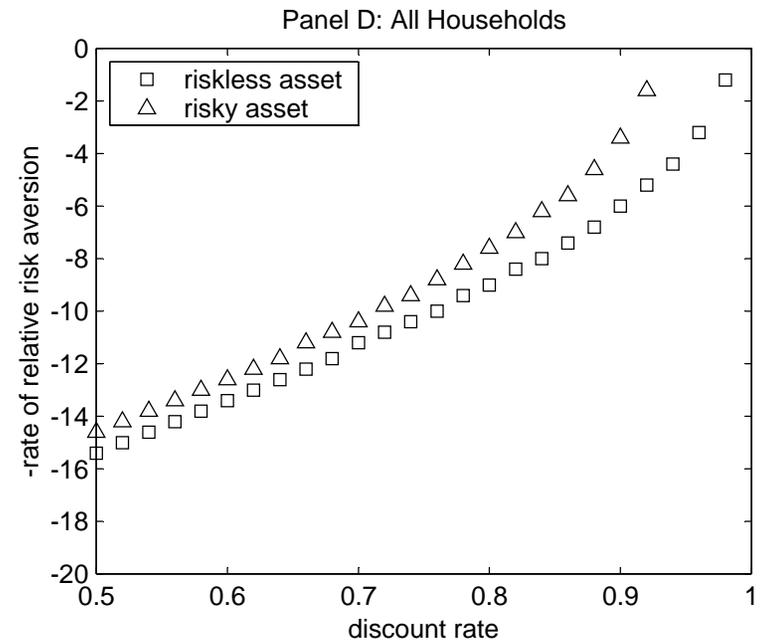
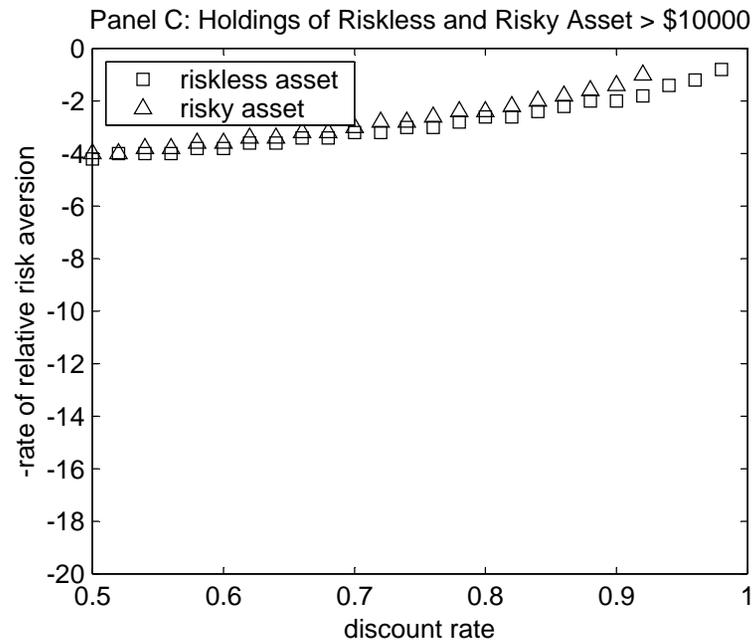
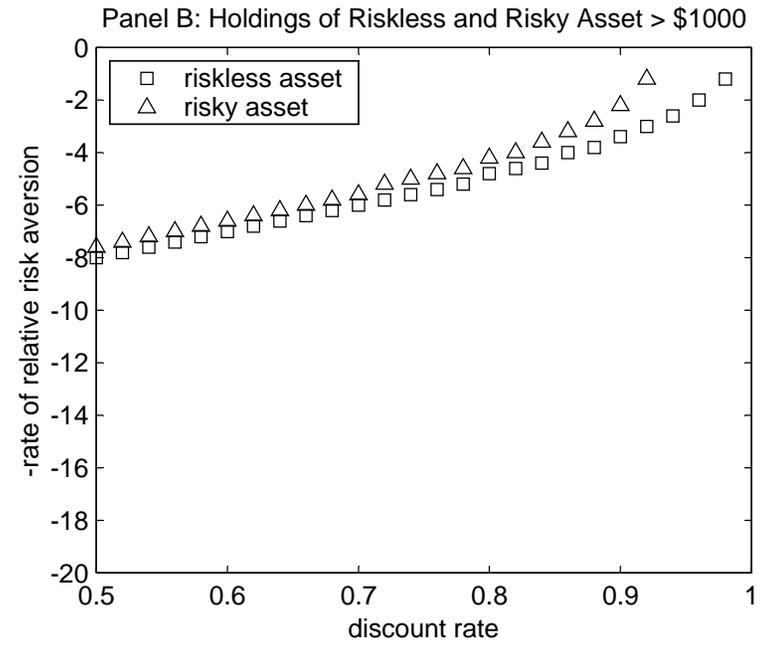
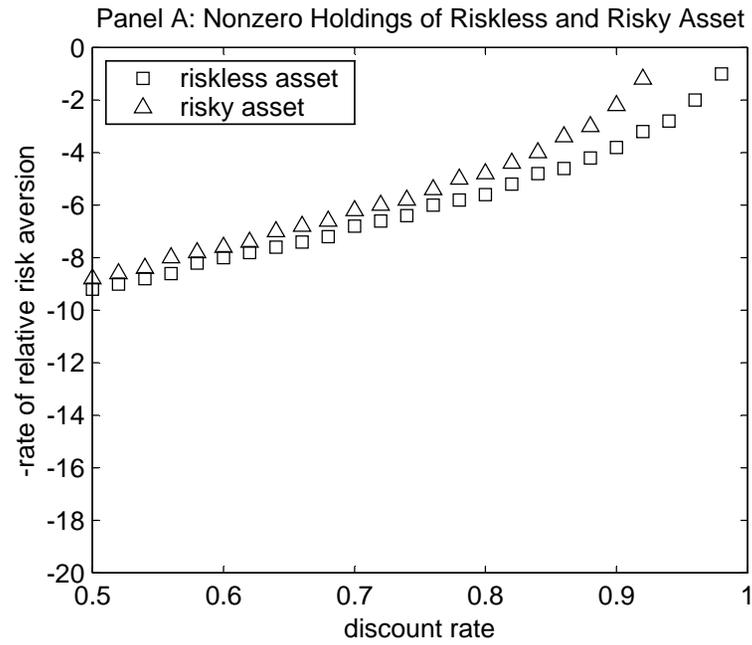


Figure 5: Euler equation error analysis for cohort PSID data. Cohorts are formed using age and education. Consumption growth is obtained by averaging over individual consumption growth. Squares and triangles indicate discount rate-risk aversion combinations that equate the Euler equation to zero in sample.

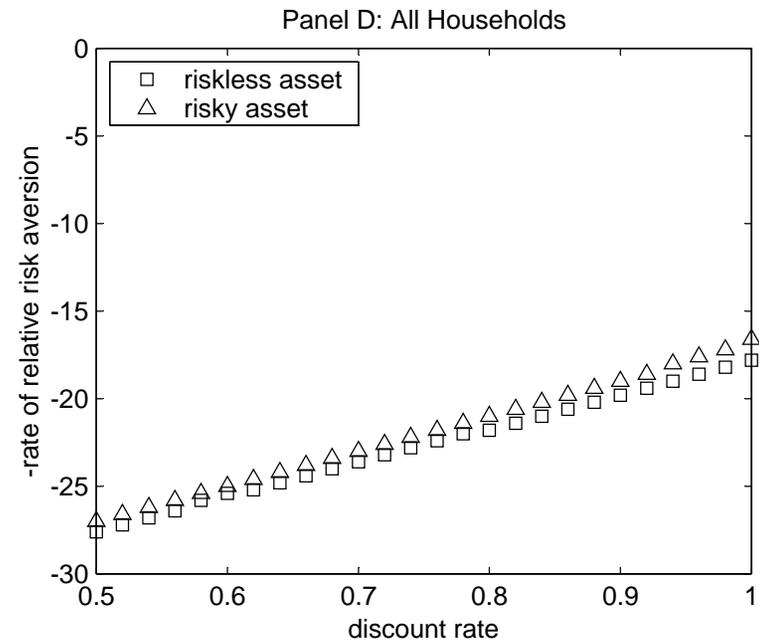
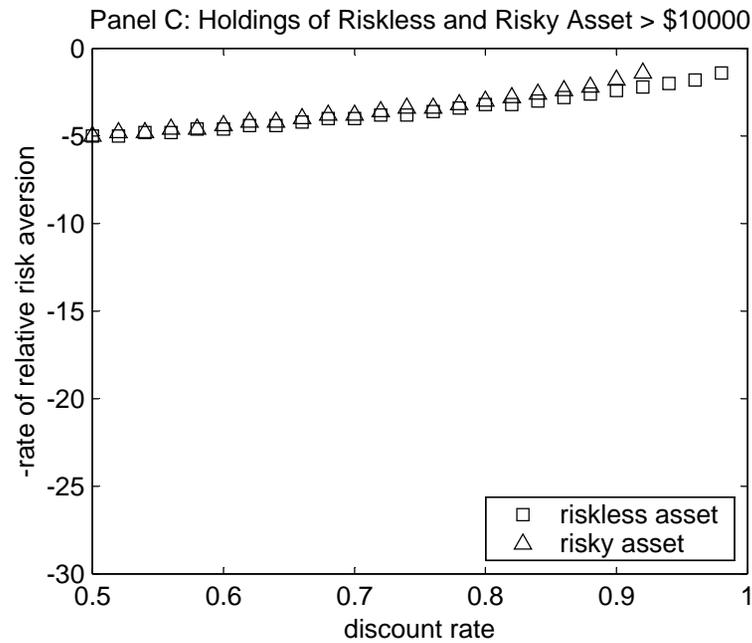
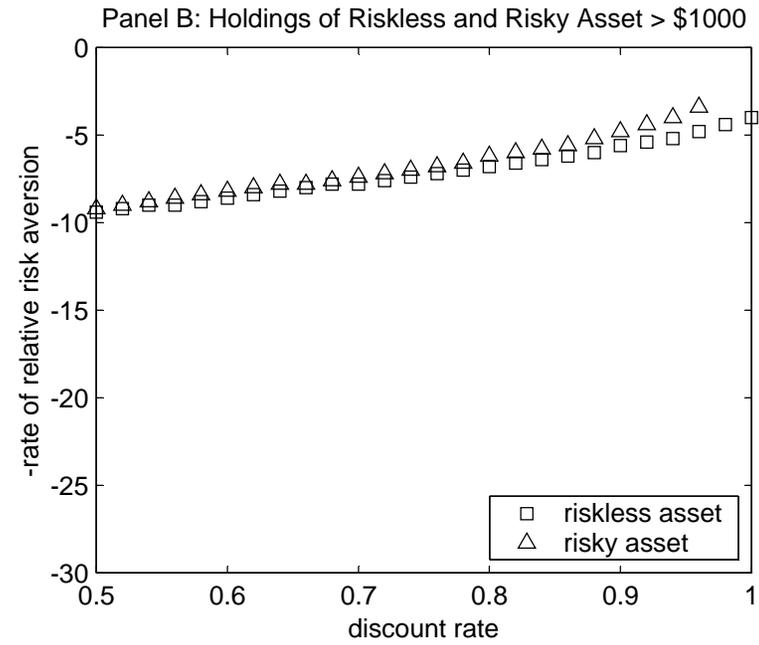
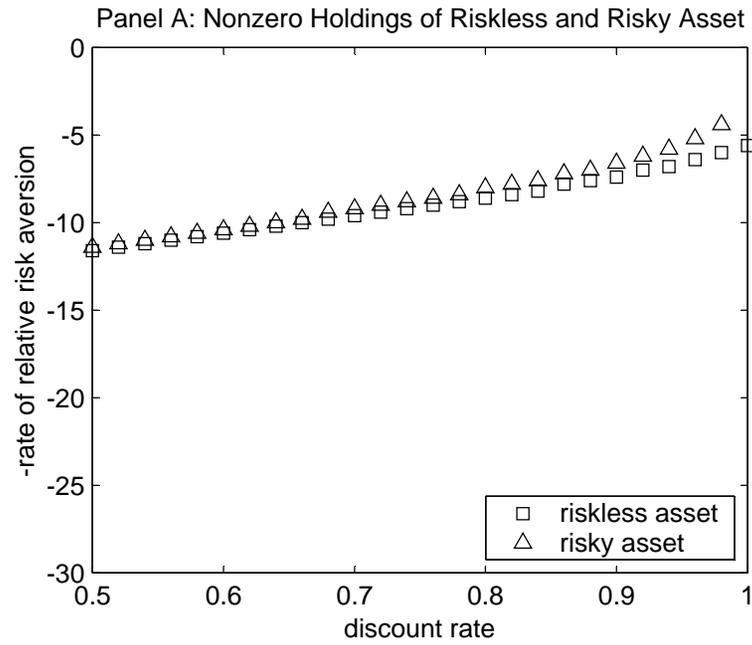


Figure 6: Euler equation error analysis for cohort PSID data. Cohorts are formed using age. Consumption growth is obtained by averaging over individual consumption. Squares and triangles indicate discount rate-risk aversion combinations that equate the Euler equation to zero in sample.

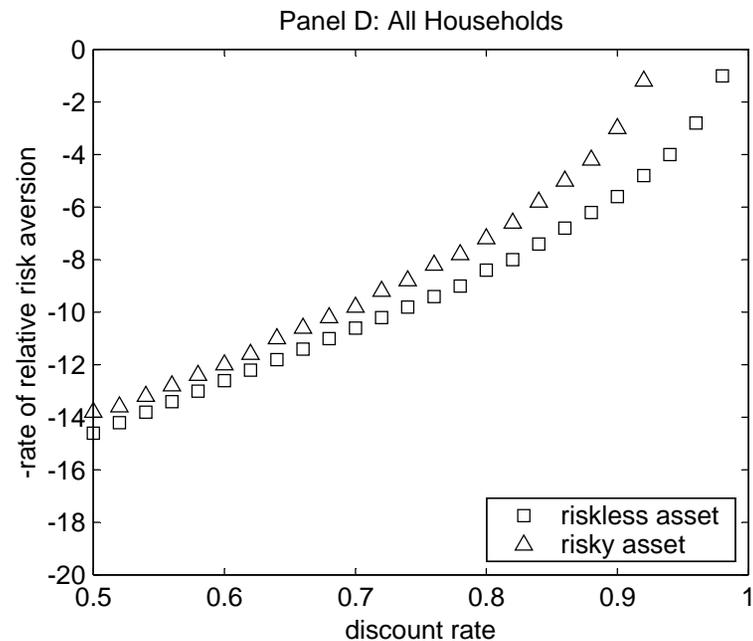
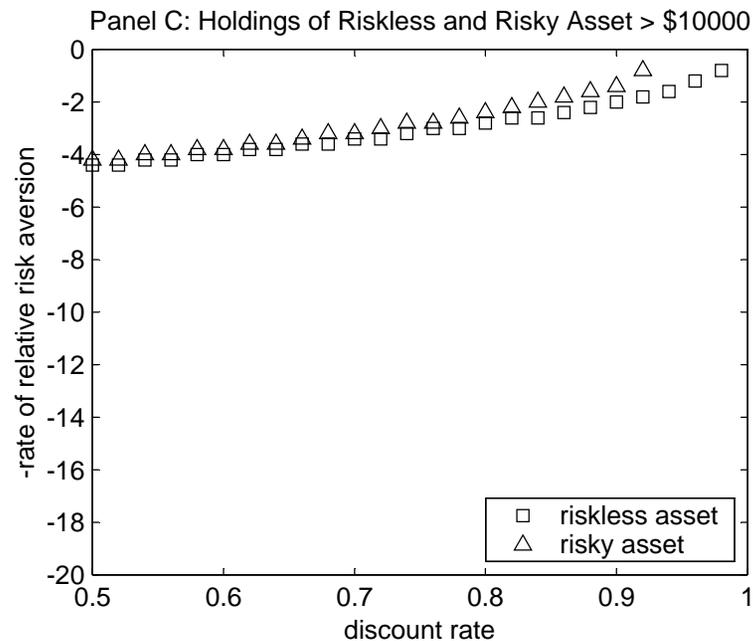
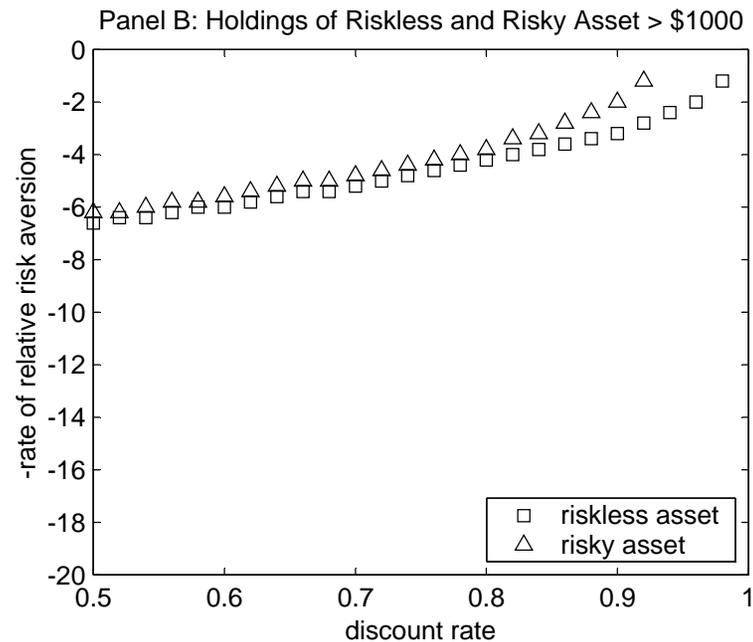
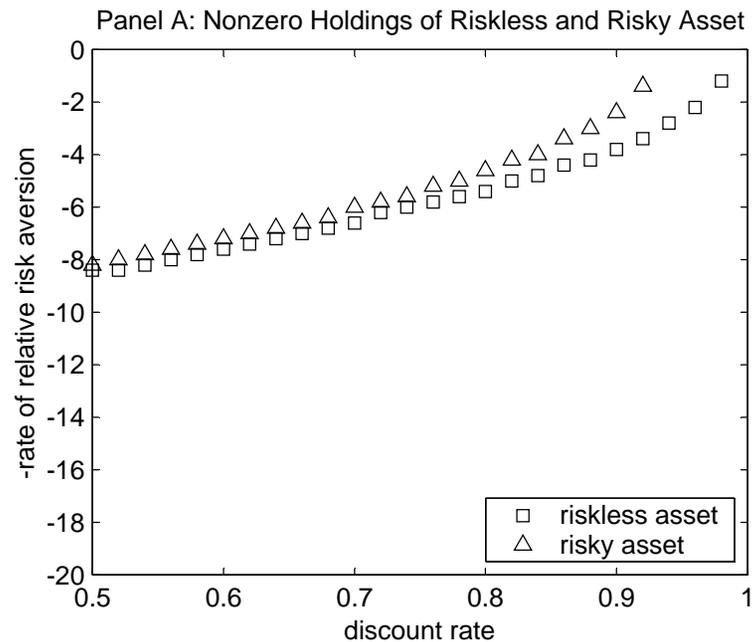


Figure 7: Euler equation error analysis for cohort PSID data. Cohorts are formed using age. Consumption growth is obtained by averaging over individual consumption growth. Squares and triangles indicate discount rate-risk aversion combinations that equate the Euler equation to zero in sample.

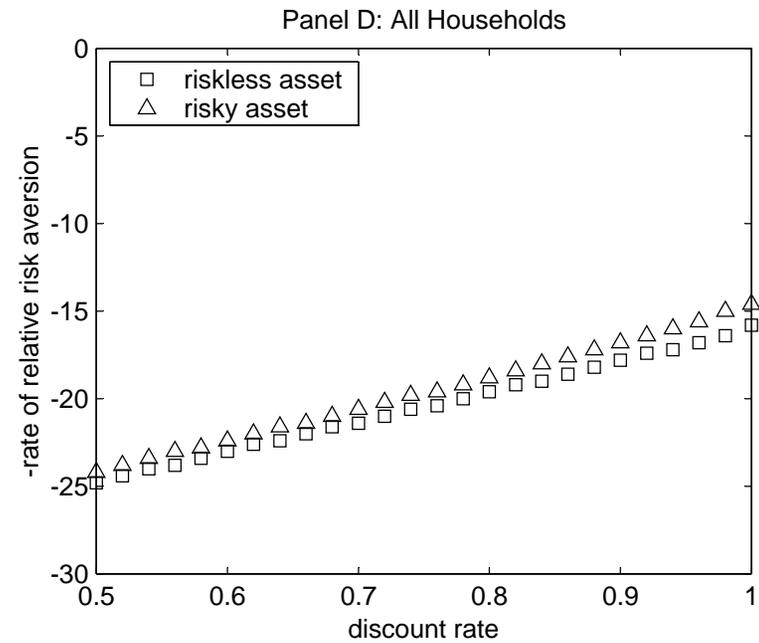
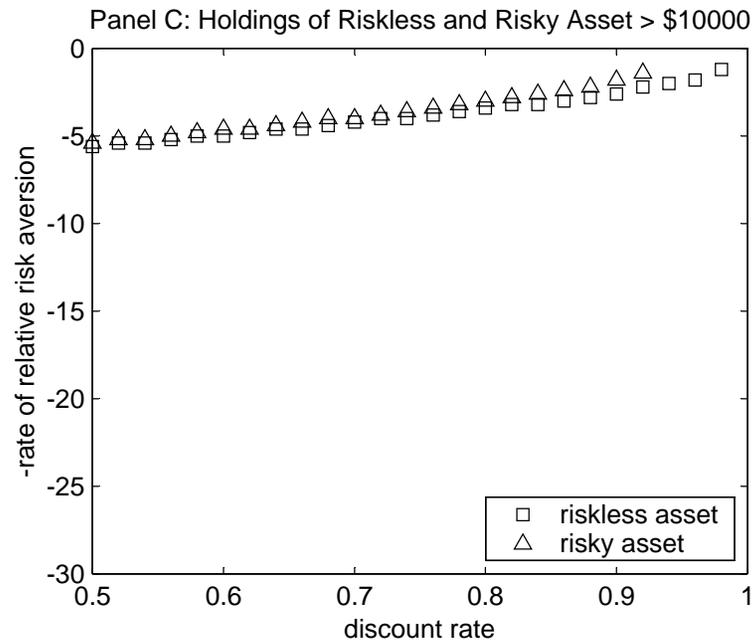
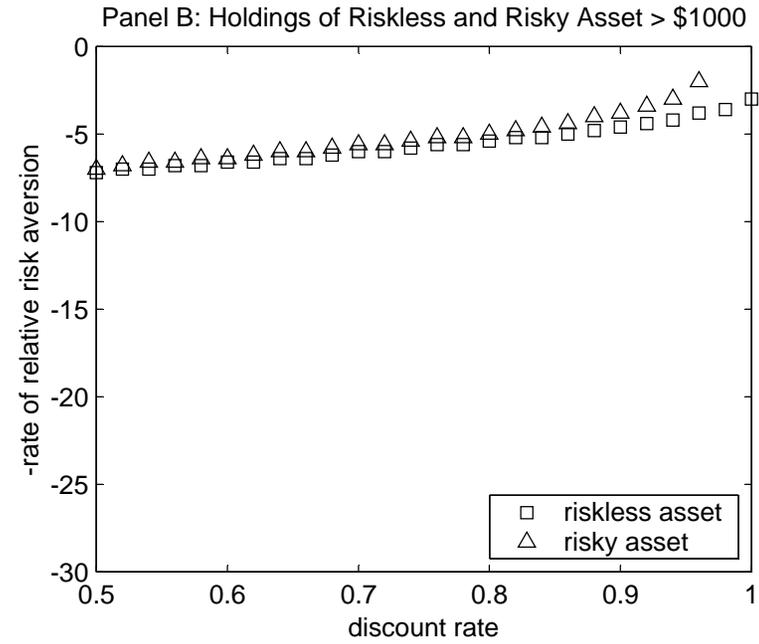
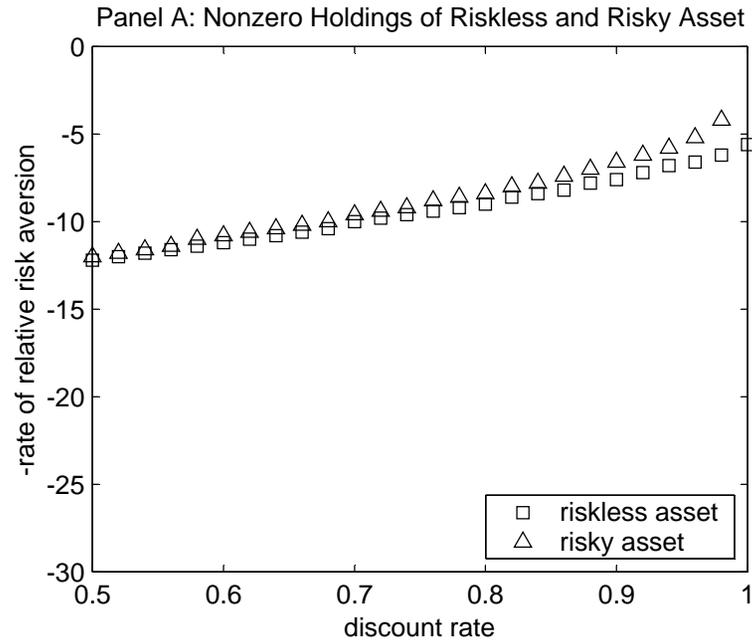


Figure 8: Euler equation error analysis for PSID "representative agent" data. The representative agent is obtained by averaging over consumption. Squares and triangles indicate discount rate-risk aversion combinations that equate the Euler equation to zero in sample.

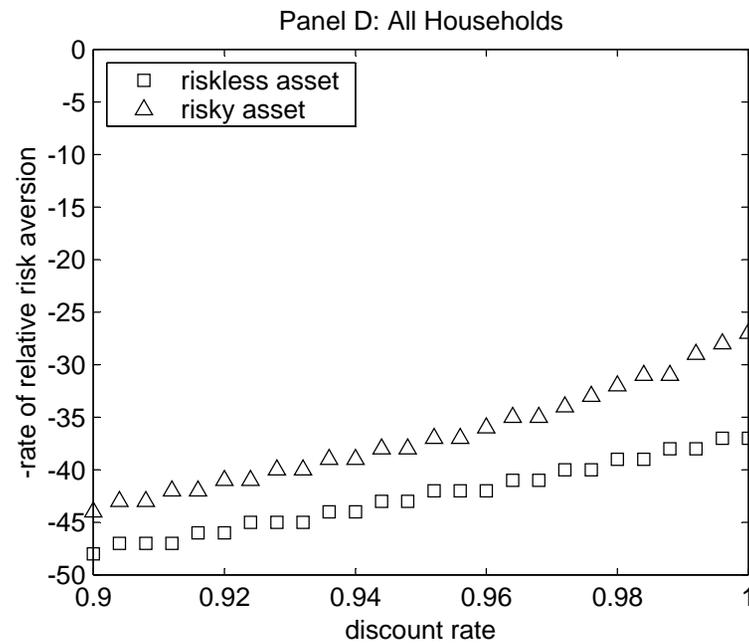
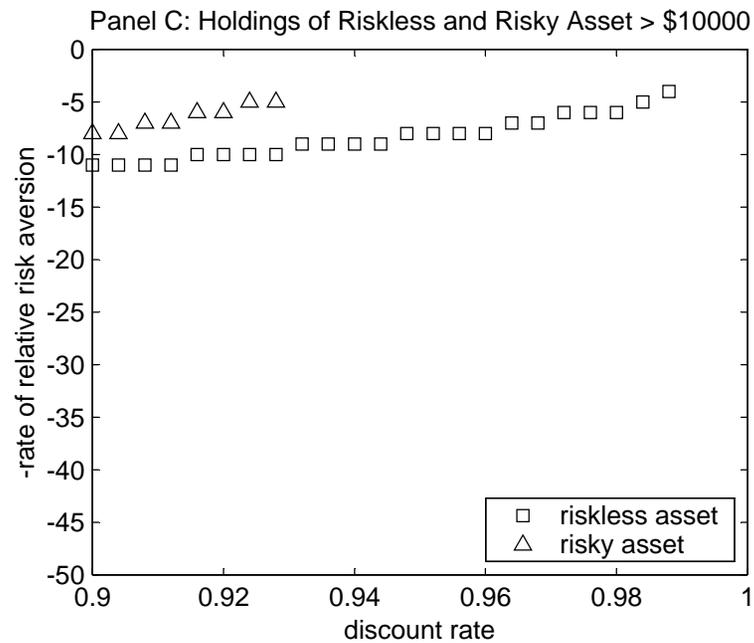
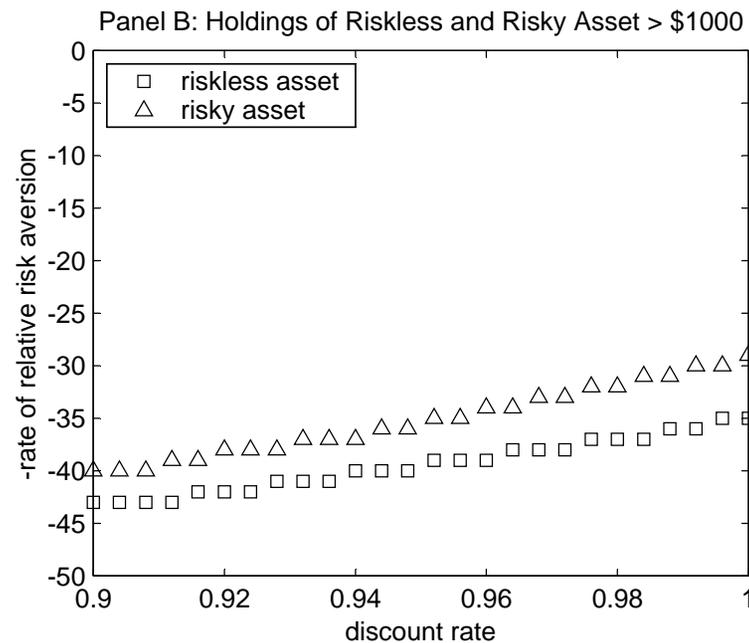
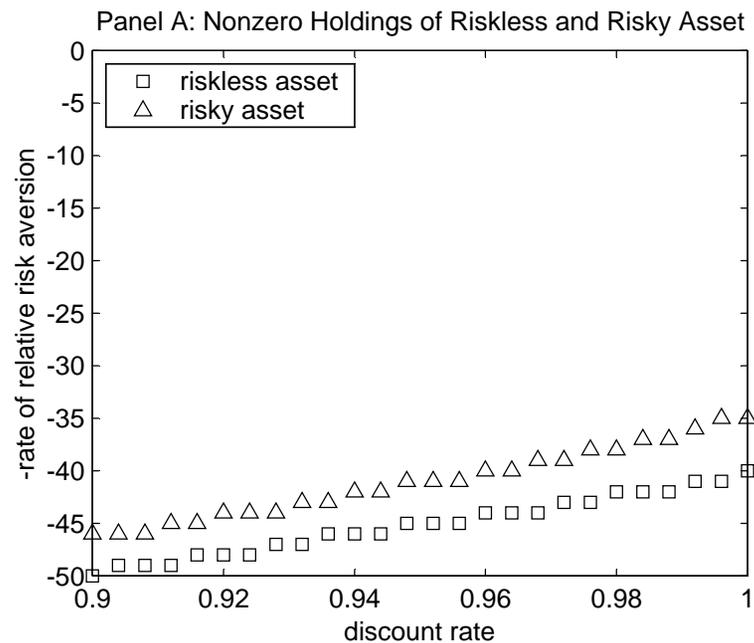


Figure 9: Euler equation error analysis for individual PSID data, exploiting the orthogonality between the Euler equation error and a lagged instrument (conditional information). All panels are for the dataset containing only households with holdings of the riskless and risky asset > \$1000.

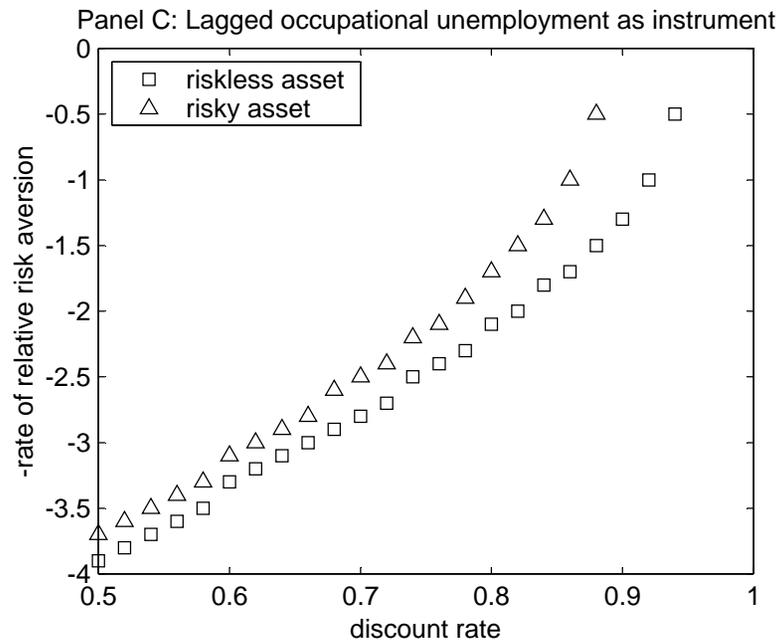
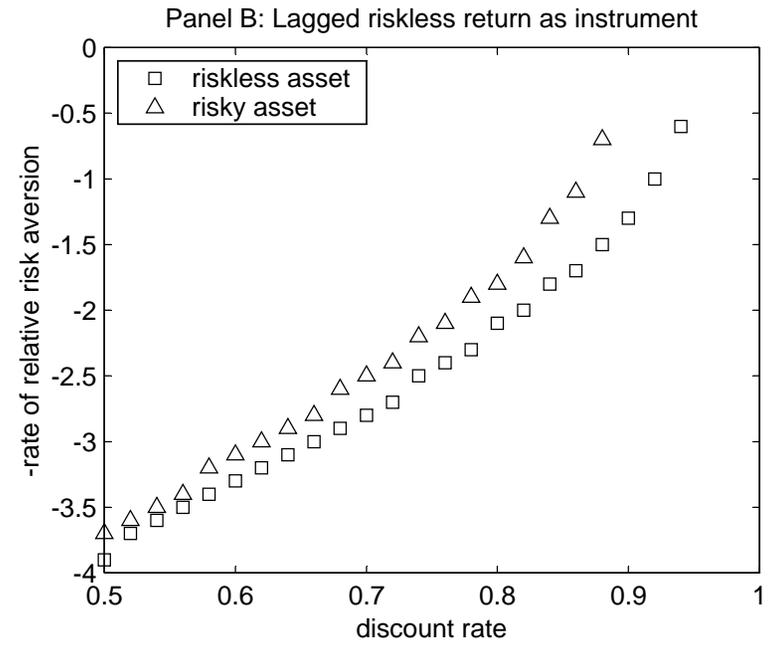
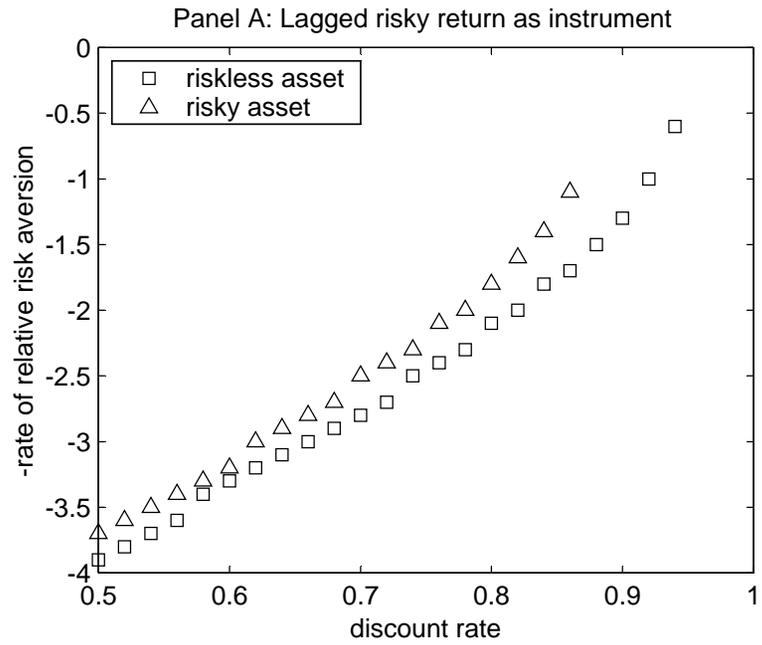


Figure 10: Euler equation error analysis for cohort PSID data, exploiting the orthogonality between the Euler equation error and a lagged instrument (conditional information). Cohorts are formed using age and education and consumption growth is obtained by averaging over individual consumption. All panels are for the dataset containing only households with holdings of the riskless and risky asset > \$1000.

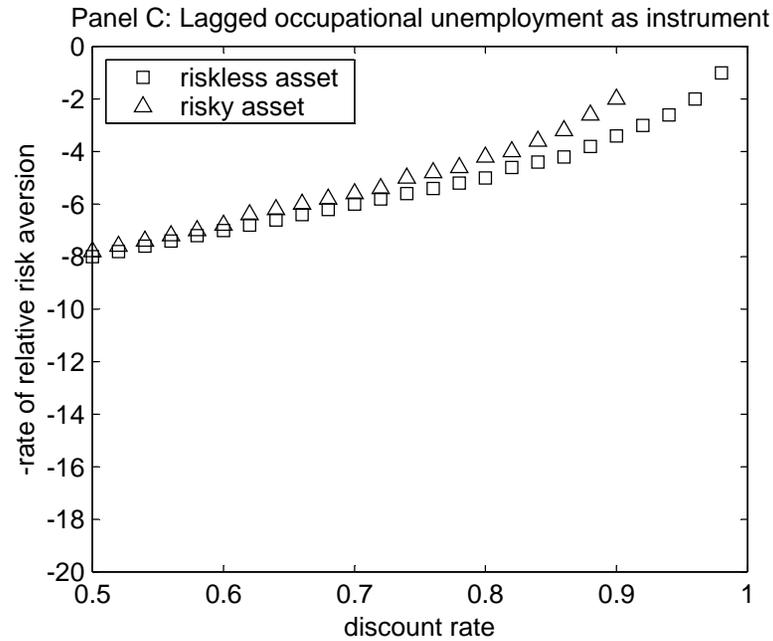
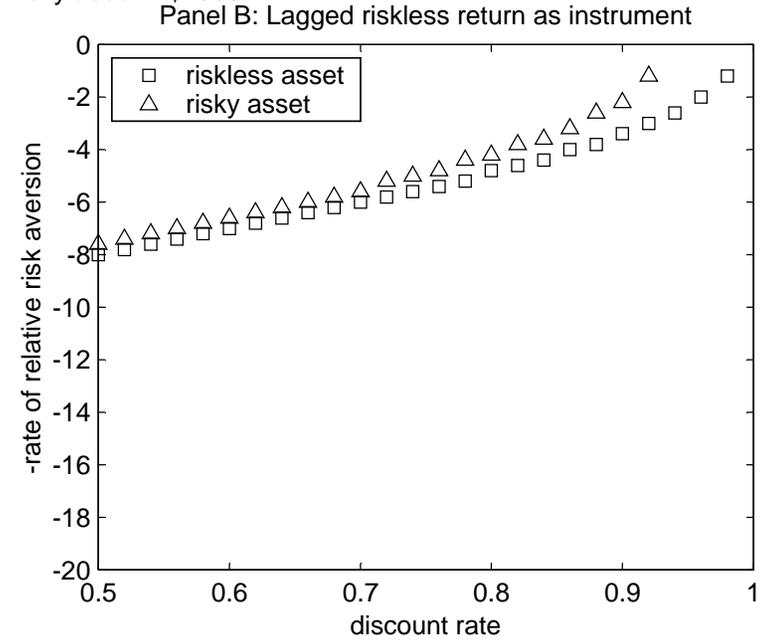
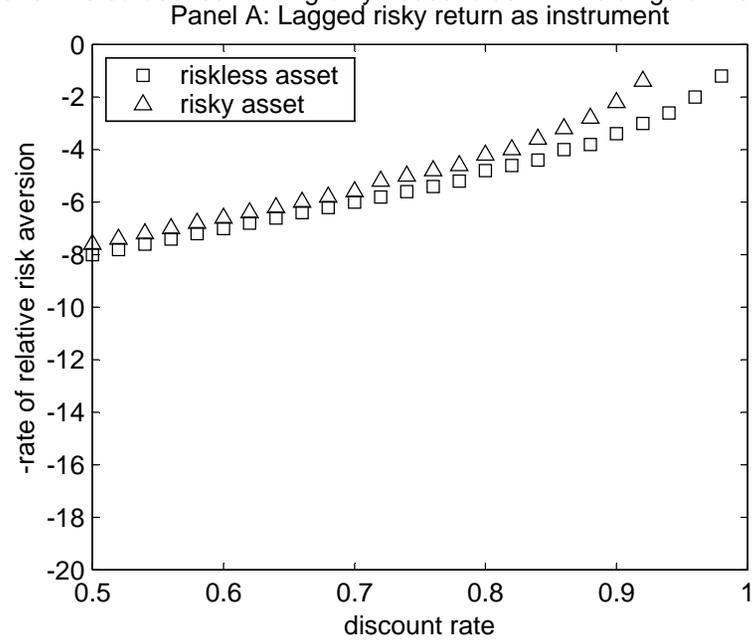


Figure 11: Euler equation error analysis for representative agent data, exploiting the orthogonality between the Euler equation error and a lagged instrument (conditional information). All panels are for the Mehra-Prescott data (yearly data 1889-1979)

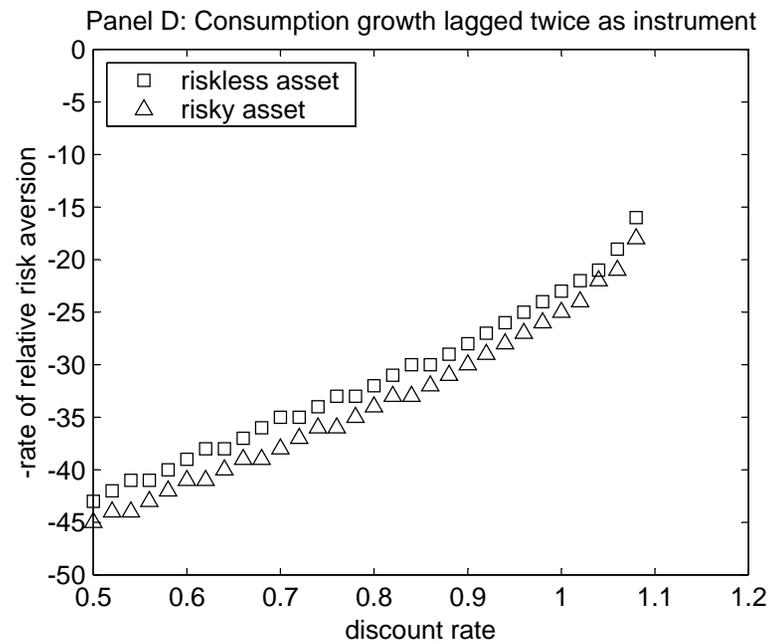
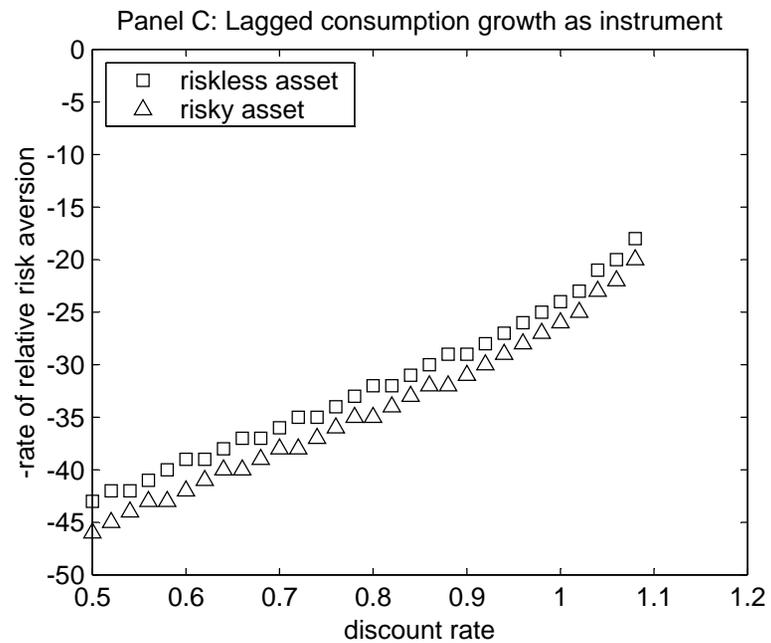
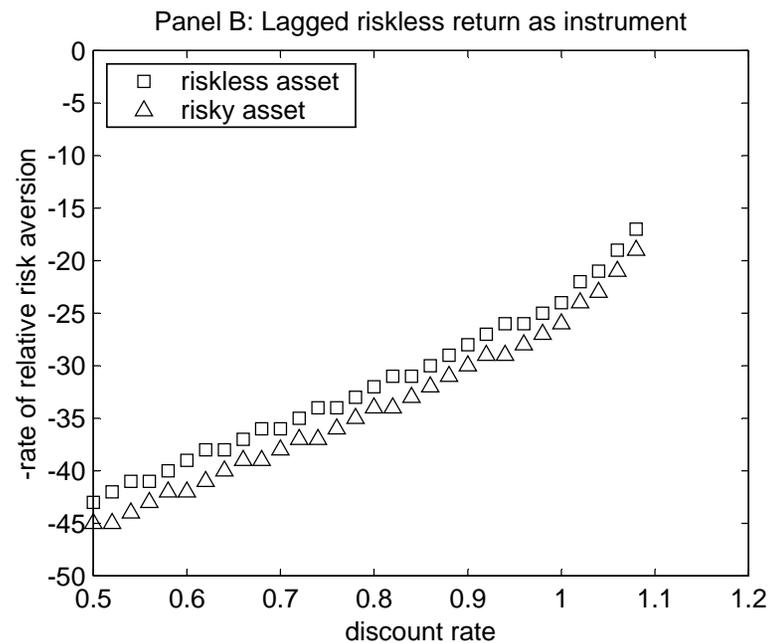
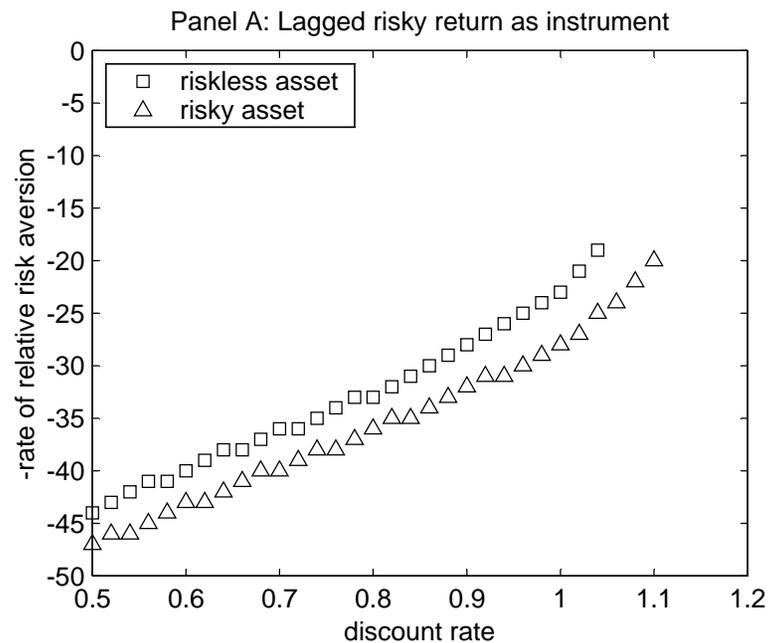
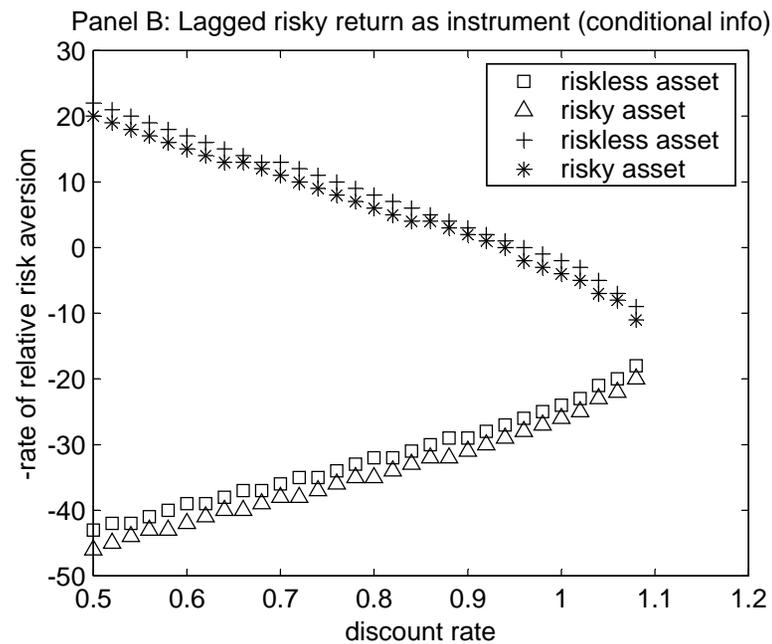
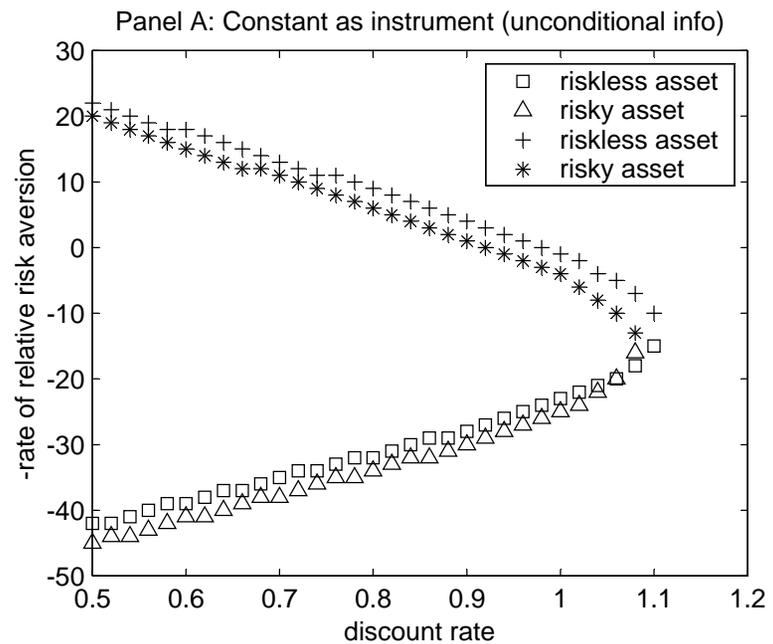


Figure 12: Euler equation error analysis for representative agent data.  
Both panels are for the Mehra-Prescott data (yearly data 1889-1979).



## Liste des publications au CIRANO\*

### Série Scientifique / *Scientific Series* (ISSN 1198-8177)

- 2002s-08 The Rate of Risk Aversion May Be Lower Than You Think / Kris Jacobs
- 2002s-07 A Structural Analysis of the Correlated Random Coefficient Wage Regression Model / Christian Belzil et Jörgen Hansen
- 2002s-06 Information Asymmetry, Insurance, and the Decision to Hospitalize / Åke Blomqvist et Pierre Thomas Léger
- 2002s-05 Coping with Stressful Decisions: Individual Differences, Appraisals and Choice / Ann-Renée Blais
- 2002s-04 A New Proof Of The Maximum Principle / Ngo Van Long et Koji Shimomura
- 2002s-03 Macro Surprises And Short-Term Behaviour In Bond Futures / Eugene Durenard et David Veredas
- 2002s-02 Financial Asset Returns, Market Timing, and Volatility Dynamics / Peter F. Christoffersen et Francis X. Diebold
- 2002s-01 An Empirical Analysis of Water Supply Contracts / Serge Garcia et Alban Thomas
- 2001s-71 A Theoretical Comparison Between Integrated and Realized Volatilities Modeling / Nour Meddahi
- 2001s-70 An Eigenfunction Approach for Volatility Modeling / Nour Meddahi
- 2001s-69 Dynamic Prevention in Short Term Insurance Contracts / M. Martin Boyer et Karine Gobert
- 2001s-68 Serial Cost Sharing in Multidimensional Contexts / Cyril Tédjédo et Michel Truchon
- 2001s-67 Learning from Strike / Fabienne Tournadre et Marie-Claire Villevall
- 2001s-66 Incentives in Common Agency / Bernard Sinclair-Desgagné
- 2001s-65 Detecting Multiple Breaks in Financial Market Volatility Dynamics / Elena Andreou et Eric Ghysels
- 2001s-64 Real Options, Preemption, and the Dynamics of Industry Investments / Marcel Boyer, Pierre Lasserre, Thomas Mariotti et Michel Moreaux
- 2001s-63 Dropout, School Performance and Working while in School: An Econometric Model with Heterogeneous Groups / Marcel Dagenais, Claude Montmarquette et Nathalie Viennot-Briot
- 2001s-62 Derivatives Do Affect Mutual Funds Returns : How and When? / Charles Cao, Eric Ghysels et Frank Hatheway
- 2001s-61 Conditional Quantiles of Volatility in Equity Index and Foreign Exchange Data / John W. Galbraith, Serguei Zernov and Victoria Zinde-Walsh
- 2001s-60 The Public-Private Sector Risk-Sharing in the French Insurance "Cat. Nat. System" / Nathalie de Marcellis-Warin et Erwann Michel-Kerjan

---

\* Consultez la liste complète des publications du CIRANO et les publications elles-mêmes sur notre site Internet :

- 2001s-59 Compensation and Auditing with Correlated Information / M. Martin Boyer et Patrick González
- 2001s-58 Resistance is Futile: An Essay in Crime and Commitment / M. Martin Boyer
- 2001s-57 The Unreliability of Output Gap Estimates in Real Time / Athanasios Orphanides et Simon van Norden
- 2001s-56 Exact Nonparametric Two-Sample Homogeneity Tests for Possibly Discrete Distributions / Jean-Marie Dufour et Abdeljelil Farhat
- 2001s-55 Les coûts de la réglementation : une revue de la littérature / Robert Gagné, Paul Lanoie, Pierre-Carl Micheud et Michel Patry
- 2001s-54 Testing for structural Change in the Presence of Auxiliary Models / Eric Ghysels et Alain Guay
- 2001s-53 Environmental Regulation and Productivity: New Findings on the Porter Hypothesis / Paul Lanoie, Michel Patry et Richard Lajeunesse
- 2001s-52 The Aftermarket Performance of Initial Public Offerings in Canada / Maher Kooli et Jean-Marc Suret
- 2001s-51 Capital Structure and Risk Management / Karine Gobert
- 2001s-50 The Underpricing of Initial Public Offerings: Further Canadian Evidence / Maher Kooli et Jean-Marc Suret
- 2001s-49 How Innovative Are Canadian Firms Compared to Some European Firms? A Comparative Look at Innovation Surveys / Pierre Mohnen et Pierre Therrien
- 2001s-48 A Tale of Two Ports / Ngo Van Long et Kar-yiu Wong
- 2001s-47 Wage Policy of Firms: An Empirical Investigation / Stéphanie Lluís
- 2001s-46 Forecasting Some Low-Predictability Time Series Using Diffusion Indices / Marc Brisson, Bryan Campbell et John W. Galbraith
- 2001s-45 The Importance of the Loss Function in Option Pricing / Peter Christoffersen et Kris Jacobs
- 2001s-44 Let's Get "Real" about Using Economic Data / Peter Christoffersen, Eric Ghysels et Norman R. Swanson
- 2001s-43 Fragmentation, Outsourcing and the Service Sector / Ngo Van Long, Ray Riezman et Antoine Soubeyran
- 2001s-42 Nonlinear Features of Realized FX Volatility / John M. Maheu et Thomas H. McCurdy
- 2001s-41 Job Satisfaction and Quits: Theory and Evidence from the German Socioeconomic Panel / Louis Lévy-Garboua, Claude Montmarquette et Véronique Simonnet
- 2001s-40 Logique et tests d'hypothèse : réflexions sur les problèmes mal posés en économétrie / Jean-Marie Dufour
- 2001s-39 Managing IT Outsourcing Risk: Lessons Learned / Benoit A. Aubert, Suzanne Rivard et Michel Patry
- 2001s-38 Organizational Design of R&D Activities / Stefan Ambec et Michel Poitevin